

1 **Maternal and infant prediction of the child BMI trajectories; studies across two generations of Northern**
2 **Finland birth cohorts**

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23 **Abstract:**

24 **Background/objective:** Children BMI is a longitudinal phenotype, developing through interplays between
25 genetic and environmental factors. Whilst childhood obesity is escalating, we require a better understanding
26 of its early origins and variation across generations to prevent it.

27 **Subjects/Methods:** We designed a cross-cohort study including 12,040 Finnish children from the Northern
28 Finland Birth Cohorts 1966 and 1986 (NFBC1966 and NFBC1986) born before or at the start of the obesity
29 epidemic. We used group-based trajectory modelling to identify BMI trajectories from 2 to 20 years. We
30 subsequently tested their associations with early determinants (mother and child) and the possible
31 difference between generations, adjusted for relevant biological and socioeconomic confounders.

32 **Results:** We identified four BMI trajectories, 'stable-low' (34.8%), 'normal' (44.0%), 'stable-high' (17.5%) and
33 'early-increase' (3.7%). The 'early-increase' trajectory represented the highest risk for obesity. We analysed
34 a dose-response association of maternal pre-pregnancy BMI and smoking with BMI trajectories. The
35 directions of effect were consistent across generations and the effect sizes tended to increase from earlier
36 generation to later. Respectively for NFBC1966 and NFBC1986, the adjusted risk ratios of being in the early-
37 increase group were 1.08 (1.06-1.10) and 1.12 (1.09-1.15) per unit of pre-pregnancy BMI and 1.44 (1.05-1.96)
38 and 1.48 (1.17-1.87) in offspring of smoking mothers compared to non-smokers. We observed similar
39 relations with infant factors including birthweight for gestational age and peak weight velocity. In contrast,
40 the age at adiposity peak in infancy was associated with the BMI trajectories in NFBC1966 but did not
41 replicate in NFBC1986.

42 **Conclusion:** Exposures to adverse maternal predictors were associated with a higher risk obesity trajectory
43 and were consistent across generations. However, we found a discordant association for the timing of
44 adiposity peak over a 20-year period. This suggest the role of residual environmental factors, such as
45 nutrition, and warrants additional research to understand the underlying gene-environment interplay.

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51 **Key words:** Epidemiology, life-course, BMI trajectories, adiposity peak, obesity

52 INTRODUCTION

53 In 2016, the World Health Organisation (WHO) reported a global prevalence of 41 million children under the
54 age of five and 340 million between 5 and 19 years old with overweight or obesity ¹. Over the last four
55 decades, obesity in the latter age category has increased tenfold ². Obesity, including during childhood, is
56 currently being explained by interplays between a polygenetic build-up ^{3,4} and interrelated environmental
57 risk factors acting upon body composition (epidemiologically proxied from BMI). Whilst the polygenetic
58 structure has remained stable over the last four decades, including in the population studied in the current
59 research ⁵, we are observing major changes in the set of environmental risk factors affecting the risk of
60 obesity, with growing evidence to support the role of the environment in the early years of life. At the
61 population level, BMI and changes in BMI during childhood are therefore highly heterogeneous and different
62 paths to obesity may coexist throughout the life-course, some of them starting very early and influencing
63 child's growth patterns ⁶⁻⁸.

64 So far, a vast majority of studies have considered childhood BMI from a cross-sectional manner without
65 accounting for the longitudinal effect such as changes over time. Various analytical strategies have been
66 proposed to study child BMI development. Conventional growth modelling methods assume that a single
67 growth trajectory approximates the entire population and that covariates influence each individual in the
68 same manner ⁹. However, these models cannot account for the heterogeneity of BMI development over
69 longer periods ^{10,11}. In contrast, in growth mixture models (GMM), the attention is put on relationships among
70 individuals and the longitudinal characteristics of the measures. They assume that the population is made of
71 latent groups, sharing the same pattern over time. GMM is a flexible modelling approach that can provide
72 quantitative insights in the longitudinal aspect of BMI changes throughout the life course.

73 Previous longitudinal studies have described multiple BMI trajectories defined from latent growth trajectory
74 analysis supporting more than one BMI trajectories in childhood ¹²⁻¹⁵. As to whether early biological and/or
75 psychosocial factors, classically associated with the child BMI ¹⁶, influenced each BMI trajectory in a
76 comparative manner and whether the strength of such association is affected from one generation to
77 another remains a debatable area of research. In the current study, we hypothesized that maternal and early
78 childhood factors are determinants of the BMI trajectory a child embarks. We further hypothesized that the
79 strength (*i.e.* effect size) of the associations linking a risk factor to the child BMI trajectory could be modified
80 from one generation to another. Testing such hypotheses might help identify potential shared and
81 generation specific risk factors and further advance the understanding of child BMI development.

82 To study the association of early risk factors with childhood BMI trajectories and its evolution over time, we
83 performed a BMI trajectory analysis in two birth cohorts about one generation apart, using GBTM (Group
84 Based Trajectory Modelling) in SAS PROC TRAJ ¹⁰. Importantly, we studied the Northern Finland Birth Cohort

85 (NFBC)1966, pre-dating the obesity epidemic and the NFBC1986, born 20 years later, with prospective
86 recruitment at the start of the obesity epidemic in Finland. While consistent evidence supports replicability
87 of effects of early life factors on the child BMI and the risk of obesity¹⁷, we may anticipate important
88 generational effects depending on contextual differences in terms of feeding and nurturing practices and
89 changes in the environmental exposures.

