



Gene flow in European coal tits (Aves, Passeriformes, *Periparus ater*): low among Mediterranean populations but high in a continental contact zone

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Gene flow in European coal tits (Aves, Passeriformes, *Periparus ater*): low among Mediterranean populations but high in a continental contact zone

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Abstract

Extant phylogeographic patterns of Palearctic terrestrial vertebrates are generally believed to have originated from glacial range fragmentation. Post-Pleistocene range expansions have led to the formation of secondary contact zones among genetically distinct taxa. For coal tits (*Periparus ater*), such a contact zone has been localized in Germany. In this study, we quantified gene flow between Fennoscandian and southern European coal tits using a set of 13 microsatellite loci. STRUCTURE analysis revealed four genetic clusters two of these on Mediterranean islands. German populations were genetically admixed but introgression of southern alleles was evident for Fennoscandian populations. In the South, we found negligible introgression of northern alleles (and haplotypes) but slight admixture of two southern genetic clusters in the Pyrenees and on the Balkan Peninsula and near complete sorting of these two allelic lineages on the islands of Corsica and Sardinia. Genetic distinctiveness of the Mediterranean island populations reflects general patterns of endemism in the Corso-Sardinian fauna and the Cypriot fauna. Wide-range gene flow in Central Europe suggests a broad zone of intergradation between subspecies of the coal tit rather than a narrow contact zone. This is in accordance with low morphological and bioacoustic differentiation of European coal tit populations.

Key words: phylogeography – island populations – microsatellites – subspecies – glacial refugia

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3 1 **Introduction**

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5 2 Evolutionary biologists widely agree that glacial impact considerably shaped phylogeographic
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7 3 patterns and speciation of terrestrial vertebrates in the Palearctic (Avice & Walker, 1998;
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9 4 Hewitt, 2000, 2004; Lovette, 2005; Zink *et al.*, 2008; Stewart *et al.*, 2010). Pleistocene
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11 5 separation of Eastern and Western Palearctic populations led to divergence of gene pools
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13 6 among distant refugia, in a few extreme cases across a large extant distributional gap, such as
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15 7 seen in the marsh tit, *Poecile palustris* (Tritsch *et al.*, 2017) or in the azure-winged magpie,
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17 8 *Cyanopica cyanus* (Zhang *et al.*, 2012). Other East-West lineage splits dating back to
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19 9 Pleistocene events were reconstructed for example in corvids (Haring *et al.*, 2007, 2012) and
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21 10 tits (Kvist *et al.* 2003; Päckert *et al.*, 2005; Kvist & Rytönen, 2006). One noticeable result
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23 11 from Holocene range expansion is the spatial overlap of genetically distinct populations that
24
25 12 is manifested in secondary contact zones of a highly variable extent (Woodruff, 1973; Haffer,
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27 13 1989; Aliabadian *et al.*, 2005). In Western Europe, the apparent spatial clustering of
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29 14 secondary contact zones among terrestrial vertebrate sister taxa was the result of postglacial
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31 15 expansion from southern glacial refugia (Hewitt, 2000; Schmitt, 2007). Parapatry along sharp
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33 16 and narrow hybrid zones is typically found at geographic barriers, such as the European
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35 17 mountain systems that separate two larger glacial refugia from the continent: i) the Iberian
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37 18 Peninsula in the Pyrenees (Fig. 1A; birds: Helbig *et al.*, 2001; Pons *et al.*, 2011; Backström *et*
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39 19 *al.*, 2013; Kuhn *et al.*, 2013; reptiles: Mila *et al.*, 2013; insects: Vasquez *et al.*, 1994; Shuker
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41 20 *et al.*, 2005; Bella *et al.*, 2009); ii) the Italian Peninsula in the Alps (Fig. 1A) II; birds:
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43 21 Hermansen *et al.*, 2011; rodents: Sutter *et al.*, 2013; Giménez *et al.*, 2017; insects: Flanagan *et*
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45 22 *al.*, 1999). Apart from parapatry across mountain ranges, narrow hybrid zones of a wide
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47 23 latitudinal extent exist in Central Europe (Fig. 1A), the best-studied examples being those of
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49 24 crows (*Corvus c. corone*, *C. c. cornix*: Haas *et al.*, 2009, 2010; Wolf *et al.*, 2010; Poelstra *et*
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51 25 *al.*, 2014a, b; other birds: Secondi *et al.*, 2011), the house mouse (*Mus m. musculus*, *M. m.*
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53 26 *domesticus*: Macholán *et al.*, 2008; Giménez *et al.*, 2017) and hedgehogs (*Erinaceus*
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27 *europaeus*, *E. roumanicus*: Berggren *et al.*, 2005; Bolfiková & Hulva, 2012; Waters *et al.*
28 [2013: Fig. 1B]; Pfäffle *et al.*, 2014).

29 In several bird species pairs, zones of secondary contact and hybridization are not
30 restricted to a narrow band, but extend along a wide longitudinal range into Eastern Europe
31 (Fig. 1B), such as found in flycatchers (*Ficedula*: Sætre *et al.*, 2001; Hogner *et al.*, 2012a),
32 reed warblers (*Acrocephalus*: Reifová *et al.*, 2016), nightingales (*Luscinia*: Vokurkova *et al.*,
33 2013), tits (*Cyanistes*: Woodruff, 1973; Stervander *et al.*, 2015) and Old World buntings
34 (*Emberiza*: Irwin *et al.*, 2009).

35 In a few other examples, pre-mating barriers were either not established during a short
36 separation time in refuge areas or they simply broke down in secondary contact, which had
37 led to merging of divergent genetic lineages. The signal from mitochondrial markers might
38 then remain the only testimony of past (Pleistocene) lineage separation, presently contrasted
39 by narrow or wide-range gene flow (Fig. 1C), as suggested for some passerine bird species
40 (Zink *et al.*, 2008; Päckert *et al.*, 2010; Hogner *et al.* 2012b; Block *et al.*, 2015). Interbreeding
41 and merging of gene pools between cryptic genetic lineages of a phenotypically uniform
42 species has been sometimes termed “speciation in reverse” (e.g. in birds Webb *et al.*, 2011).
43 However, in the strict sense, reverse speciation is more appropriately applied to those
44 examples where gene pools become largely absorbed due to hybridization with one of the two
45 parental species running the risk of going extinct (in fishes: Seehausen *et al.*, 2008; Taylor *et al.*,
46 2006; Hudson *et al.*, 2013; Bath *et al.*, 2014; in Darwin’s finches Kleindorfer *et al.*, 2014).

47 Our study focuses on the recent evidence of secondary range overlap in Western
48 Europe among a north-eastern and a south-western mitochondrial lineage of the coal tit,
49 *Parus ater* (Pentzold *et al.*, 2013). In this study, we aim at verifying the extent and degree
50 of nuclear gene flow among the two coal tit lineages using nuclear markers (microsatellites).
51 We expect significant gene flow at least in the region of considerable mtDNA lineage overlap
52 at a contact zone extending throughout Germany (Fig. 2). We also expect nuclear gene flow

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53 to extend across a wider range than mitochondrial introgression, as was shown for another
54 parid hybrid zone (*Parus major*/ *P. minor*: Kvist & Rytönen, 2006).

56 **Material and methods**

57 **Study species**

58 Eight divergent mitochondrial lineages are presently known in the coal tit. These lineages are
59 distributed across large parts of the Palearctic, the mountain forests of China, the Himalayas,
60 Karakoram and Hindu Kush as well as on Taiwan (Tietze *et al.*, 2011). Across the Western
61 Palearctic, four distinct mtDNA lineages of the coal tit occur (Fig. 2): i) the north-eastern
62 Palearctic (*ater* subspecies group; distributed from Northern Europe across the Eurasian
63 continent to the Pacific coast and Japan), ii) Central and Southern Europe (*abietum* subspecies
64 group) including the British Isles and the islands of Corsica and Sardinia, iii) North Africa
65 and iv) Cyprus (Martens *et al.*, 2006, Tietze *et al.*, 2011, Pentzold *et al.*, 2013). Range overlap
66 of the south-western *abietum* and the north-eastern *ater* lineages could so far be restricted to
67 the German populations only (Fig. 2). Due to a lack of reliable morphological and bioacoustic
68 distinctiveness of north-eastern versus south-western Palearctic coal tits, the spatial dimension
69 of the contact zone cannot be delimited by geographical variation of phenotypes or song types
70 (Tietze *et al.*, 2011; Pentzold *et al.*, 2013, 2016).

72 **Sampling and multilocus genotyping**

73 We sampled 166 birds from Russia, Kyrgyzstan, Kazakhstan, Finland, Norway, Germany,
74 French Pyrenees, Corsica and Sardinia. DNA preparation was conducted either using the
75 innuPREP DNA Mini Kit (muscle tissue, Analytik Jena AG, Germany) or the PEQLAB
76 GOLD Blood DNA Mini Kit (blood samples, PEQLAB Biotechnologie GmbH, Germany),
77 following the manufacturers' advice.

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3 78 New microsatellite loci for *P. ater* were identified by Ecogenics GmbH (Zürich,
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5 79 Switzerland) based on an enriched DNA library. Size selected genomic DNA was ligated into
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7 80 SNXforward/SNX reverse-linker (Hamilton *et al.*, 1999) and enriched by magnetic bead
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9 81 selection with biotin-labelled oligonucleotide repeats ((CT)₁₃, (GT)₁₃, (GTAT)₇, (GATA)₇;
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11 82 Gautschi *et al.*, 2000 a,b). A total number of 528 recombinant colonies were screened and 415
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13 83 gave a positive signal after dot-blot hybridization. Plasmids from 48 positive clones were
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15 84 Sanger sequenced and primers were designed for 16 microsatellite inserts of which nine
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17 85 (listed in Table 1) were finally used for amplification of polymorphic microsatellite loci using
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19 86 the protocol described by Schuelke (2000). For this protocol a M13(-21) tail (18 bp) was
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21 87 adhered to the 5' end of the forward primer. The reverse primer remained unmodified. In
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23 88 addition to the two regular PCR primers, a fluorescent-labelled universal M13(-21) primer
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25 89 was added to the reaction mixture. The reaction contained 10 to 40 ng of template DNA,
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27 90 0.04 µM of the M13-forward primer, 0.16 µM of the reverse primer and the labelled M13
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29 91 primer, 0.2 mM of each dNTP, 1 µL of 10x PCR reaction buffer "complete" and 0.5 units of
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31 92 DFS-Taq DNA polymerase (Bioron GmbH, Germany) in a total volume of 10 µL. The
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33 93 thermo-treatment consisted of two successive steps: a) amplification of the microsatellite
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35 94 fragment with M13 fusion, b) labelling of the fragment with the fluorescent dye. The PCR
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37 95 program was 95 °C for 10 min followed by 30 cycles of 30 s of 95 °C, 45 s of 50 °C (Parate8)
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39 96 or rather 56 °C (all other Parate loci) and 45 s of 72 °C (step a), followed by eight cycles of
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41 97 30 s of 95 °C, 45 s of 53 °C and 45 s of 72 °C (step b) and a final elongation at 72 °C for 30
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43 98 min.
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48 99 Four additional primer pairs targeting microsatellite loci were obtained from previous
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50 100 studies on *Poecile atricapillus* (Table 1). We tested for cross amplification with *P. ater*
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52 101 samples in a total volume of 10 µL containing 10 to 40 ng of template DNA, 0.3 µM of each
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54 102 primer, 0.2 mM of each dNTP, 1 µL of 10x PCR reaction buffer "complete" and 0.5 units of
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56 103 DFS-Taq DNA polymerase (Bioron GmbH, Germany). The thermo-cycling protocol was as
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104 follows: 94 °C for 5 min followed by 35 cycles of 94 °C for 30 s, 60 °C (Pat2-43) or 57 °C
105 (PmaC25, PmaTGAn33, Pma69) for 30 s, 72 °C for 45 s and a final elongation at 72 °C for 5
106 min.