90

91 **METHODS**

92 **Study population**

93 The study was based on the two Northern Finland Birth Cohorts initiated 20 years apart from the same region
94 (the two northernmost provinces of Finland: Oulu and Lapland) and founder population. NFBC1966 recruited
95 pregnant women with a due date between the 1st of January and the 31st of December 1966 (12 055 mothers,
96 12 231 babies, 96.3% of all births from this period in the region). NFBC1986 included pregnant women with
97 an expected delivery date between 1st of July 1985 and 30th of June 1986 (9362 mothers, 9479 babies, 99%
98 of all births from this period in the region). A total of 12 058 and 9432 babies were born alive in the NFBC1966
99 and 1986 respectively.

100 **Data collection of the child BMI measures**

101 The mothers entered the study around the 16th gestational week for NBC1966 and the 10th to 12th gestational
102 weeks for NFBC1986. Pregnant mothers were followed throughout pregnancy. Children's height and weight
103 measures were collected by linking data from questionnaires, Health and Welfare records, clinical
104 examination and national registers. Briefly, in Finland, a child welfare nurse checks up on infants every month
105 during the first few months and then once a year, usually around their birthday. When they start school at
106 seven years of age, the school nurse takes over the yearly check-ups. Children's measurements data were
107 completed by self-reported measurements at 14 years for NFBC1966 and 7-8 years for NFBC1986. At 16 years
108 old, NFBC1986 members were invited to a clinical examination with a trained nurse.

109 **Exclusion criteria**

110 We excluded preterm babies (<37 gestational weeks) and multiple births (N=2 199) (Supplementary Figure
111 S1). We calculated BMI (kg/m²) from height and weight. It is recommended and customary to use weight-
112 for-length rather than BMI to measure growth during the first two years of life¹⁸. Therefore, we chose to
113 model BMI in 16 age-windows, one per year from two to 16.9 years-old and, due to the scarcity of the data
114 in late adolescence, the last group comprised ages from 17 to 20. Individuals with less than three repeated
115 BMI measurements, required for model stability, were excluded (N=7 159 altogether in NFBCs). Attrition in
116 the data could be due either to non-attendance to check-ups or to non-retrieved or lost records. There were

117 little differences between the children included from the model and those excluded: the mothers of the
118 excluded children were more often single and less educated than the mothers of the included children
119 (Supplementary table S1).

120 The study comprised 12 040 individuals (51.9% male), 6864 from NFBC1966 (53.7% male) and 5176 from
121 NFBC1986 (49.4% male).

122 **Maternal data**

123 In the pregnancy questionnaire, mothers were asked to report their pre-pregnancy weight. The
124 corresponding BMI was calculated using height measurement from the first prenatal visit. Maternal
125 education was categorised as elementary, vocational or secondary, matriculation level, matriculation and
126 beyond. Maternal smoking status at eight weeks of pregnancy was categorized as smokers and no smokers
127 and their marital status as married or cohabitating, single, widowed or divorced.

128 **Infant and child data**

129 The first rise in the BMI curve called adiposity peak (AP), around nine months of age and the nadir of the
130 curve called adiposity rebound (AR) around six years were obtained from random effects models fitted from
131 0 to 18 months and from 18 months to 13 years as described elsewhere¹⁹. The peak height velocity (PHV)
132 and peak weight velocity (PWV) in infancy were derived from parametric growth curves^{20,21}. These methods
133 are developed in the supplementary materials. Gestational age at birth was calculated from the mother's last
134 menstrual period in NFBC1966, and ultrasound in priority before the last menstrual period in NFBC1986.
135 Birthweight was adjusted for gestational age, sex and cohort. BMI measures were sex and cohort
136 standardized, for each age-windows as described previously^{3,4}. Briefly, the z-scores were calculated
137 internally, derived from NFBC1966 and NFBC1986, comprising respectively 96% and 99% of all children born
138 in the north of Finland at the time.

139 **Statistical analysis**

140 **Developmental Growth trajectories**

141 We used GBTM from the PROC TRAJ procedure in SAS software, based on Nagin's approach to group-based
142 modelling²², to determine BMI growth trajectories. This approach consists in fitting a mixture of parametric
143 models to the data, using the maximum likelihood method, handling missing data under the missing at
144 random assumption. PROC TRAJ fits longitudinal data to discern rather than assume two or more distinctive
145 trajectories or latent groups of individuals and estimate their prevalence in the population. Group-based
146 trajectory modelling has been successfully applied to BMI in large cohort studies^{13,14,23}.

147 We used PROC TRAJ with BMI modelled as censored normal to identify subgroups which share a similar BMI
148 growth from two to 20 years in both cohorts modelled together. We started by fitting a quartic polynomial

149 to the models, increasing one by one the number of groups up to 7. The first step of the modelling consisted
150 in choosing the optimal model. Bayesian Information Criteria (BIC) values are compared and the model
151 with the smallest absolute value is chosen. The second step consisted in determining the shape of each
152 trajectory by identifying the best polynomial order. Model adequacy was evaluated by calculating the
153 average posterior probabilities of belonging to each group, they should be at least 0.7 with the closest to 1
154 reflecting the best discrimination between groups. We also calculated the Odds of Correct Classification, they
155 should be over 5, and sought for a close correspondence between the estimated and actual percentages for
156 each group and a reasonably tight confidence intervals of each trajectory¹⁰. A group should represent at least
157 1% of the population⁹. Parsimony and a priori knowledge of the topic combined with an evaluation of the
158 graphical shape of the trajectories should be taken into account.

159 **Descriptive and association analysis**

160 All analyses were performed using SAS software 9.4 (SAS Institute Inc, Cary, North Carolina). Characteristics
161 of participants were presented as frequencies for categorical variables and means and SD for continuous
162 variables. We used non-parametric tests for comparisons between groups and χ^2 test for categorical
163 variables. Multivariate models were applied to calculate risk ratios (RR) and their 95% confidence intervals in
164 associations between maternal, infancy and childhood predictors and BMI trajectories, trajectory two being
165 the reference. The confounders for the models were selected according to existing literature and previous
166 analyses¹⁹. We tested three models for the maternal parameters: unadjusted, adjusted for maternal
167 education and adjusted for parity, maternal education, maternal age and smoking (Fig. 2a)/pre-pregnancy
168 BMI (Fig. 2b). Birthweight was not included in model 2a-b as it might be in the causal pathway between
169 prenatal factors and the BMI trajectories. For birthweight, we tested three models: unadjusted, adjusted for
170 pre-pregnancy BMI and adjusted for parity, pre-pregnancy BMI, maternal age, maternal education, smoking.
171 For childhood predictors, we tested four models: unadjusted, adjusted for birthweight z-scores, adjusted for
172 pre-pregnancy BMI and the last model adjusted for parity, pre-pregnancy BMI, maternal age, maternal
173 education, smoking and birthweight z-scores.

174

175 **RESULTS**

176 **Characteristics of the study population**

177 On average, mothers enrolled in the NFBC1986 cohort were 3 cm taller, 2 months younger and had a 0.9 BMI
178 smaller in comparison to NFBC1966 mothers (Supplementary Table S2). In addition, they also had fewer
179 children, were more often smoker and better educated. Furthermore, offspring born to the NFBC1986
180 mothers were 86 grams heavier, had a 3.7% increase in PHV and a 2.3% decrease in PWV compared to their

181 NFBC1966 counterparts. They were also younger at the time of AP by 3.6 weeks and AR by 7.7 months
182 suggesting distinct early growth patterns.

183 **BMI trajectories**

184 In an exploratory analysis, we modelled the trajectories separately by cohort (Supplementary Fig. S2 and
185 table S3). To compare the cohort effect, we modelled the pooled cohorts, controlling for sex in the modelling
186 process. The model converged well using the default starting values (codes in Supplementary material).
187 During the modelling of the trajectories, we could not identify the best number of trajectories based on the
188 Bayesian Index Criterion (BIC), i.e. BIC continued decreasing through all seven tested models. In this situation
189 Nagin advised to use more subjective criteria¹⁰. Between the visual analysis of the trajectory graphs, the
190 objective of the study and the knowledge of the other goodness of fit criteria (Supplementary Fig. S3 and
191 table S4), we were able to identify four trajectories (polynomials 4, 3, 4, 3) as the optimal model for the
192 studied population (Fig. 1). The trajectories were named according to their position in the graph (low to high)
193 as 1: 'stable-low' (34.8% of total population, 34.4% of NFBC1966 and 35.4% of NFBC1986), 2: 'normal' (44.0%,
194 44.9% of NFBC1966 and 42.9% of NFBC1986), 3: 'stable-high' (17.5%, 13.3% of NFBC1966 and 17.7% of
195 NFBC1986), 4: 'early-increase' (3.7%, 3.4% of NFBC1966 and 4.0% of NFBC1986). We observed that,
196 compared to the other trajectories, the early-increase trajectory started already at a higher point, with a
197 steeper curve. The cohort and sex prevalence per group are presented in Supplementary table S5.

198 In both cohorts, from trajectory one to four, we observed a stepwise increase in pre-pregnancy BMI and
199 smoking during pregnancy (Table 1). The effect was consistent between cohorts both in terms of direction
200 and magnitude. Parity and maternal marital status were associated to trajectories in a stepwise manner in
201 NFBC1966 only. We also noted a decrease in the proportion of non-instrumental vaginal deliveries from
202 group one to four in both cohorts.

203 The early life determinants followed the same stepwise pattern described earlier (Table 2). Birthweight was
204 on average higher in NFBC1986 for each group trajectory and increased gradually by 245 and 270g between
205 trajectory one to four in NFBC1966 and NFBC1986 respectively. We observed the same trend in PHV with
206 higher PHV in NFBC1986 and an increase of 1.1 cm/year from group one to four in NFBC1966 only. PWV
207 increased by 2.13 and 1.99 kg/year, for NFBC1966 and NFBC1986 respectively. At AP, we observed a stepwise
208 increase in BMI between trajectory one and four, 1.2 and 0.8 kg/m² in NFBC1966 and NFBC1986 respectively.
209 Changes in age at AP showed the same trend with only one week increase for NFBC1966. AR occurred earlier
210 in NFBC1986 than in NFBC1966 and we observed a dramatic stepwise decrease of 2.6 and 2.9 years between
211 the low stable and the early increase trajectory for NFBC1966 and NFBC1986 respectively. It is interesting to
212 notice that in trajectory four, BMI at AP and AR are high, above 18 kg/m² in both cohorts.

213 **Associations between maternal factors and BMI trajectories**

214 We observed differences between trajectories, showing positive associations of pre-pregnancy BMI and
215 maternal smoking with stable high and early increase trajectories (Fig. 2). In Figure 2a, in the fully adjusted
216 model compared to the normal trajectory, pre-pregnancy BMI was associated to a lower risk of belonging to
217 the stable-low trajectory in both NFBC1966 and NFBC1986. From high-stable to early-increase trajectories,
218 the risk steadily increased up to a RR of 1.08 (95% CI 1.06-1.10) for NFBC1966 and 1.12 (95% CI 1.09-1.15) for
219 NFBC1986. Maternal smoking was associated to a 34% (adjusted-RR (aRR): 1.34, 95% CI 1.14-1.59) and 42%
220 (aRR: 1.42, 95% CI 1.22-1.64) higher risk of being in the high-stable trajectory for NFBC1966 and NFBC1986
221 respectively (Fig. 2b). The risk increased up to 44% (aRR: 1.44, 95% CI 1.05-1.96) and 48% (aRR: 1.48, 95% CI
222 1.17-1.87) in the early-increase trajectory for NFBC1966 and NFBC1986 respectively.