107 Specimens were genotyped at 13 loci (Table 1) using the reaction conditions outlined
108 above and labelled PCR fragments were run on a 16-column ABI 3130xl capillary sequencer
109 (Applied Biosystems). The alleles were scored using the STRand Analysis Software vers. 2.4
110 (UC Davis, veterinary genetics lab, <http://www.vgl.ucdavis.edu/STRand>; Toonen and
111 Hughes, 2001).

112 The software package MICROCHECKER 2.2.3 (van Oosterhout *et al.*, 2004) was
113 used to test the probability that experimental errors occurred during microsatellite genotyping,
114 i.e. large allelic dropout, scoring errors due to misinterpretation of stutter bands and null
115 alleles.

116
117 **Diversity and Divergence**

118 Summaries of allele sizes and the existence and frequencies of population specific alleles
119 (private alleles) were calculated using the program CONVERT vers. 1.31 (Glaubitz, 2004),
120 which was also employed to generate input files for various software packages. Linkage
121 between loci was determined using ARLEQUIN vers 3.5.1.3 (Excoffier *et al.*, 2005). The
122 same software package was used to calculate locus-specific observed and expected
123 heterozygosities (H_O , H_E) for each sample population and to test for locus-specific deviations
124 from Hardy Weinberg expectations (HWE). Population specific deviations from HWE
125 (excess or deficiency of heterozygosity) across all loci were explored using inbreeding
126 coefficients (F_{IS}), calculated with the software FStat (ver. 2.9.3.2; Goudet 1995) using a
127 randomisation test (3600 randomisations) to test for significance. The same software was used
128 to estimate the mean number of alleles per locus and populations as well as the mean allelic
129 richness (AR) per population across all loci. For these, we analysed a reduced dataset

consisting of 145 specimens that belong to 14 distinct populations with a minimum of five specimens per population (Table 2, Table S1).

Divergence between populations was estimated using F-statistics (inferred from microsatellite allele frequencies) and Φ -statistics (inferred from mitochondrial nucleotide sequences) by pairwise F_{ST} and Φ_{ST} values as well as by non-hierarchical and hierarchical locus-by-locus analysis of molecular variance (AMOVA with F_{CT} and Φ_{CT} values as a measure of divergence among groups) using ARLEQUIN. 20000 permutations were performed to test for significance of these values. All p-values obtained from tests implementing multiple comparisons (i.e. test for deviation from HWE expectation, test for linkage between loci) were Bonferroni corrected to adjust the significance threshold (Rice, 1989). In order to depict divergence between populations, pairwise F'_{ST} values were used in a distance matrix to construct a UPGMA phenogram with MEGA v.6 (Tamura *et al.*, 2013).

Bayesian inference of the population structure

Non-spatial Bayesian inference of population structure was performed using the software package STRUCTURE vers. 2.3.3. (Pritchard *et al.*, 2000; Falush *et al.*, 2003). STRUCTURE runs were performed i) under the *a priori* assumption of genetic admixture and correlated allele frequencies and ii) under a LOCPRIOR model that allows for classification of the individuals into groups, which are given to the algorithm as an *a priori* parameter (Hubisz *et al.*, 2009). The model was run under two different LOCPRIOR settings: i) by classifying the individuals of the complete data set ($n = 166$) according to their assignment to mitochondrial lineages (inferred from the data set by Pentzold *et al.*, 2013) and ii) by assigning the individuals to 14 local populations of $n \geq 5$ (total sampling $n = 145$, Table 2).

All STRUCTURE runs were conducted for 1–10 putative genetic clusters (K) with ten replicates for each value of K. The number of MCMC runs was 10^5 with a burn-in period of 25 000 throughout all model runs. For further processing of the STRUCTURE output,

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STRUCTURE HARVESTER (Earl & von Holdt, 2012) was used. In order to select the most likely number of genetic clusters (K), we used the approach by Evanno *et al.* (2005). STRUCTURE analysis was also used to estimate the extent of genetic admixture in different populations according to the method described by Randi (2008). Accordingly, we used a threshold $q > 0.80$ for the assignment of individuals to a cluster or we classified individuals as admixed individuals, if the proportion of membership was $q < 0.80$ (see Randi, 2008).

Spatial Bayesian clustering was performed using the software packages TESS vers. 2.3.1 (Chen *et al.*, 2007) and GENELAND vers. 4.03 (Guillot *et al.*, 2005b). Unlike the non-spatial models of STRUCTURE, the spatially explicit models implemented in TESS and GENELAND consider the geographic coordinates of the samples, but do not consider affiliation as a model parameter. Data exploration under different models is recommendable, because explanatory power of spatial *versus* non-spatial models depends on demographic scenarios to be tested and clustering output from one model-based method might reveal a finer scaled spatial structure that other models fail to detect (François & Durand 2010). Both admixture models in TESS (BYM and CAR model, Durand *et al.*, 2009a, b) were run with the complete dataset of $n = 166$ individuals (MCMC iterations: 10^5 , burn-in period: 20 000, Kmax: 2–10, five replicates for each Kmax for both models). Unlike STRUCTURE or TESS the standard models of GENELAND do not account for admixture, but assign posterior probabilities of cluster membership to the single individuals. A benefit of GENELAND is that it can correct for the occurrence of null alleles. Although from a biological perspective it seems obvious to assume the allele frequencies to be correlated between populations, the respective model was assessed to systematically overestimate the number of clusters (Guillot *et al.*, 2005a). Hence according to the authors' recommendations the analysis was conducted in two steps: first, resolving the number of populations (K) using the D model (*frequency model* = uncorrelated; K: 1–10); second, deriving the correct population assignment by applying the F model (*frequency model* = correlated) with a fixed number of K (which was

determined in the first step, Guillot *et al.*, 2005a, GENELAND documentation). The model parameters were: 10^6 MCMC iterations, Thinning = 1000, Null allele model = TRUE and ten replicates per analysis step. The outputs of STRUCTURE and TESS both were further processed with CLUMPP (Jakobsson & Rosenberg, 2007). For visualization, DISTRUCT (Rosenberg, 2004) was used.

Admixture rate in the hybrid zone

For estimating admixture rate in the hybrid zone we applied demographic modelling based on Approximate Bayesian Computation (ABC) using the program DIYABC v. 2.0.4 (Cornuet *et al.*, 2010). We first used both microsatellite and mitochondrial control-region sequences and included individuals from which both data were available. We chose samples from Norway and Finland ($n = 18$) with q -values from STRUCTURE above 0.8 to their cluster ‘*ater*’ to represent the ‘northern’ parental population and samples from French Pyrenees with q -values from STRUCTURE above 0.8 to their cluster ‘*abietum*’ to represent the ‘southern’ ($n = 8$) parental population. This was based on further evidence that these populations represented only one mitochondrial lineage each and were clearly separated from, but still related to the admixture populations in the STRUCTURE analysis from the microsatellite data. In the central European admixture population, we included samples from Schleswig Holstein, Harz, Saxony, Palatine Forest and Black Forest ($n = 47$, all Germany).

We started by constructing four historical models (Fig. 3): 1) parental ‘northern’ population and ‘southern’ population were split from each other at time t_2 and come into a contact at time t_1 to form the admixture population, 2) the parental populations split first from each other and the central European population was split later from the northern population, 3) the northern and central populations split first from each other and the southern population was split later from the northern population and 4) the parental populations split first from each other and the central European population was split later from the southern population.

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3 208 Mutation rate for the microsatellite data was set to 10^{-4} – 10^{-3} as was done with another tit
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5 209 species, the blue tit *Cyanistes caeruleus* (Hansson *et al.*, 2014). For the mitochondrial
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7 210 sequences, we applied the substitution rate 1.156×10^{-8} calibrated for coal tit control region in
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9 211 Pentzold *et al.* (2013), HKY+Gamma model with gamma = 0.09 (as suggested by the test for
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11 212 the substitution model implemented in MEGA v. 6.06; Tamura *et al.*, 2013). As the fit of
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13 213 observed data with the simulated data was poor, we next performed the same analysis
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15 214 separately for the microsatellite data and the mitochondrial data. For the microsatellite data,
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17 215 uniform prior distributions for effective population sizes were set to 10–10 000 and for
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19 216 coalescence times, t1 was set to 10–1000 and t2 to 10–4000. The uniform prior distribution
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21 217 for the admixture rate was 0.001–0.999. The priors for effective population sizes and
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23 218 coalescence times were changed for mitochondrial analyses to $N_e = 1000$ – $1\,000\,000$ for
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25 219 ‘northern’, ‘southern’ and ancient populations, $N_e = 1000$ – $2\,000\,000$ for the population at the
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27 220 contact zone, t1 = 10–20 000 and t2 = 1000–2 000 000, as the fit of the observed and
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29 221 simulated data was poor when using the same priors as for the microsatellite data. Altogether,
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31 222 4 000 000 data sets were simulated for both microsatellite and mitochondrial data.

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35 223 For calculation of time of divergence and time since admixture we assumed a mean
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37 224 generation time of approximately two years in tits (as applied by Hansson *et al.*, 2014;
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39 225 compare 1.5 years for the great tit, *Parus major*, in Qu *et al.* (2015) and 2.26 years for the
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41 226 willow tit, *Poecile montanus*, in Kvist *et al.* (2001)). We expect time estimates inferred from
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43 227 mitochondrial DNA to correspond with the onset of lineage splitting and admixture caused by
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45 228 paleoclimatic events more accurately, because nuclear loci will generally reach coalescence
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47 229 slower and at a later stage of evolution (Palumbi *et al.*, 2001). Therefore, time estimates
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49 230 inferred from microsatellite data are generally supposed to post-date paleoclimatic events that
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51 231 triggered lineage divergence.