223 **Associations between early life factors and BMI trajectories**

224 The association between BW z-score and BMI trajectories (Fig. 3a) showed, in both cohorts, the stepwise
225 pattern described in Fig. 2. A higher BW z-score was associated to a higher risk of belonging to the highest
226 trajectory, with an aRR of 1.18 (95% CI 1.05-1.33) in NFBC1966 and 1.28 (95% CI: 1.12-1.46) in NFBC1986 (Fig.
227 3a). PWV (Fig. 3b) was associated with a decreased risk of belonging to the low-stable trajectory (aRR: 0.92,
228 95% CI 0.92-0.93 in NFBC1966 and aRR: 0.89, 95% CI: 0.88-0.90 in NFBC1986). High PWV was associated with
229 an 8% and 7% higher risk of being in the early-increase trajectory, for NFBC1966 and NFBC1986 respectively.
230 Regarding PHV, the pattern was divergent between the cohorts (Fig. 3c). PHV was positively associated with
231 trajectory four (aRR:1.014, 95% CI: 1.002-1.026) in NFBC1966, but no association was found in NFBC1986.
232 The stepwise pattern described earlier was maintained in the association between age at AP and the four
233 trajectories in NFBC1966. The adjusted RR were ranging from 0.989 (95% CI: 0.986-0.991) in the lowest
234 trajectory to 1.015 (95% CI: 1.008-1.022) in the early-increase trajectory (Fig. 3d). However, there was no
235 association in NFBC1986.

236

237 **DISCUSSION**

238 To our knowledge, this is the first time that BMI latent growth trajectories were modelled in two birth cohorts
239 set 20 years apart. The specific study design relied on using two separate prospective birth cohorts from the
240 same founder population born 20 years apart. Whilst the causal genomic factors affecting BMI development
241 are highly likely to be stable from one generation to another, we have observed important changes in the
242 environmental risk factors associated to BMI development from the mid 60's and mid 80's in Europe. We
243 identified four BMI z-scores trajectories from 2 to 20 years in the combined NFBC studies. Our findings
244 suggested that, in both cohorts, offspring of high pre-pregnancy BMI or smoking mothers had more chances
245 of belonging to the more adverse childhood BMI trajectory. Conversely, children of low pre-pregnancy BMI
246 or non-smoking mothers had better chances in following more favourable trajectories. Our results also

247 suggested that the child's faster weight and height gains in infancy associated with adverse trajectories in
248 both cohorts. Furthermore, we uncovered that age at AP was associated with every BMI trajectory in
249 NFBC1966, but the association was lost in NFBC1986, suggesting that adiposity measured around 9 months
250 of age might not be a stable determinant of later adiposity during childhood.

251 Our findings about the association of adverse maternal factors and BMI trajectories were in line with others
252 ^{15,25}. The age of adiposity rebound decreased in a stepwise manner from group trajectories one to four,
253 consistent with a higher risk of obesity. We observed the same pattern between the two cohorts which
254 suggests that the effect associated to the variation of pre-pregnancy BMI and smoking remained over 20
255 years. They are important factors to consider in future generations and cohort studies. Interestingly, we
256 observed that the amplitude of the effect differed. Although the average pre-pregnancy BMI was lower in
257 NFBC1986 than in NFBC1966, its effect on the trajectories appeared stronger in the early increase trajectory.
258 The association of birthweight with the trajectories was reflecting the strong link shown with pre-pregnancy
259 BMI, unchanged over a 20-year period. A Danish study showed a stable association between birthweight and
260 childhood overweight across almost 50 years ²⁶, those results were supported by a study comparing both
261 NFBC studies ²⁷. Infancy peak velocities occurred around the first month of life and were associated with BMI
262 trajectories in both cohorts. Birthweight as an indicator of foetal growth and the peak velocities as indicators
263 of early postnatal growth could be expected to be associated with the trajectories. However, there are still
264 some important areas of debate pertained by mismatched findings between epidemiological observations
265 and the causal inference made by Mendelian randomisation ^{28,29} or the measures of genetic overlap between
266 these early adiposity phenotypes ³⁰. The relationship between these early growth phenotypes (BW, PWV and
267 PHV) and the child BMI trajectory from 2 years onwards may still need clarifications as highlighted by our
268 present observations. Possible inter-individuals and -generational differences in childcare and early nutrition
269 may be important sources of moderation of the above relationships. These inter-individual and generational
270 factors might explain the large confidence intervals and differential effect size or the lack of replication as
271 these observed for PHV.

272 One of the main findings of this study was that the age at AP followed a different pattern of association
273 between the two cohorts with a stepwise association observed between the age at AP and BMI trajectories
274 in NFBC1966 only. The current literature in the field shows contrasting findings. Evidence from Swedish and
275 Dutch birth cohorts supported a positive association between the BMI at adiposity peak and the later risk of
276 obesity ^{31,32}. However, the generalisation of such association is currently being debated by two recent GWAS
277 supporting distinct molecular factors regulating infant and child BMI ^{30,33}. Furthermore, in western
278 populations, it was reported that BMI at adiposity peak is getting lower in more contemporary cohorts. This
279 may seem counterintuitive with the increasing prevalence of childhood obesity during the last decades but
280 it aligned with a study based on European cohorts establishing that children from contemporary cohorts had

281 a lower BMI at two years, a greater BMI growth velocity and earlier age at AR than children from older cohorts
282 ³⁴. Altogether, these findings warrant a better understanding of the nature of the association linking the BMI
283 of a child in infancy and during childhood to support evidence-based recommendations for parents and
284 health care professionals. The present observation for an inconsistent effect between two generations of
285 birth cohort from the same founder population may allow us to speculate about an indirect (confounded),
286 association between the timing of the adiposity peak and the risk of being in an adverse BMI trajectory.

287 Although, we are lacking quantitative or qualitative indicators to explain this generational difference in the
288 age at AP and PHV in infancy, it was meaningful, and it might highlight important moderating factors. The
289 Finnish society underwent a massive change between 1966 and 1986, from agricultural to high-tech society.
290 This transformation was accompanied with better pre- and post-natal care, but it also brought convenient
291 energy-dense food affecting both adult and children nutrition. We might speculate that the differences in
292 PHV and AP between cohorts, observed so soon after birth, might be due to early nutrition. Although,
293 breastfeeding data in NFBC was incomplete, when included in the model (data not shown), it did not alter
294 the results. The changes in PHV and AP between the cohorts might also indicate some residual confounding
295 that we were unable to analyse. Historically, breastfeeding in Finland, like in other European countries,
296 decreased from the second World War to its record low in the 1960s and 1970s until it started to increase
297 again ^{35,36}. Following this trend, it is likely that NFBC1966 infants were less often breastfed than the
298 NFBC1986. Exclusive breastfeeding for five months has been shown to modulate the timing of AP and AR and
299 BMI velocities in Avon Longitudinal Study of Parents And Children study ³⁷. Exclusive breastfeeding for six
300 months reduced the associations of birthweight and early weight gain on fat mass in three year old children
301 in a Danish cohort ³⁸.

302 **Strengths and limitations**

303 One of the great strengths of this study was the richness of data, we were able to closely follow any variation
304 in childhood BMI through 16 age-windows. Another highly valuable strength resided in the use of two birth
305 cohorts, born before and at the start of the obesity epidemic, originating from the same geographic area of
306 Finland and characterised by a genetically homogeneous population. Nevertheless, limitations should be
307 considered. One limitation of this study would reside in the harmonization of variables between cohorts,
308 such as paternal data, type of infant feeding or maternal weight gain during pregnancy which could not be
309 reciprocated in both cohorts. Due to model requirements, many individuals from both cohorts were
310 excluded. In addition, we should acknowledge that BMI measures the ratio between weight and height. Each
311 of these two measures are susceptible to describe their own trajectories during childhood which might affect
312 the BMI trajectories described in this report. One of the future steps to undertake, to grow our understanding
313 of the biological and environmental mechanisms being at play would be the development of analytical
314 strategy modelling child height and weight trajectory simultaneously.

315 There are few statistical methods available to model children BMI. Modelling approaches using mixed-effect
316 models, latent curve analysis, hierarchical modelling or growth-curve modelling are offering measures of
317 individual growth profiles against the mean and may provide a health professional with derived phenotypes
318 such as the age at adiposity rebound. In contrast, the GBTM used in this study, is a person-centred data-
319 driven process and assumes that the population is composed of latent groups, they do not assume the one-
320 size-fits-all approach. These subgroups are homogeneous within their trajectory but distinct from other
321 trajectories, each following the same behaviour over time. This latent approach captures more information,
322 especially the longitudinal relationships at the child level but, is specific to the modelled population. Unlike
323 growth models, GBTM are limited in the obtention of distinct phenotypes that could be directly translated
324 as clinical measures. Nonetheless, they seem to present a new set of tools to study individual variation in
325 response to clinical interventions and randomized trials³⁹.

326 **Conclusion**

327 Our results add new insights to the study of childhood obesity by using two generations of Finnish birth
328 cohorts, initiated before and at the start of the obesity epidemic. In both cohorts, detrimental maternal
329 factors were associated to adverse BMI trajectories, independent of time. However, we were observing a
330 larger amplitude of the effects in the younger cohort suggesting moderation by a more obesogenic
331 environment. Our findings support evidence for very early mechanisms in the first months of life linked to
332 childhood obesity and affected over the course of a generation. Finally, the cross-cohort design exemplified
333 by this research might be a powerful way to detect indirect associations such as the one linking early variation
334 at the time of the adiposity peak and later BMI trajectories. Further research, and methodological
335 development are warranted to identify the intergenerational changes that might help revealing gene-
336 environment interplays.

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338

339 **Data Availability Statement:** Data is available from the Northern Finland Birth Cohort (NFBC) for
340 researchers who meet the criteria for accessing confidential data. Please, contact NFBC project center
341 (NFBCprojectcenter@oulu.fi) and visit the cohort website (www.oulu.fi/nfbc) for more information.

342 **Ethical approval:** All procedures performed were in accordance with the 1964 Helsinki declaration. The Ethics
343 Committee of the Northern Ostrobothnia Hospital District has approved the NFBC1966 and NFBC1986
344 studies.

345 **Informed consent:** Mothers gave their informed consent in the beginning of the NFBC1966 and 1986 data
346 collections. Written informed consent has been obtained from the cohort participants in the 31- and 46-year
347 data collections.

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357 **Conflict of Interest**

358 The authors, RN, JM, MM, MRJ and SS declare that they have no competing interest.

359

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462 **LEGENDS.**

463

464 **Table 1.** Characteristic of maternal variables according to the group trajectories. BMI: Body Mass Index.
465 Statistical tests performed between trajectories by cohorts and between cohorts by trajectories.