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54 232 Furthermore, the ratio between male to female gene flow (m_m/m_f) was calculated
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56 233 according to the equations in Hedrick *et al.* (2013). The main assumptions implemented in
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this approach are that populations can be described according to the island model and that populations are in migration-drift equilibrium. The approach uses divergence levels caused by female gene flow as $F_{ST(f)}$ values derived from mtDNA and estimates divergence levels caused by male gene flow $F_{ST(m)}$ from microsatellites (Eqn. 7a; Hedrick *et al.*, 2013). Both divergence levels were used to calculate the m_m/m_f ratio (Eqn. 7b; Hedrick *et al.*, 2013). Given the island model and migration drift equilibrium as basic assumptions, the estimates were alternatively performed i) for the total set of populations and ii) under exclusion of Mediterranean island populations for all continental Eurasian populations.

Results

Microsatellite genotyping

Deviations from HWE were predominately found at loci Parate 6 (six populations) and Parate 8 (seven populations), but also in single populations at loci PmaC25, Parate 15, Parate 16, Parate 3 and Parate 2. The occurrence of HWE deviations at these loci was associated with the presence of null alleles (Table S1). Most HWE deviations were found in the population from Schleswig-Holstein, which also had the highest positive F_{IS} value (Table 2). Furthermore, in this population we found evidence of linkage disequilibrium among alleles at six loci (Table S1). None of the other study populations showed a signal of linkage disequilibrium except two island populations from Cyprus and Corsica (at two loci each; Table S1). Our analyses did not provide evidence of genotyping artefacts due to large allele dropout and misscoring of genotypes due to stuttering was almost absent except at locus Parate 8 in a single population (Cyprus). These analyses suggested that loci Parate 6 and Parate 8 should be treated with a precaution in subsequent analyses (see below).

Diversity and Divergence

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Given the null allele bias at loci Parate 6 and Parate 8, we excluded these data from estimating diversity and divergence for 14 populations ($n = 145$). The mean allele number over loci varied between 9.9 (Schleswig-Holstein) to 4.5 (Central Asia; Table 2) due to variation in sample sizes. However, allelic richness that is corrected for differences in sample sizes, ranged at similar values across populations (4.1–5.2), except for Sardinia, where it was considerably lower than in other populations (3.5; Table 2). Based on the expected heterozygosity (H_E), the genetic variation appeared relatively high and fairly constant between the samples (0.75–0.86), excluding again Sardinia that had the lowest H_E (0.63; Table 2). Most population samples showed moderate deficit of heterozygotes as indicated by positive inbreeding coefficients (Table 2). This was the most pronounced in the three German populations from Schleswig Holstein, Harz and Saxony, but also in Norway.

Non-hierarchical AMOVA indicated divergence between populations ($H'_{ST} = 0.065$, $P < 0.0001$). Pairwise F_{ST} values were calculated (Table S2) and depicted as an UPGMA phenogram reflecting the existence of four clusters of populations (Fig. 4). There is a significantly high amount of genetic divergence between these four groups as indicated by a hierarchical AMOVA (microsatellites: $F_{CT} = 0.076$, $P < 0.001$; mtDNA: $\Phi_{CT} = 0.662$, $P < 0.001$). Two of the Mediterranean island populations exhibited the strongest divergence (Cyprus vs. all, $F_{ST} = 0.13$ – 0.28 , Sardinia vs. all, $F_{ST} = 0.10$ – 0.26 ; Table S2). For Cyprus, this was apparent not only in terms of high F_{ST} values, but also by a considerably high accumulation of private alleles despite the small sample size (Table 2). In comparison, Φ_{ST} values inferred from the mtDNA data set (control-region sequences) were much higher than F_{ST} values from the microsatellite data, but likewise indicated the strongest divergence between island and continental populations (Cyprus vs. all, $\Phi_{ST} = 0.58$ – 0.99 , Sardinia vs. all, $\Phi_{ST} = 0.41$ – 1.00).

The continental populations can be considered as two population clusters (Fig. 4). North-eastern Eurasian populations (Russia and Fennoscandia) are divergent from the central

and south-western European populations (Fig. 4). To give a general picture, pairwise F_{ST} values within these two groups were lower than values between populations of both groups (Table S2). Maximum divergence between the two continental groups was observed between southern European populations (Pyrenees, Greece) and north-eastern Palearctic populations (Russian Far East, Central Asia; $F_{ST} = 0.075$ – 0.106 ; Table S2). Lowest divergence was observed between German populations from the zone of overlap and all other continental Eurasian populations ($F_{ST} = 0.00$ – 0.088 ; Table S2).

For the entire set of populations, the ratio between male to female gene flow (m_m/m_f) was estimated based on AMOVA performed for microsatellite data ($F_{ST} = 0.065$) and mtDNA ($F_{ST(f)} = 0.667$). These estimates suggest that male gene flow contributes less to the total divergence ($F_{ST(m)} = 0.13$) than female gene flow and m_m/m_f was 13.41. When we limited the analysis to Eurasian continental populations (under exclusion of island populations) the ratio slightly increased to 19.96 with $F_{ST} = 0.029$ (microsatellites), divergence caused by female gene flow $F_{ST(f)} = 0.556$ (mtDNA) and divergence caused by male gene flow $F_{ST(m)} = 0.059$ (microsatellites).

Bayesian inference of population structure

For the complete data set ($n = 166$) under the admixture – frequency-correlated model, Evanno's ΔK separated two large clusters ($K = 2$) as the most plausible population structure (Fig. S1). There was a high level of admixture between these two groups all across Europe and only few populations appeared to be pure representatives of either of the two clusters: i) Central Asian and Far East Russian populations of the north-eastern cluster and ii) the island population from Sardinia of the south-western cluster. All continental European study populations included a high number of genetically admixed individuals. Using LOCPRIOR (classification according to mitochondrial lineages), a population structure with four genetic groups ($K = 4$) resulted as the most plausible situation (Fig. S1). Likewise for the reduced

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3 311 population dataset (n = 145 for 14 local populations) four genetic clusters (K = 4) were
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5 312 identified by the ΔK method to be the most plausible scenario, independent of the model
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7 313 applied (with and without LOCPRIOR; Figs 5, S1). The spatial differentiation pattern under
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9 314 K = 4 was as follows: In the West Palearctic, four groups were distinguished (Fig. 5): i) The
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11 315 north-eastern Palearctic cluster including Central Asian, Far East Russian and Fennoscandian
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13 316 populations, ii) the south-western Palearctic cluster including all central and southern
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15 317 continental European populations, iii) the Mediterranean cluster comprising of populations
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17 318 from Corsica and Sardinia (Fig. 5) and iv) the Cypriot population. No signs of genetic
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19 319 admixture were found in Far East Russia, on Sardinia and on Cyprus (Fig. 5). Local genetic
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21 320 admixture was found all across Europe: i) introgression of southern European alleles into
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23 321 Fennoscandian populations, ii) introgression of north-eastern Palearctic alleles into German
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25 322 populations, but (near) absence of north-eastern alleles in the Pyrenean and Greek populations
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27 323 (Fig. 5), iii) wide admixture of two southern clusters (continental: orange; Mediterranean
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29 324 islands: yellow) and introgression of continental European alleles into the Corsican island
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31 325 population (but not into the Sardinian population; Fig. 5).

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35 326 Spatial clustering studied with GENELAND identified three clusters which exactly
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37 327 matched the phylogeographic pattern of three mitochondrial lineages: a north-eastern
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39 328 Palearctic cluster (Russian Far East, Central Asia, Fennoscandia), a south-western Palearctic
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41 329 cluster (central and southern Europe, including Corsica and Sardinia) and Cyprus (Fig. 6).
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43 330 This pattern was identical in nine out of ten model runs. In a single run, Corsica and Sardinia
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45 331 together represented one separate group whereas the north-eastern and south-western coal tits
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47 332 were united in a second cluster (the number of possible clusters in the analysis was fixed at
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49 333 K = 3). Both of the two spatially explicit admixture models that were run in TESS reflected
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51 334 the same population subdivision inferred by GENELAND (Fig. 6, S2). Neither TESS nor
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53 335 GENELAND separated populations from Corsica and Sardinia from the south-western
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55 336 mainland group.
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338 Admixture rate in the hybrid zone

339 For both the mitochondrial and the microsatellite data set the best fit was for the admixture
340 model (scenario 1, Fig. 2). However, parameter estimates inferred from microsatellite data do
341 not seem reasonably interpretable for yielding extremely recent and unreliable estimates for
342 times of divergence and time since admixture (during the last six centuries). For the
343 mitochondrial data, scenario 1 was the best model with high support (posterior probabilities
344 0.9740 and 0.9983 for the direct and logistic regression approaches, respectively), also
345 supported by all the used 40 summary statistics. Type I and II errors were small; type I errors
346 were 0.006 and 0.192 and type II errors were 0.011 and 0.049 for the direct and logistic
347 approaches, respectively. Modes of the effective population sizes for Northern Europe
348 (Norway and Finland) were 280 000 (95 % HPD = 105 000–940 000), for the hybrid zone
349 9 820 000 (95 % HPD = 3 120 000–9 910 000) and for the Southern Europe 595 000 (95 %
350 HPD = 222 000–981 000). The admixture was estimated to have occurred 57 600 generations
351 ago (95 % HPD = 12 400–95 900) with an admixture rate of 0.216 (95 % HPD = 0.079–
352 0.422) relative to the northern population and divergence of southern and northern lineages 1
353 260 000 (95 % HPD = 521 000–5 710 000) generations ago. Applying a mean generation time
354 of two years, our estimates correspond to a mean time of divergence of 2.52 mya (95% HPD
355 1.04–11.42 mya) and a mean time since admixture of 0.114 mya (95% HPD 0.024–0.192
356 mya).

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358 Discussion**359 Patterns of gene flow in Europe**

360 Despite considerable genetic differences, the European range of secondary overlap among
361 south-western and north-eastern coal tit lineages does not match the general pattern of a
362 narrow and geographically restricted secondary contact zone (Fig. 1A; compare Haffer, 1989;

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3 363 Aliabadian, 2005). All German populations appeared to be genetically strongly admixed and
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5 364 introgression of south-western alleles extended northward into Fennoscandian populations,
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7 365 where corresponding (south-western) mtDNA haplotypes were absent in our study (but see
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9 366 Johnsen *et al.*, 2010). A moderate deficit of heterozygotes and linkage disequilibrium in three
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11 367 German populations (strongest in Schleswig-Holstein) are indicative of local admixture of
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13 368 two diverged genetic lineages and are typically found in populations from the centre of a
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15 369 hybrid zone (Jiggins & Mallet, 2000; Alexandrino *et al.*, 2005; Brelsford & Iwrin, 2009). The
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17 370 actual eastward and southward extent of the contact zone is far from being fully described,
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19 371 because to a lesser degree the northwestern alleles and haplotypes were present in southern
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21 372 European populations (Greece).