466 **Table 2.** Characteristics of infancy and childhood variables according to the group trajectories. PHV: Peak
467 Height Velocity, PWV: Peak Weight Velocity, AP: Adiposity Peak, AR: Adiposity Rebound. Statistical tests
468 performed between trajectories by cohorts and between cohorts by trajectories.

469 **Fig. 1:** BMI z-scores trajectories of NFBC studies from 2 to 20 years. Solid lines represent the trajectories and
470 dashed lines the 95% confidence intervals. Group trajectories, 1: Stable-low (34.8%, N=4195), 2: Normal
471 (44.0%, N=5299), 3: Stable-high (17.5%, N=2106) ND 4: Early-increase (3.7%, N=440).

472 **Fig. 2:** Forest-plot of unadjusted and adjusted Risk Ratios (RR) between maternal parameters and BMI z-
473 scores trajectory classes. Fig. 2a: pre-pregnancy BMI (kg/m^2), Fig. 2b: maternal smoking during pregnancy.
474 RR with 95% confidence intervals. ●: unadjusted; ▲: adjusted for maternal education; ■: adjusted for pre-
475 pregnancy BMI (Fig. 2b); ▼: adjusted for parity, maternal age, maternal education, maternal smoking (Fig.
476 2a) / maternal age, parity, maternal education and pre-pregnancy BMI (Fig. 2b). G1: Stable Low group
477 trajectory; G2: Normal group trajectory (reference); G3; Stable high group trajectory; G4; Early Increase group
478 trajectory.

479 **Fig. 3:** Forest-plot of unadjusted and adjusted Risk Ratios (RR) between early growth parameters and BMI z-
480 scores trajectory classes. Fig. 3a: birthweight z-scores, Fig. 3b: peak weight velocity in infancy (kg/year), Fig.
481 3c: peak height velocity in infancy (cm/year), Fig. 3d: age at adiposity peak (years). RR with 95% confidence
482 intervals. ●: unadjusted; ▲: adjusted for birthweight z-score (Fig. 3b, 3c and 3d); ■: adjusted for pre-
483 pregnancy BMI; ▼: adjusted for parity, pre-pregnancy BMI, maternal age, maternal education, maternal
484 smoking (Fig. 3a) / birthweight z-score, parity, pre-pregnancy BMI, maternal age, maternal education,
485 maternal smoking (Fig. 3b, 3c and 3d). G1: Stable Low group trajectory; G2: Normal group trajectory
486 (reference); G3; Stable high group trajectory; G4; Early Increase group trajectory.

Table 1

| | Trajectories | | | | | | | | <i>p</i> value |
|--------------------------|---------------|-------------------|-----------|-------------------|----------------|-------------------|-------------------|-------------------|-----------------------|
| | 1: Stable-low | | 2: Normal | | 3: Stable-high | | 4: Early-increase | | |
| | N=4195 | | N=5299 | | N=2106 | | N=440 | | |
| | N | mean±SD | N | mean±SD | N | mean±SD | N | mean±SD | |
| Maternal pre-pregnancy | | | | | | | | | |
| weight (kg) | | | | | | | | | |
| NFBC1966 | 2250 | 57.8 ± 8.2 | 2933 | 59.9 ± 8.9 | 1136 | 61.4 ± 9.8 | 223 | 63.6 ± 10.1 | <0.0001 |
| NFBC1986 | 1801 | 57.2 ± 8.5 | 2187 | 59.5 ± 9.1 | 895 | 62.4 ± 10.3 | 201 | 66.6 ± 14.0 | <0.0001 |
| <i>p</i> value | | 0.0187 | | 0.0440 | | 0.775 | | 0.325 | |
| Maternal pre-pregnancy | | | | | | | | | |
| BMI (kg/m ²) | | | | | | | | | |
| NFBC1966 | 2171 | 22.5 ± 3.0 | 2814 | 23.3 ± 3.2 | 1091 | 24.2 ± 3.5 | 215 | 25.1 ± 3.8 | <0.0001 |
| NFBC1986 | 1795 | 21.5 ± 3.0 | 2175 | 22.3 ± 3.2 | 891 | 23.3 ± 3.6 | 201 | 25.1 ± 5.1 | <0.0001 |
| <i>p</i> value | | <0.0001 | | <0.0001 | | <0.0001 | | 0.295 | |
| Parity | | | | | | | | | |
| NFBC1966 | 2358 | 3.0 ± 2.2 | 3078 | 3.0 ± 2.2 | 1188 | 2.8 ± 2.1 | 234 | 2.6 ± 2.1 | 0.0006 |
| NFBC1986 | 1828 | 1.6 ± 2.0 | 2212 | 1.5 ± 1.9 | 914 | 1.3 ± 1.7 | 206 | 1.4 ± 1.7 | 0.1 |
| <i>p</i> value | | <0.0001 | | <0.0001 | | <0.0001 | | <0.0001 | |
| | N | % | N | % | N | % | N | % | <i>p</i> value |
| Maternal Smoking | | | | | | | | | |
| NFBC1966 | 2321 | | 3008 | | 1167 | | 230 | | 0.0038 |
| Smoker | 309 | 13.3 | 388 | 12.9 | 194 | 16.6 | 41 | 17.8 | |
| No smoker | 2012 | 86.7 | 2620 | 87.1 | 973 | 83.4 | 189 | 82.2 | |
| NFBC1986 | 1827 | | 2208 | | 912 | | 205 | | <0.0001 |
| Smoker | 302 | 16.5 | 400 | 18.1 | 236 | 25.9 | 62 | 30.2 | |
| No smoker | 1525 | 83.5 | 1808 | 81.9 | 676 | 74.1 | 143 | 69.8 | |
| <i>p</i> value | | 0.0037 | | <0.0001 | | <0.0001 | | 0.0024 | |
| Maternal marital status | | | | | | | | | |
| NFBC1966 | 2358 | | 3078 | | 1185 | | 234 | | 0.0419 |
| Married/Cohabiting | 2291 | 97.1 | 2995 | 97.3 | 1139 | 96.1 | 222 | 94.9 | |
| Single | 56 | 2.4 | 71 | 2.3 | 35 | 3.0 | 8 | 3.4 | |
| Widowed /Divorced | 11 | 0.5 | 12 | 0.4 | 11 | 0.9 | 4 | 1.7 | |
| NFBC1986 | 1831 | | 2213 | | 917 | | 206 | | 0.13 |
| Married/Cohabiting | 1750 | 95.6 | 2124 | 96.0 | 881 | 96.1 | 194 | 94.2 | |
| Single | 56 | 3.0 | 73 | 3.3 | 32 | 3.5 | 8 | 3.9 | |
| Widowed /Divorced | 25 | 1.4 | 16 | 0.7 | 4 | 0.4 | 4 | 1.9 | |

| p value | 0.0028 | | 0.0225 | | 0.33 | | 0.95 | | |
|-------------------------------------|-------------------|------|-------------------|------|---------------|------|---------------|------|-------------------|
| Operative delivery | | | | | | | | | |
| NFBC1966 | 839 | | 1087 | | 441 | | 97 | | 0.0009 |
| Non-instrumental vaginal deliveries | 642 | 76.5 | 822 | 75.6 | 323 | 73.2 | 59 | 60.8 | |
| Caesarian Section | 90 | 10.7 | 131 | 12.1 | 54 | 12.3 | 26 | 26.8 | |
| Others (vacuum extraction, forceps) | 107 | 12.8 | 134 | 12.3 | 64 | 14.5 | 12 | 12.4 | |
| NFBC1986 | 1834 | | 2218 | | 918 | | 206 | | <0.0001 |
| Non-instrumental vaginal deliveries | 1545 | 84.2 | 1828 | 82.4 | 714 | 77.8 | 158 | 76.7 | |
| Caesarian Section | 177 | 9.7 | 278 | 12.5 | 133 | 14.5 | 31 | 15.1 | |
| Others (vacuum extraction, forceps) | 112 | 6.1 | 112 | 5.1 | 71 | 7.7 | 17 | 8.3 | |
| p value | <0.0001 | | <0.0001 | | 0.0004 | | 0.0155 | | |

Table 2

| | Trajectories | | | | | | | | <i>p</i> value |
|-----------------------------|--------------|-------------------|--------|-------------------|-------------|-------------------|----------------|-------------------|-------------------|
| | 1 | | 2 | | 3 | | 4 | | |
| | Stable-low | | Normal | | Stable-high | | Early-increase | | |
| | N=4195 | | N=5299 | | N=2106 | | N=440 | | |
| | N | % | N | % | N | % | N | % | |
| Sex (%male) | | | | | | | | | |
| NFBC1966 | 2361 | 51.8 | 3081 | 56.2 | 1188 | 50.6 | 234 | 54.7 | 0.0010 |
| NFBC1986 | 1834 | 48.7 | 2218 | 51.4 | 918 | 45.7 | 206 | 50.5 | 0.0286 |
| <i>p</i> value | | 0.0458 | | 0.0005 | | 0.0276 | | 0.38 | |
| | N | mean±SD | N | mean±SD | N | mean±SD | N | mean±SD | <i>p</i> value |
| Birthweight (grams) | | | | | | | | | |
| NFBC1966 | 2361 | 3450 ± 467 | 3081 | 3568 ± 481 | 1188 | 3642 ± 520 | 234 | 3695 ± 465 | <0.0001 |
| NFBC1986 | 1834 | 3519 ± 440 | 2218 | 3662 ± 460 | 918 | 3739 ± 479 | 206 | 3789 ± 517 | <0.0001 |
| <i>p</i> value | | 0.0001 | | <0.0001 | | <0.0001 | | 0.0369 | |
| Birthweight z-score | | | | | | | | | |
| NFBC1966 | 2361 | -0.14 ± 0.94 | 3081 | 0.08 ± 0.95 | 1188 | 0.25 ± 1.03 | 234 | 0.35 ± 0.94 | <0.0001 |
| NFBC1986 | 1834 | -0.19 ± 0.93 | 2218 | 0.10 ± 0.96 | 918 | 0.27 ± 1.00 | 206 | 0.36 ± 1.11 | <0.0001 |
| <i>p</i> value | | 0.063 | | 0.49 | | 0.46 | | 0.93 | |
| PHV in Infancy (cm/year) | | | | | | | | | |
| NFBC1966 | 2081 | 50.31 ± 3.72 | 2726 | 50.67 ± 3.71 | 1069 | 50.92 ± 3.90 | 209 | 51.41 ± 4.06 | <0.0001 |
| NFBC1986 | 1785 | 52.42 ± 6.72 | 2150 | 52.49 ± 6.74 | 893 | 52.61 ± 6.83 | 199 | 52.28 ± 6.54 | 0.95 |
| <i>p</i> value | | <0.0001 | | <0.0001 | | <0.0001 | | 0.2705 | |
| PWV in infancy (kg/year) | | | | | | | | | |
| NFBC1966 | 2145 | 12.03 ± 1.43 | 2806 | 13.07 ± 1.56 | 1096 | 13.65 ± 1.89 | 213 | 14.16 ± 2.07 | <0.0001 |
| NFBC1986 | 1798 | 11.65 ± 2.39 | 2167 | 13.01 ± 2.85 | 898 | 13.42 ± 3.05 | 200 | 13.64 ± 3.40 | <0.0001 |
| <i>p</i> value | | <0.0001 | | <0.0001 | | <0.0001 | | 0.0029 | |
| Age AP (years) | | | | | | | | | |
| NFBC1966 | 1817 | 0.75 ± 0.03 | 2427 | 0.76 ± 0.03 | 959 | 0.77 ± 0.04 | 188 | 0.77 ± 0.04 | <0.0001 |
| NFBC1986 | 1753 | 0.70 ± 0.02 | 2114 | 0.69 ± 0.02 | 882 | 0.69 ± 0.02 | 195 | 0.69 ± 0.02 | <0.0001 |
| <i>p</i> value | | <0.0001 | | <0.0001 | | <0.0001 | | <0.0001 | |
| BMI AP (kg/m ²) | | | | | | | | | |
| NFBC1966 | 1817 | 17.5 ± 0.7 | 2427 | 18.1 ± 0.7 | 959 | 18.4 ± 0.8 | 188 | 18.7 ± 0.9 | <0.0001 |
| NFBC1986 | 1753 | 17.2 ± 0.6 | 2114 | 17.7 ± 0.6 | 882 | 17.9 ± 0.7 | 195 | 18.0 ± 0.8 | <0.0001 |