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24 373 Despite all limitations of our sampling, northward allelic introgression and strong
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26 374 differences between pairwise Φ_{ST} values and F_{ST} values hint to wider spatial extent of nuclear
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28 375 gene flow as compared to mtDNA introgression. Such mito-nuclear discordance, can arise
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30 376 from selection against hybrids and/or sterility of the F_1 heterogametic sex (Haldane's rule:
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32 377 Davies & Pomiankowski, 1995; Wu *et al.*, 1996) as suggested to be the case in other
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34 378 passerine hybrid zones (European *Ficedula* flycatchers: Tegelström & Gelter, 1990; Sætre *et*
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36 379 *al.*, 2001; Qvarnström *et al.*, 2010; Far East Russian great tits: Kvist & Rytkönen, 2006).
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38 380 However, in coal tits there was no evidence of hybrid sterility or selection against hybrids
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40 381 from cross-fostering experiments with individuals from European and Afghan populations
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42 382 (Löhrl, 1994). Therefore, with given certainty we can rule out selection against hybrids in
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44 383 admixed European coal tit populations, too.

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47 384 Secondly, sex-biased dispersal is considered as another possible cause of mito-nuclear
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49 385 discordance (reviewed by Prugnolle & de Meus, 2002; in birds: Kvist & Rytkönen, 2006;
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51 386 Illera *et al.*, 2011; Lin *et al.*, 2011). Though the common paradigm of female-biased dispersal
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53 387 in birds (Clarke *et al.*, 1997; Petit & Excoffier, 2009) has recently been challenged (Li &
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55 388 Merilä, 2010; Both *et al.*, 2012; Dobson, 2013), there is only very scarce information on sex-

specific dispersal distances for many bird species, including the coal tit (except Dietrich *et al.*, 2003). Because of such data deficiency further evidence from field studies is required to substantiate assumptions on any putative correlation between sex-biased dispersal and mitochondrial discordance in coal tits.

Thirdly and lastly, extreme ratios between male to female gene flow might arise from stochastic effects when comparing different levels of genetic diversity, e.g. high allelic variation of microsatellite loci and sequence variation between deeply divergent lineages (Karl *et al.*, 2014; Putman & Carbone, 2014). Due to relatively long coalescence times, incomplete lineage sorting of nuclear markers might blur spatial patterns of genetic variation. In the coal tit, this is reflected by strong admixture of two southern European allelic clusters (yellow and orange for $K = 4$) in continental populations on the one hand and a near-complete allelic lineage sorting in island populations of Corsica and Sardinia on the other hand. This is in accordance with low parameters of genetic variation on these islands and with the general assumption that density-dependent processes, such as founder effects and genetic drift, are most effective in island populations (Waters *et al.* 2013; birds: Padilla *et al.*, 2015). Even during short evolutionary time spans, fast lineage sorting derived from ancestral polymorphisms in founder populations can occur in organisms with high dispersal ability, as inferred from a comparison of historical and extant Mediterranean populations of hawkmoths (Hyles; Mende & Hundsdoerfer, 2013).

Genetic admixture on the European continent

Extant phylogeographic patterns and lineage diversification in the coal tit are likely to have emerged from glacial range fragmentation (Martens *et al.*, 2006; Pentzold *et al.*, 2013) as suggested for other tit species (Kvist *et al.*, 2003, 2005; Päckert *et al.*, 2013; Stervander *et al.*, 2015; Tritsch *et al.*, 2017). Our time estimates inferred from the mitochondrial data set support a scenario of lineage divergence close to the Pliocene-Pleistocene boundary at a mean

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time of divergence of 2.5 mya (in accordance with Päckert *et al.*, 2012). Our mean
coalescence-based estimate for time since admixture of 0.114 mya pre-dates a Holocene post-
glacial expansion and thus suggest that admixture of north-eastern and south-western gene
pools could have already started in southern refuges during the late Pleistocene. This
assumption is supported by sound evidence that northward dispersal of forest birds from
Mediterranean refuges already started before the onset of the Holocene, because fossil
remains of forest birds from interglacial periods have been found north of the Alps across
Central Europe up to a latitude of 50° N (Holm & Svenning, 2014). Furthermore, mean
coalescence time estimates are rather rough estimates, because there is no reliable empirical
value of coal-tit generation time and several authors have applied shorter generation times for
tits (Garant *et al.*, 2005; Qu *et al.*, 2015) that would shift our time since admixture estimates
closer to a Holocene expansion scenario.

The wide range of mitochondrial introgression and nuclear gene flow in central
European coal tits is indicative of a partial reversal of Pleistocene divergence patterns (for a
similar case in North American chickadees compare Manthey *et al.* 2012). The
phylogeographic pattern in continental European coal tits matches a broad trans-European
zone of intergradation at the subspecies level similarly to e.g. in Eurasian nuthatches, *Sitta*
europaea (Red'kin & Konovalova, 2006). Unlike in the latter species, phenotypical variation
of continental European coal tits is very subtle and body size parameters and plumage
coloration vary along a pan-European cline with phenotypical extreme forms *vieirae* and
abietum in the South and *ater* in the North (Wolters, 1968; Niethammer, 1943; Glutz von
Blotzheim & Bauer, 1993; Martens, 2012). Furthermore, the vocal repertoire of coal tits is
remarkably uniform throughout continental Eurasia and seems to provide a less effective
pre mating barrier compared to songs of other tit species (Thielcke, 1973; Tietze *et al.*, 2011;
Pentzold *et al.*, 2016). In contrast, in many cases of asymmetric gene flow across narrow
hybrid zones among Holarctic bird taxa (regardless of their taxonomic rank), assortative

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3 441 mating seems to be associated with strong divergence of vocal repertoires (Haavie *et al.*,
4 442 2004; Helbig *et al.*, 2001; Päckert *et al.*, 2005; Kvist & Rytönen, 2006; Sattler *et al.*, 2007;
5 443 Vorkurková *et al.*, 2013; Shipilina *et al.*, 2017). Generally, separation of gene pools is
6 444 strongly enhanced by the variation of morphological and behavioural traits that play a key
7 445 role for species recognition, e.g. in *Ficedula* flycatchers (Sætre *et al.*, 2003; Backström *et al.*,
8 446 2013; Ellegren *et al.*, 2012). In contrast, it seems that phenotypical and behavioural
9 447 differentiation between northern and southern European coal tits is too subtle to provide an
10 448 effective premating barrier.

11 449 The same holds true for potential segregation of ecological or climatic niches in
12 450 secondary contact (in tits and chickadees: Päckert *et al.*, 2005; Zhao *et al.*, 2012; Taylor *et al.*,
13 451 2014). Webb *et al.* (2011) pointed out that merging of genetic lineages might be more likely
14 452 to occur in generalist species having a lower probability of evolving unique adaptations. This
15 453 argument may apply to the coal tits as well, because despite a strong adaptation to coniferous
16 454 forests, they exploit a great variety of food resources. In those regions where the species has
17 455 adapted to deciduous forests, coal tits use a broader range of the tree's canopy and trunk than
18 456 many other parid species do (Glutz von Blotzheim & Bauer, 1993; Gosler & Clement, 2007).
19 457 Habitat structure might also have a considerable effect on local population structure, because,
20 458 in mixed conifer-broadleaved forests of Ussuriland (Far Eastern Russia), population densities
21 459 of coal tits were estimated 2.5 to 3 times higher compared to pure spruce-fir taiga forests of
22 460 the upper mountain-forest belt (Nazarenko 1984).

23 461 Third and last, there is in fact recent evidence of spatial variation in an adaptive trait of
24 462 European coal tits. Schmoll & Kleven (2011) found differences in sperm size between coal
25 463 tits from Norway and Germany, as was reported among European and Afro-Canarian blue tits
26 464 (*Cyanistes caeruleus* and *C. teneriffae*; Gohli *et al.*, 2014). Whether in blue tits these
27 465 differences would constitute an effective post-mating barrier, cannot be judged due to a lack
28 466 of range overlap in the field and missing evidence from experimental studies. In European

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3 467 coal tits, intraspecific differences in sperm morphology do not seem to effectively prevent
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5 468 gene flow across the European contact zone of coal tits.
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9 470 **Allopatric differentiation on Mediterranean Islands**

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11 471 Typically, in the central areas of a species' range, the degree of gene flow is often high,
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13 472 whereas it is low at the range margins (Kvist *et al.*, 2007; Lehtonen *et al.*, 2009; Küpper *et al.*,
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15 473 2012; Päckert *et al.*, 2013). In widespread Palearctic bird species greatest phylogeographic
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17 474 structure is often observed at the southwestern range margins, e.g. in the Mediterranean and
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19 475 on the southern European peninsulae (revision in Steward *et al.*, 2010; birds: Tietze *et al.*
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21 476 2011; Brambilla *et al.*, 2008). Since genetic drift and lineage sorting are most effective in
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23 477 small isolated populations, genetic distinctiveness of island populations is a common
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25 478 phylogeographic pattern.
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29 479 At the global scale levels of vertebrate endemism are significantly higher on islands when
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31 480 compared to the same ecoregions on the adjacent mainland (Fa & Funk, 2007; Kier *et al.*,
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33 481 2009). In the coal tit, the population from Cyprus (*cypristes*) stands out as a genetically and
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35 482 phenotypically distinctive form that dates back to a more ancient (though still Pleistocene)
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37 483 colonization (Pentzold *et al.*, 2013). Phylogenetic studies have revealed complex circum-
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39 484 Mediterranean phylogeographic patterns including distinct island lineages on Cyprus
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41 485 (Voelker & Light, 2011) and highly distinctive populations and even endemic species or
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43 486 subspecies. Apart from the famous examples of the extinct megafauna from Cyprus
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45 487 (Hadjisterkotis & Masala, 1995) weak insular endemism has also been postulated for the
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47 488 extant Cypriote herpetofauna (Böhme & Wiedel 1994) and the Cypriote avifauna (Förschler
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49 489 & Randler, 2009; Randler *et al.*, 2012).

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52 490 Genetic distinctiveness of Corsican and Sardinian coal tit populations was less
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54 491 manifest than that of *P. a. cypristes*. The shallow genetic divergence of *P. a. sardus* from its
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56 492 continental relatives (see also Tietze *et al.*, 2011; Pentzold *et al.*, 2013) contrasts the long
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3 493 evolutionary histories of some Corso-Sardinian faunal elements (reviewed in Ketmaier &
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5 494 Caccone, 2013). Accordingly, rather ancient (pre-Pleistocene) Corso-Sardinian species-level
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7 495 lineages have been found in amphibians and reptiles (Rodríguez *et al.*, 2017; Salvi *et al.*,
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9 496 2010, 2017; Fritz *et al.*, 2012). In a Corsican endemic frog species, *Discoglossus montalentii*,
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11 497 phylogeographic structure in microallopatry was found even within the island (Bisconti *et al.*,
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13 498 2013). But also in highly mobile vertebrates, such as birds, several endemic species occur on
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15 499 these islands, such as the Corsican nuthatch, *Sitta whiteheadi* (Pasquet *et al.*, 2013) and the
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17 500 Corsican finch, *Carduelis corsicana* (Förschler *et al.*, 2009), which also breeds on the
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19 501 Balearic Islands and on a few smaller neighbouring islands. In addition, there are distinct
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21 502 genetic lineages at the subspecies level in other bird species (Pons *et al.*, 2016).
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24 503 Subtle genetic admixture of the Corsican population might imply that this island does
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26 504 or did receive more influx from continental populations than the Sardinian population, e.g.
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28 505 due to its closer proximity to the mainland and along a North-South migratory pathway of
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30 506 migrants and/or dispersers. However, a greater number of local samplings from both islands
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32 507 would be needed to reliably confirm this hypothesis. Moreover, dispersal behaviour of coal
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34 508 tits is quite variable (Löhrl, 1974; Glutz von Blotzheim & Bauer, 1993; Gosler & Clement,
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36 509 2007) and seems to depend on the availability of food resources (Löhrl, 1974; Harrap &
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38 510 Quinn, 1996). On Corsica, the breeding phenology of coal tits has strongly adapted to local
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40 511 food peaks (Blondel *et al.*, 1988) and such adaptive processes might effectively have
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42 512 contributed to the fixation of genetic lineages on islands e.g. in Corsican blue tits (*Cyanistes*
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44 513 *caeruleus ogliastreae*; Porlier *et al.*, 2012). Generally, the genetic composition of the Corsican
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46 514 coal tit population might be the result of both incomplete lineage sorting during a short
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48 515 separation time (Pentzold *et al.*, 2013) and recent gene flow from irregular influx of
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50 516 continental vagrant individuals and/or dispersers.
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54 517 These examples demonstrate that the circum-Mediterranean phylogeographic pattern in
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56 518 the coal tit is partly or often paralleled in other island endemics of the Corso-Sardinian fauna.
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Traditionally, phenotypical distinctiveness has been a crucial factor for species delimitation and in fact, distinct genetic lineages of the coal tit in North Africa (*atlas* subspecies group) and on Cyprus (ssp. *cypriotes*) are corroborated by the differences in the plumage coloration (Harrap & Quinn, 1996; Gosler & Clement, 2007) and partially by subtle differences in song (Tietze *et al.*, 2011; Pentzold *et al.*, 2016). A deeper understanding of the range-wide intraspecific differentiation in the coal tit will therefore benefit (i) from an integrative taxonomic approach and (ii) from broad population sampling across gradients of genetic introgression (e.g. in narrow hybrid zones that exist for example in the Himalayas).

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998 **Figure captions**

999

1000 Figure 1: Spatial patterns of secondary contact and hybrid zones among divergent genetic
1001 lineages (schematized haplotype networks; right) of terrestrial vertebrates in the Western
1002 Palearctic; A) narrow contact zone along geographic barriers (Pyrenees, Alps) or of a wide
1003 latitudinal North-South extent; B) broad intergradation zone often along phenotypical clines;
1004 C) range wide merger and local co-occurrence of distinct genetic lineages (reversal of past
1005 lineage divergence); Europe outline map inferred from www.freeworldmaps.net.

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1007 Figure 2: Distribution range and phylogeographic pattern of the coal tit, *Periparus ater*
1008 (modified from Pentzold et al. 2013; sampling sites of mtDNA data indicated by black dots;
1009 range boundaries in light brown according to BirdLife International 2017); pie charts indicate
1010 percentages of haplotypes belonging to four different clusters (Scandinavia/Russia, W and
1011 SW Europe, North Africa and Cyprus; subspecies included in the mtDNA dataset listed at the
1012 corresponding clusters) indicated by different colours; study populations (abbreviations): BF
1013 = Black Forest, Cors = Corsica, Cyp = Cyprus, Fin = Finland, Grec = Greece, Mor =
1014 Morocco, Nor = Norway, PF = Palatine Forest, Pyr = French Pyrenees, Sard = Sardinia, Sax =
1015 Saxony, SH = Schleswig-Holstein (strongly divergent North African subspecies *atlas* not
1016 included in our population genetic study); coal tit drawing: K. Rehbinder, University of
1017 Mainz.

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1019 Figure 3: Historical models used for DIYABC analyses; pop1 = northeastern populations
1020 from Finland and Norway; pop2 = populations from the German zone of overlap (Pentzold et
1021 al. 2012); pop3 = southwestern population from the French Pyrenees.

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3 1023 Figure 4: UPGMA phenogram inferred from pairwise F_{ST} values (computed with MEGA v.6,
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5 1024 and listed in Table S2).
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9 1026 Figure 5: Genetic variation of 14 Western European and Mediterranean coal tit populations (n
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11 1027 = 145) based on 13 microsatellite loci; STRUCTURE analysis under the admixture –
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13 1028 frequency-correlated model without locpriors *a priori* defined, STRUCTURE plots for K = 2
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15 1029 to K = 4 (left); threshold $q > 0.8$ for assignment of individuals to genetic clusters (according
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17 1030 to Randi 2008) indicated for the most plausible scenario of K = 4; coloured bars above the
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19 1031 plots indicate individual assignment to three mitochondrial lineages (control region; data from
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21 1032 Pentzold et al. 2013); grey bars above indicate regional origin of samples; right: a) estimate of
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23 1033 most plausible K = 4 according to Evanno et al. (2005; ΔK plotted against the number of
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25 1034 modelled genetic clusters) and b) according to L(K) (Prichard2000); abbreviations of
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27 1035 populations: BF = Black Forest, Cor = Corsica, Cyp = Cyprus, PF = Palatine Forest, Sard =
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29 1036 Sardinia, Sax = Saxony, Schl. Holstein = Schleswig-Holstein,.
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35 1038 Figure 6: Spatial clustering of coal tit populations as inferred from GENELAND analysis;
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37 1039 assignment probabilities of the individuals to spatial clusters identified by GENELAND
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39 1040 displayed in a contour map for a) the northwestern, b) the south-western and c) the Cypriot
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41 1041 cluster respectively. The spatial membership probability is visualized by colour: bright yellow
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43 1042 indicates a high-, dark red a low assignment probability; black dots: sampling localities.
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48 1044 Figure S1: Genetic variation in the complete data set of the Western European and
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50 1045 Mediterranean coal tits (n = 166) based on 13 microsatellite loci; STRUCTURE analysis
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52 1046 under the admixture – frequency-correlated model with locpriors *a priori* defined (assignment
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54 1047 according to mtDNA lineages), STRUCTURE plots for K = 2 to K = 4 (left); right: estimate
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56 1048 of most plausible K according to Evanno et al. (2005; ΔK plotted against the number of
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3 1049 modelled genetic clusters) a) under the admixture – frequency-correlated model without
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5 1050 locpriors *a priori* defined ($K = 2$) and b) with locpriors *a priori* defined ($K = 4$); assignment
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7 1051 according to mtDNA lineages, control region: coloured bars above the plots.
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11 1053 Figure S2: Spatial clusters as inferred by the spatial explicit CAR admixture model of TESS
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13 1054 ($K_{\max} = 3$). The TESS admixture models did not distinguish more than three units even if the
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15 1055 number of possible clusters in model is higher ($K_{\max} > 3$). The DIC criterion (arithmetic
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17 1056 mean of ten replicates) confirmed three spatial clusters.
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Tables

Table 1: Characteristics and variation of nine newly identified microsatellite loci for population genetic analysis of coal tits. The Parate loci were amplified using the protocol described by Schuelke (2000). For this protocol an 18-bp long M13(-21) tail (5'-TGTAAACGACGGCCAGT-3') was adhered to the 5' end of the forward primer (= Label); * based on sequence clone; ** based on experimental optimization during microsatellite primer design; *** minimum and maximum sizes based on analyses of 14 populations.

GenBank accession numbers to be provided upon acceptance.

newly designed primers					
Locus	Primer sequence 5'-3'	GenBank accession no	Repeat Motif*	T _m (°C)**	Allele Size (bp)***
Parate 01	F: (Label) TCCTGGAGCACATTATGTCATATG	MG970333	(TAGA) ₁₄ (CAGA) ₄	56	209 - 271
	R: AATCTGCTGCTCCATACCTTGG				
Parate 02	F: (Label)-AGGGACAGAAATTGTGCAAGG	MG970334	(ATCT) ₁₄	56	189 - 373
	R: TGCATTTCATGCATACATAGACAC				
Parate 03	F: Label-TGTGTCTGCAAAAGGCCAAG	MG970335	(ATAG) ₁₄	56	117 - 177
	R: CAAAGCCTTCATCTGCTTGG				
Parate 06	F: (Label)-TTCAGTGCAGGTGCATAAATTG	MG970336	(CTAT) ₁₅	56	219 - 355
	R: GGCCAAAGAGAAAGTAGGGGTGTAG				
Parate 07	F: (Label)-CTCCCAAGAGAGTCTGTGTCTG	MG970337	(CTAT) ₁₂	56	166 - 199
	R: AAGGCTTTTGAAACAGGAGAAAG				
Parate 08	F: (Label)-TTGTAAACGACCTTGCACCTC	MG970338	(CA) ₂₀	50	90 - 153
	R: AGGCAGTAAACCCCTCAT TGG				
Parate 09	F: (Label)-GGCACAGATGCATATTTTGTTCAC	MG970339	(GT) ₁₃	56	122 - 136
	R: TGCACAAATCATGCTTAATCCTC				
Parate 15	F: (Label)-TCACAAAAAGGCAATTGTCAG	MG970340	(TC) ₁₂ (C) ₄ (TC) ₇ CC(TC) ₄	56	129 - 204
	R: GGAGACAGGAGAGCAGCAAC				
Parate 16	F: (Label)-CTTTCTTGAATGCTCAGATTGC	MG970342	(CT) ₂₇	56	166 - 263

	R: CAAGCCCATGTTCAAGGTTC				
primers from previous studies					
Pat2-43	F: ACAGGTAAGTCAGAAAATGGAAG				
	R: (Label)-GTATCCACAGAGTCTTTTGCTGATG			60	126 - 213
PmaTGAn33	F: (Label)-TTCCCCAAAGTATCCCTGCATC				
	R: AAACCATATCACCCAGTGCC	AY260539 Saladin <i>et al.</i> (2003)	(GATA)14GAT(GATA)8	57	258 - 398
Pma69u	F: (Label)-CCCAGACAAAAGCATCACTGG				
	R: GACAGTTCACATAGCCCTGG	AB094107 Kawano (2003)	(TG)6	57	214 – 222
PmaC25	F: CGTCCTGCTGTTTGTAATTCTG				
	R: (Label)-CCATGAACCATTTTATAGGGTG	AY260526 (Saladin <i>et al.</i> , 2003)	(GAT)11	57	313 – 349

Table 2: Diversity parameters and inbreeding coefficient (F_{IS}) for the surveyed population samples of in total 145 individuals. Loci Parate 06 and Parate 08 were excluded due to the presence of null alleles. # = population number, n = number of sampled individuals, HO = observed heterozygosity, HE = expected heterozygosity, * Far East = pooled samples from Far East Russia (n = 11) and Japan (n = 2); ** Schleswig-Holstein = pooled samples from Itzehoe (n = 15), Amrum (n = 3) and Brilit (n = 1).

#	Population	n	Mean number of alleles per locus	Mean allelic richness	Private alleles	Mean HO	Mean HE	F_{IS}	p-value (F_{IS})
1	Central Asia	5	4.455	4.089	1	0.777	0.753	-0.036	0.692
2	Far East*	13	7.818	4.717	5	0.825	0.814	-0.015	0.688
3	Finland	11	8.455	4.993	10	0.807	0.832	0.031	0.204
4	Norway	13	8.545	4.992	7	0.769	0.839	0.087	0.016
5	Schleswig-Holstein**	19	9.909	5.026	7	0.740	0.833	0.115	0.0003
6	Harz	12	8.273	4.886	2	0.758	0.822	0.082	0.018
7	Saxony	5	5.364	4.770	4	0.755	0.824	0.094	0.080
8	Palatine Forest	7	7.364	5.237	3	0.831	0.848	0.022	0.373
9	Black Forest	8	7.364	5.175	3	0.849	0.860	0.014	0.336
10	Pyrenees	13	8.636	4.803	4	0.816	0.819	0.003	0.473
11	Greece	12	7.909	4.718	14	0.760	0.807	0.062	0.084
12	Corsica	10	6.273	4.249	6	0.755	0.765	0.015	0.414
13	Sardinia	10	4.818	3.462	1	0.605	0.630	0.042	0.240
14	Cyprus	7	5.182	4.059	13	0.786	0.771	-0.020	0.683

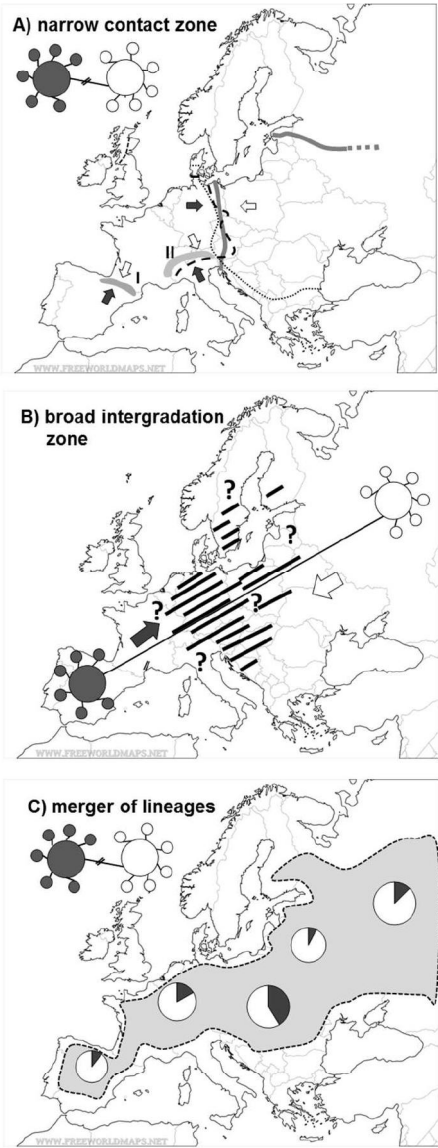


Figure 1

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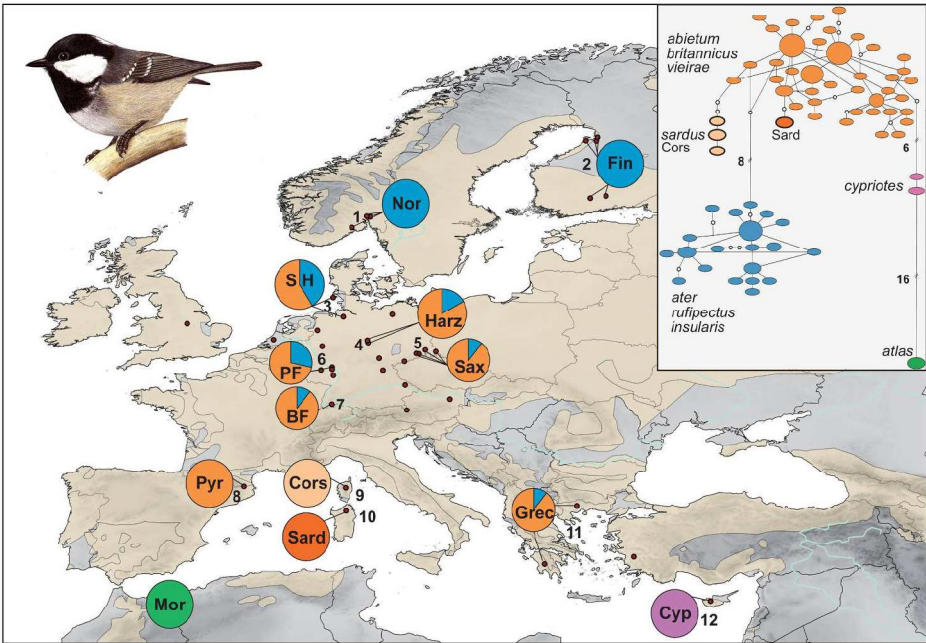
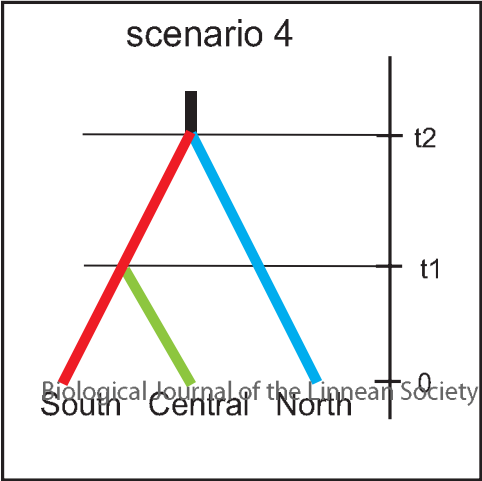
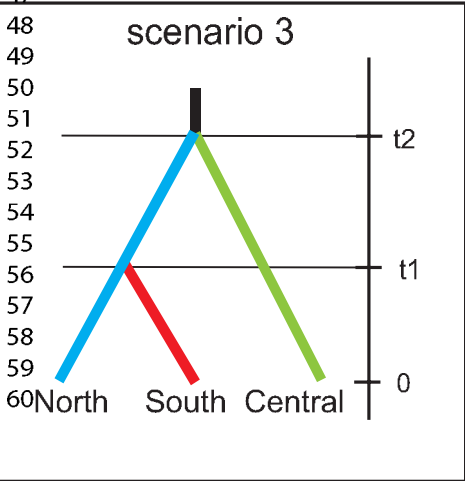
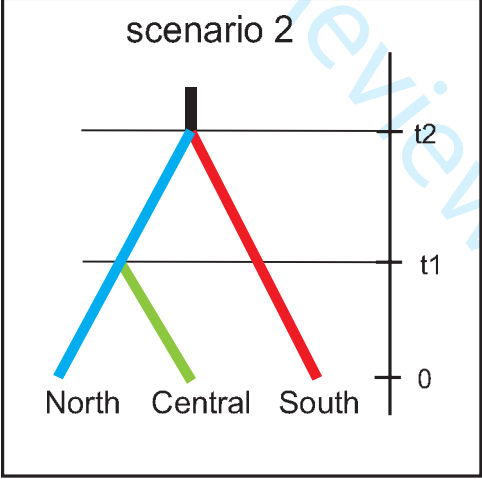
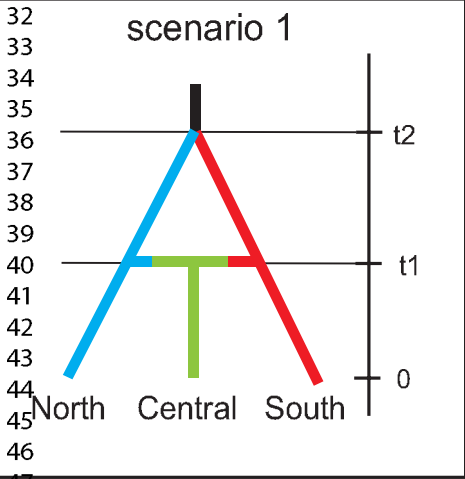


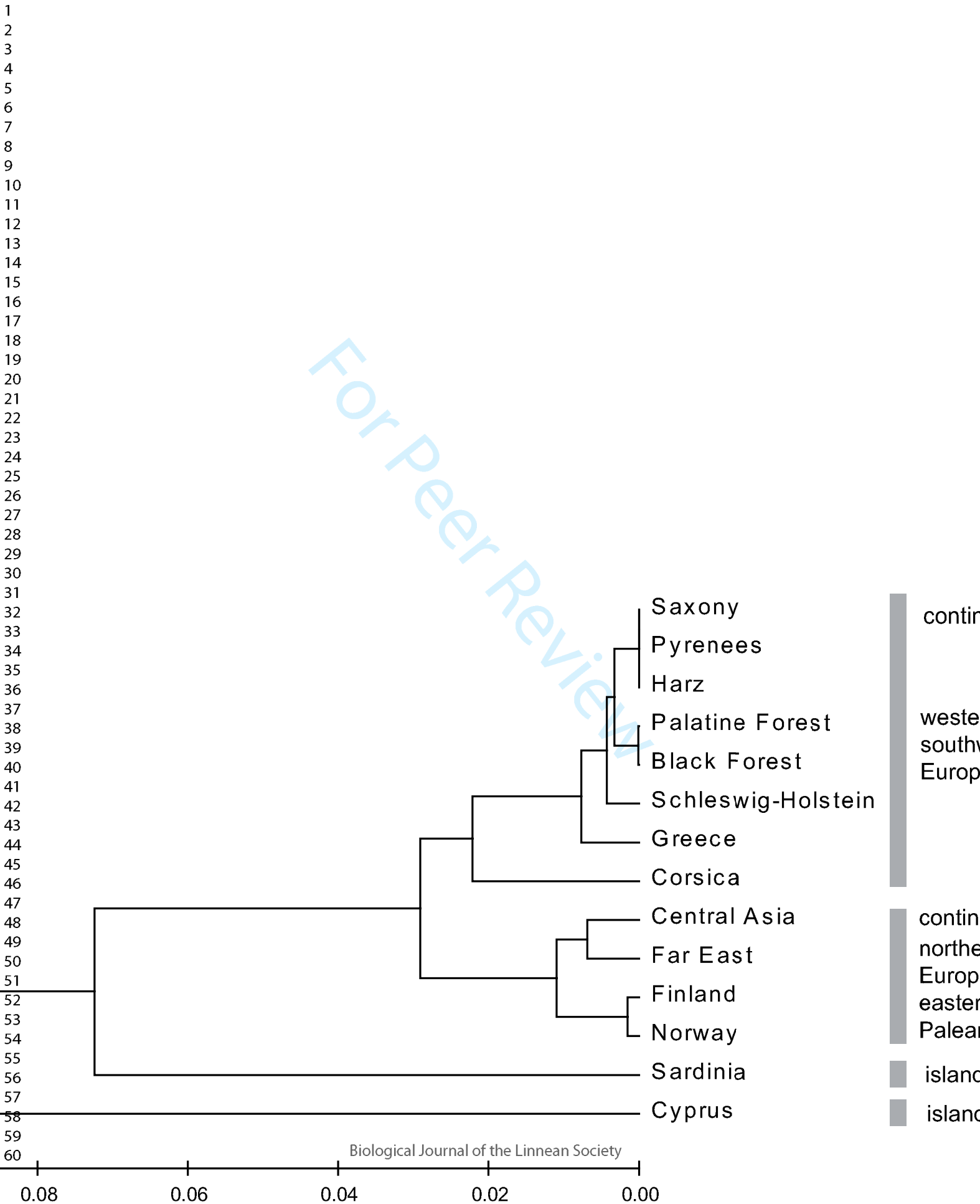
Figure 2

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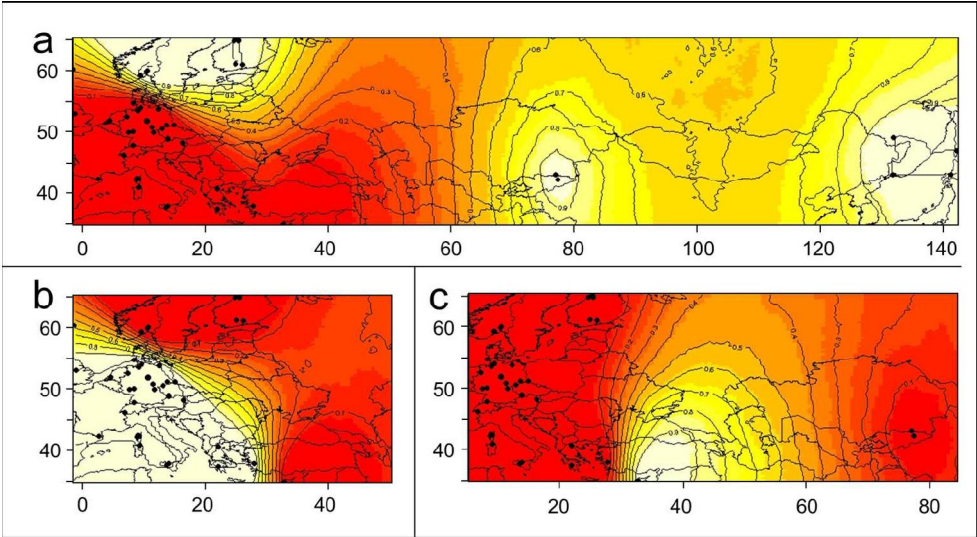
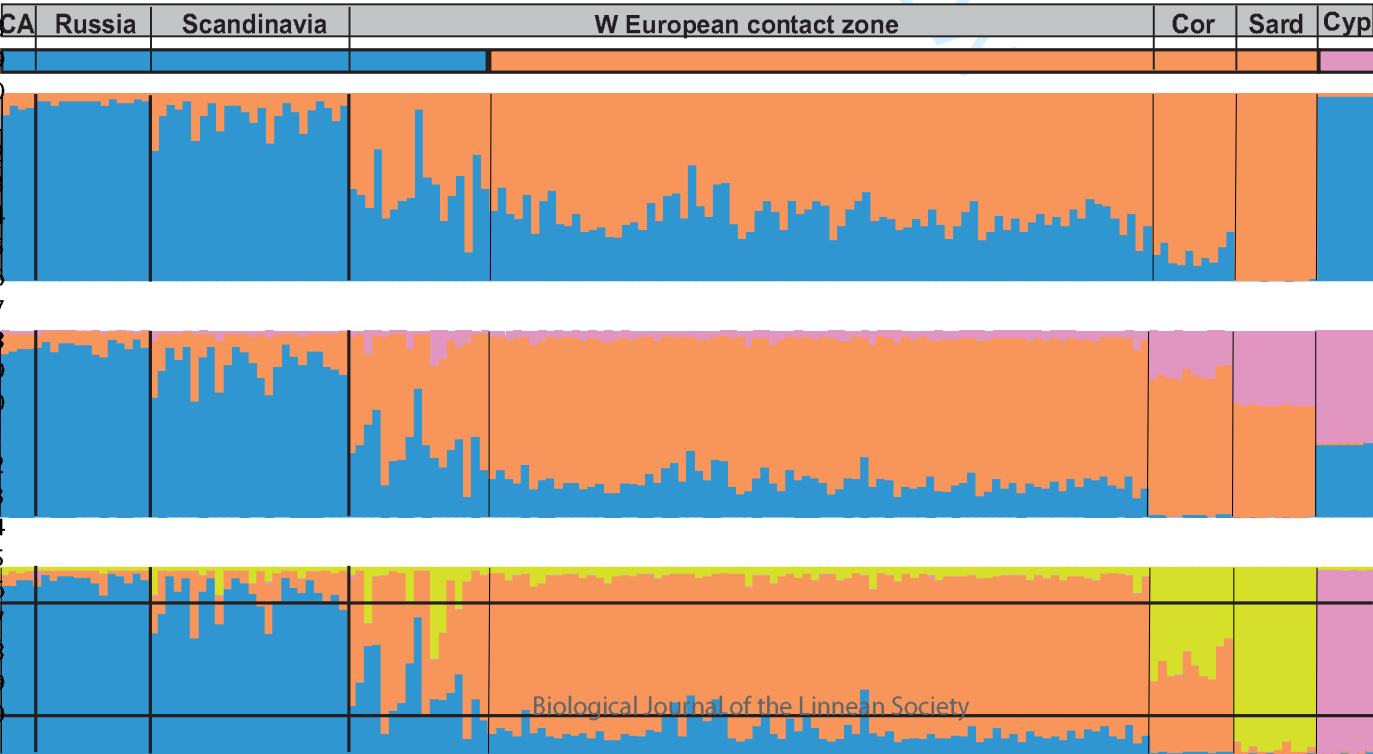


Figure 6

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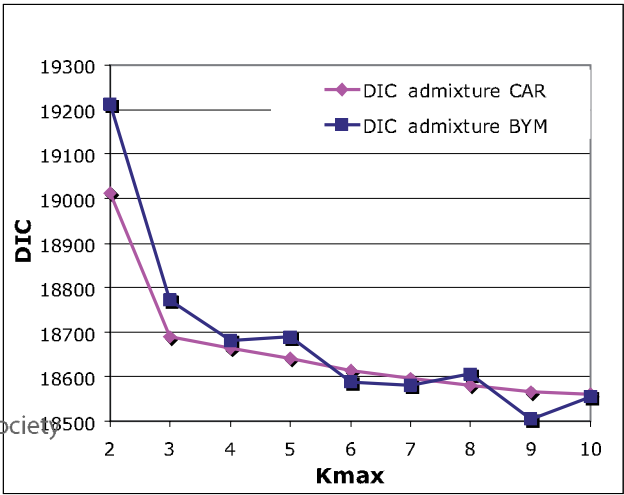
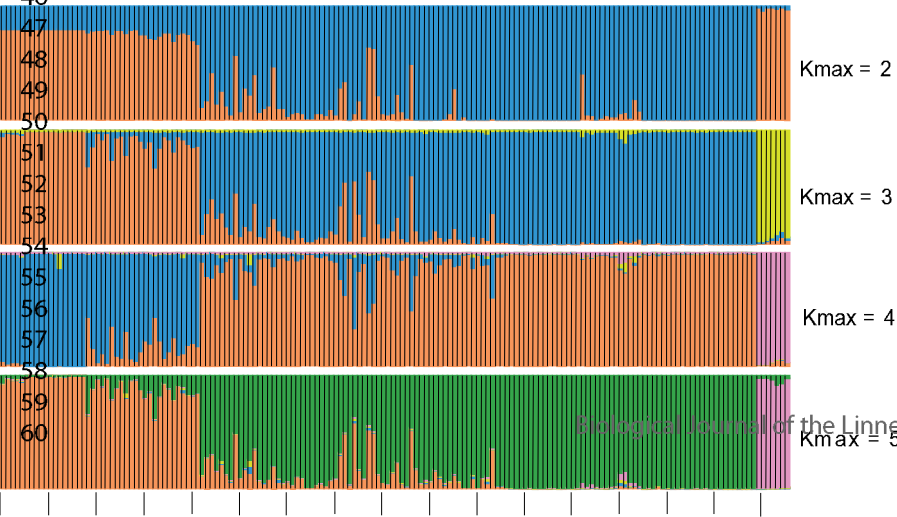


Table S1: Observed and expected heterozygosity, departure from Hardy-Weinberg equilibria, evidence of null alleles and linkage disequilibrium for each locus and population

Bonferroni corrected p-value for HWE = $0.05 / 13 = \mathbf{0.0038}$; Bonferroni corrected p-value for Linkage Disequilibrium = $0.05/78 = \mathbf{0.00064}$

Locus	Parate 01	Parate 02	Parate 03	Parate 06	Parate 07	Parate 08	Parate 09	Parate 15	Parate 16	Pat2-43	PmaTG An33	Pma69	PmaC25
Central Asia (N = 5)													
H _O	1.000	0.800	0.800	0.460	1.000	0.750	0.200	0.800	0.800	0.800	0.750	0.800	0.800
H _E	0.844	0.844	0.822	0.755	0.844	0.893	0.467	0.600	0.867	0.733	0.857	0.622	0.777
HWE	1.000	0.779	0.890	0.028	1.000	0.428	0.333	1.000	0.734	1.000	0.699	1.000	1.000
p-value													
Null allele bias	no	no	no	no	no	no	no	no	no	no	no	no	no
Linkage disequilibrium	no												

Far East (N = 13)													
H _O	0.692	1.000	0.923	0.461	0.846	0.385	0.769	0.615	0.923	0.923	0.769	0.692	0.923
H _E	0.778	0.932	0.840	0.791	0.849	0.898	0.775	0.757	0.895	0.834	0.834	0.665	0.791
HWE	0.641	0.4998	0.929	0.0004	0.2795	< 0.001	0.436	0.298	0.126	0.705	0.087	1.000	1.000
p-value													
Null allele bias	no	no	no	yes	no	yes	no	no	no	no	no	no	no
Linkage disequilibrium	no												

Locus	Parate 01	Parate 02	Parate 03	Parate 06	Parate 07	Parate 08	Parate 09	Parate 15	Parate 16	Pat2-43	PmaTG An33	Pma69	PmaC25
Finland (N = 11)													
H _O	0.909	0.818	0.909	0.455	1.000	0.545	0.636	0.636	0.800	0.800	0.909	0.636	0.818
H _E	0.896	0.931	0.866	0.857	0.823	0.922	0.645	0.757	0.821	0.884	0.913	0.688	0.922
HWE	0.079	0.491	0.896	0.002	0.597	< 0.001	0.333	0.2795	0.412	0.525	0.901	0.342	0.532
p-value													
Null allele bias	no	no	no	yes	no	yes	no	no	no	no	no	no	no
Linkage disequilibrium	no												

Norway (N = 13)													
H _O	0.769	0.846	0.846	0.769	0.923	0.846	0.692	0.692	0.769	0.923	0.769	0.462	0.769
H _E	0.775	0.898	0.853	0.745	0.843	0.935	0.751	0.877	0.902	0.902	0.883	0.649	0.898
HWE	0.875	0.177	0.171	0.152	0.594	0.057	0.466	0.189	0.009	0.697	0.1397	0.358	< 0.001
p-value													
Null allele bias	no	no	no	no	no	no	no	no	no	no	no	no	no
Linkage disequilibrium	no												

Locus	Parate 01	Parate 02	Parate 03	Parate 06	Parate 07	Parate 08	Parate 09	Parate 15	Parate 16	Pat2-43	PmaTG An33	Pma69	PmaC25
Schleswig-Holstein (N = 19)													
H _O	0.789	0.857	0.722	0.500	0.556	0.500	0.526	0.722	0.583	1.000	0.944	0.667	0.778
H _E	0.771	0.939	0.8396	0.846	0.806	0.867	0.644	0.900	0.873	0.924	0.878	0.722	0.867
HWE	0.617	0.104	0.160	0.0025	0.144	0.055	0.012	<0.001	<0.001	0.969	0.991	0.822	0.4395
P-value													
Null allele bias	no	no	no	yes	yes	yes	no	yes	yes	no	no	no	no
Linkage disequilibrium	Parate01, Parate 03, Parate06, Parate 07, Parate12, Parate15												

Harz (N = 12)													
H _O	0.750	0.583	0.917	0.500	0.917	0.583	0.667	0.750	0.667	1.000	0.750	0.583	0.750
H _E	0.786	0.942	0.902	0.833	0.819	0.895	0.746	0.812	0.819	0.948	0.8695	0.670	0.728
HWE	0.665	<0.001	0.212	0.093	0.753	<0.001	0.339	0.498	0.077	1.000	0.189	0.347	0.615
p-value													
Null allele bias	no	yes	no	yes	no	yes	no	no	no	no	no	no	no
Linkage disequilibrium													

Locus	Parate 01	Parate 02	Parate 03	Parate 06	Parate 07	Parate 08	Parate 09	Parate 15	Parate 16	Pat2-43	PmaTG An33	Pma69	PmaC25
Saxony (N = 5)													
H _O	0.800	0.800	1.000	0.800	0.600	0.800	0.400	0.800	0.800	1.00	0.750	0.750	0.600
H _E	0.756	0.933	0.844	0.933	0.800	0.889	0.711	0.844	0.867	0.889	0.857	0.714	0.844
HWE	1.000	0.227	1.000	0.256	0.4998	0.57989	0.144	0.362	0.621	1.000	0.657	1.000	0.386
p-value													
Null allele bias	no	no	no	no	no	no	no	no	no	no	no	no	no
Linkage equilibrium	no												

Locus	Parate 01	Parate 02	Parate 03	Parate 06	Parate 07	Parate 08	Parate 09	Parate 15	Parate 16	Pat2-43	PmaTG An33	Pma69	PmaC25
Palatine Forest (N = 7)													
H _O	0.857	0.857	0.857	0.286	0.714	0.429	0.714	1.000	0.857	0.857	0.857	0.857	0.714
H _E	0.857	0.923	0.868	0.846	0.659	0.945	0.703	0.890	0.923	0.923	0.956	0.780	0.846
HWE	0.928	0.513	0.829	0.001	0.838	< 0.001	0.623	0.793	0.592	0.5099	0.224	1.000	0.6099
p-value													
Null allele bias	no	no	no	yes	no	yes	no	no	no	no	no	no	no
Linkage disequilibrium	no												

Black Forest (N = 8)													
H _O	1.000	1.000	1.000	0.500	1.000	0.625	1.000	0.625	0.750	0.714	0.750	0.875	0.625
H _E	0.875	0.942	0.875	0.833	0.833	0.883	0.775	0.842	0.892	0.923	0.875	0.792	0.842
HWE	1.000	1.000	0.102	0.051	0.887	0.102	0.366	0.089	0.350	0.089	0.583	1.000	0.083
p-value													
Null allele bias	no	no	no	yes	no	no	no	no	no	no	no	no	no
Linkage disequilibrium	no												

Locus	Parate 01	Parate 02	Parate 03	Parate 06	Parate 07	Parate 08	Parate 09	Parate 15	Parate 16	Pat2-43	PmaTG An33	Pma69	PmaC25
Pyrenees (N = 13)													
H _O	0.846	0.923	0.692	0.538	0.846	0.538	0.846	0.833	0.615	1.000	0.833	0.846	0.692
H _E	0.840	0.957	0.908	0.815	0.769	0.892	0.785	0.743	0.837	0.938	0.855	0.702	0.671
HWE	0.660	0.359	0.157	0.064	0.973	< 0.001	0.646	0.713	0.157	1.000	0.413	0.942	0.067
p-value													
Null allele bias	no	no	yes	yes	no	yes	no	no	no	no	no	no	no
Linkage disequilibrium	no												

Greece (N = 12)													
H _O	0.833	0.818	0.583	0.750	0.583	0.750	1.000	0.818	0.600	0.909	0.912	0.545	0.750
H _E	0.884	0.9697	0.833	0.743	0.692	0.917	0.747	0.848	0.733	0.896	0.906	0.610	0.754
HWE	0.305	0.145	0.230	0.331	0.264	0.006	0.042	0.811	1.000	0.156	0.637	0.762	0.897
p-value													
Null allele bias	no	no	no	no	no	no	no	no	no	no	no	no	no
Linkage disequilibrium	no												

Locus	Parate 01	Parate 02	Parate 03	Parate 06	Parate 07	Parate 08	Parate 09	Parate 15	Parate 16	Pat2-43	PmaTG An33	Pma69	PmaC25
Corsica (N = 10)													
H _O	0.700	0.900	0.900	0.400	0.500	0.300	0.800	0.600	0.900	0.900	0.900	0.400	0.800
H _E	0.795	0.821	0.832	0.868	0.768	0.884	0.679	0.742	0.926	0.811	0.884	0.584	0.574
HWE	0.305	0.239	0.311	< 0.001	0.068	< 0.001	0.813	0.320	0.794	0.984	0.984	0.141	0.198
p-value													
Null allele bias	no	no	no	yes	no	yes	no	no	no	no	no	no	no
Linkage disequilibrium	Parate 03, Parate 07												

Sardinia (N = 10)													
H _O	0.800	0.900	0.600	0.800	0.778	0.600	0.200	0.333	0.800	0.556	0.889	0.500	0.300
H _E	0.900	0.795	0.684	0.816	0.752	0.737	0.189	0.582	0.621	0.712	0.791	0.426	0.479
HWE	0.667	0.623	0.341	0.751	0.255	0.165	1.000	0.137	1.000	0.203	0.010	1.000	0.4799
p-value													
Null allele bias	no	no	no	no	no	no	no	no	no	no	no	no	no
Linkage disequilibrium													

Locus	Parate 01	Parate 02	Parate 03	Parate 06	Parate 07	Parate 08	Parate 09	Parate 15	Parate 16	Pat2-43	PmaTG An33	Pma69	PmaC25
Cyprus (N = 7)													
H _O	1.000	1.000	0.714	0.429	0.714	0.000	0.571	0.429	0.714	1.000	0.857	mono.*	0.857
H _E	0.846	0.934	0.879	0.901	0.802	0.615	0.4395	0.582	0.714	0.934	0.912	mono.*	0.670
HWE	0.707	1.000	0.504	0.0003	0.4099	0.002	1.000	0.637	1.000	1.000	0.636	-	0.412
p-value													
Null allele bias	no	no	no	yes	no	yes	no	no	no	no	no	no	no
Linkage disequilibrium	Parate 01, Pat2-43												

*: mono. = monomorphic

Table S2: Matrix of pairwise F_{ST} values between all local populations (computed with Arlequin 3.1). Bold: significant F_{ST} values after Bonferroni correction for the number of comparisons at the level $p = 0.05$. (Bonferroni correction alpha (α) = 0.05 / for 91 comparisons $\alpha = 0.0005$).

	C Asia	Russia	Finland	Norway	SH	Harz	Saxony	P Forest	B Forest	Pyrenees	Greece	Corsica	Sardinia
Central Asia	-												
Russian Far East	0.014	-											
Finland	0.010	0.010	-										
Norway	0.046	0.021	0.003	-									
Schleswig-Holstein	0.088	0.059	0.038	0.018	-								
Harz	0.072	0.064	0.036	0.034	0.010	-							
Saxony	0.067	0.044	0.021	0.034	0.004	-0.007	-						
Palatine Forest	0.039	0.043	0.024	0.024	0.010	0.007	-0.014	-					
Black Forest	0.067	0.042	0.028	0.019	0.009	0.001	-0.003	0.0002	-				
Pyrenees	0.095	0.075	0.050	0.047	0.010	-0.004	-0.004	0.009	0.022	-			
Greece	0.106	0.082	0.059	0.037	0.026	0.011	0.017	0.011	0.014	0.013	-		
Corsica	0.137	0.129	0.095	0.090	0.048	0.036	0.033	0.051	0.056	0.033	0.055	-	
Sardinia	0.256	0.215	0.179	0.155	0.134	0.102	0.156	0.148	0.122	0.101	0.106	0.064	-
Cyprus	0.172	0.143	0.130	0.116	0.172	0.176	0.187	0.156	0.145	0.179	0.158	0.218	0.284