| <i>p value</i> | <0.0001 | | <0.0001 | | <0.0001 | | <0.0001 | | |
|-----------------------------|---------|-------------|---------|-------------|---------|-------------|---------|-------------|---------|
| Age AR (years) | | | | | | | | | |
| NFBC1966 | 2323 | 6.20 ± 0.59 | 3033 | 5.67 ± 0.60 | 1168 | 4.79 ± 0.71 | 231 | 3.59 ± 0.71 | <0.0001 |
| NFBC1986 | 1822 | 5.72 ± 0.72 | 2197 | 5.02 ± 0.73 | 911 | 3.92 ± 0.71 | 200 | 2.83 ± 0.42 | <0.0001 |
| <i>p value</i> | <0.0001 | | <0.0001 | | <0.0001 | | <0.0001 | | |
| BMI AR (kg/m ²) | | | | | | | | | |
| NFBC1966 | 2323 | 14.4 ± 0.5 | 3033 | 15.5 ± 0.4 | 1168 | 16.6 ± 0.6 | 231 | 18.1 ± 1.1 | <0.0001 |
| NFBC1986 | 1822 | 14.6 ± 0.5 | 2197 | 15.8 ± 0.4 | 911 | 16.95 ± 0.6 | 200 | 18.3 ± 1.1 | <0.0001 |
| <i>p value</i> | <0.0001 | | <0.0001 | | <0.0001 | | 0.0503 | | |

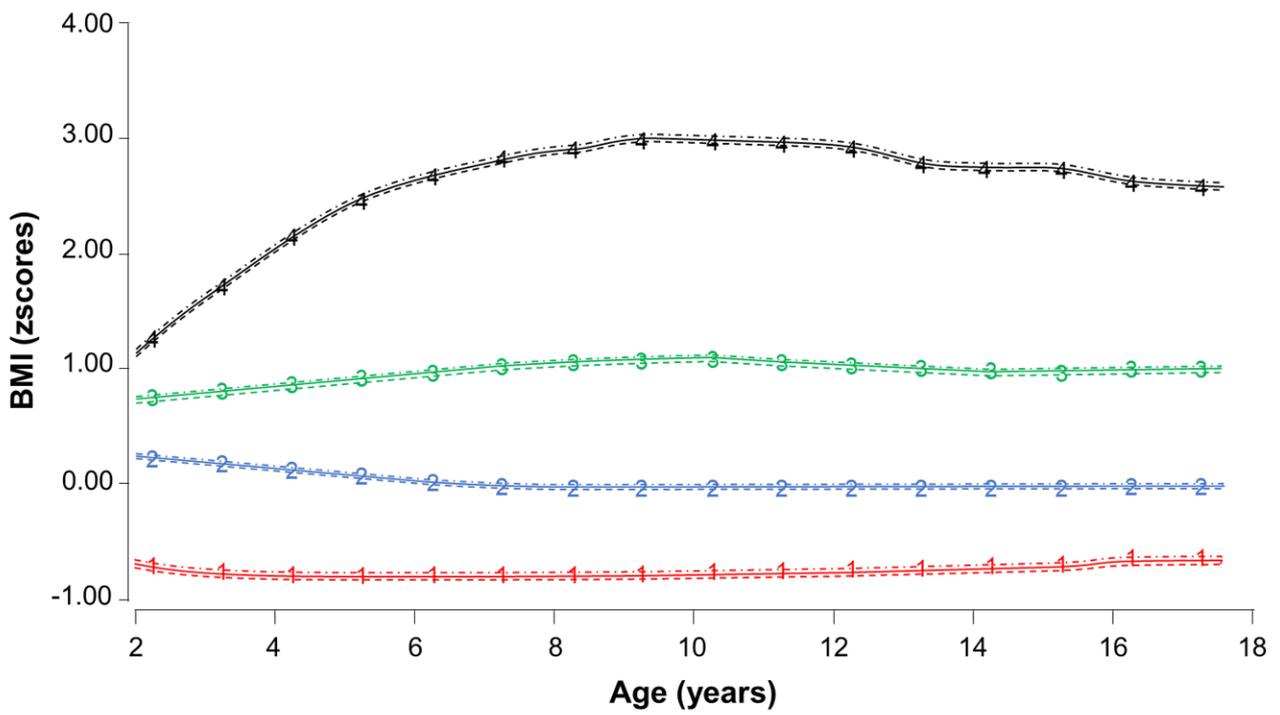


Figure 1

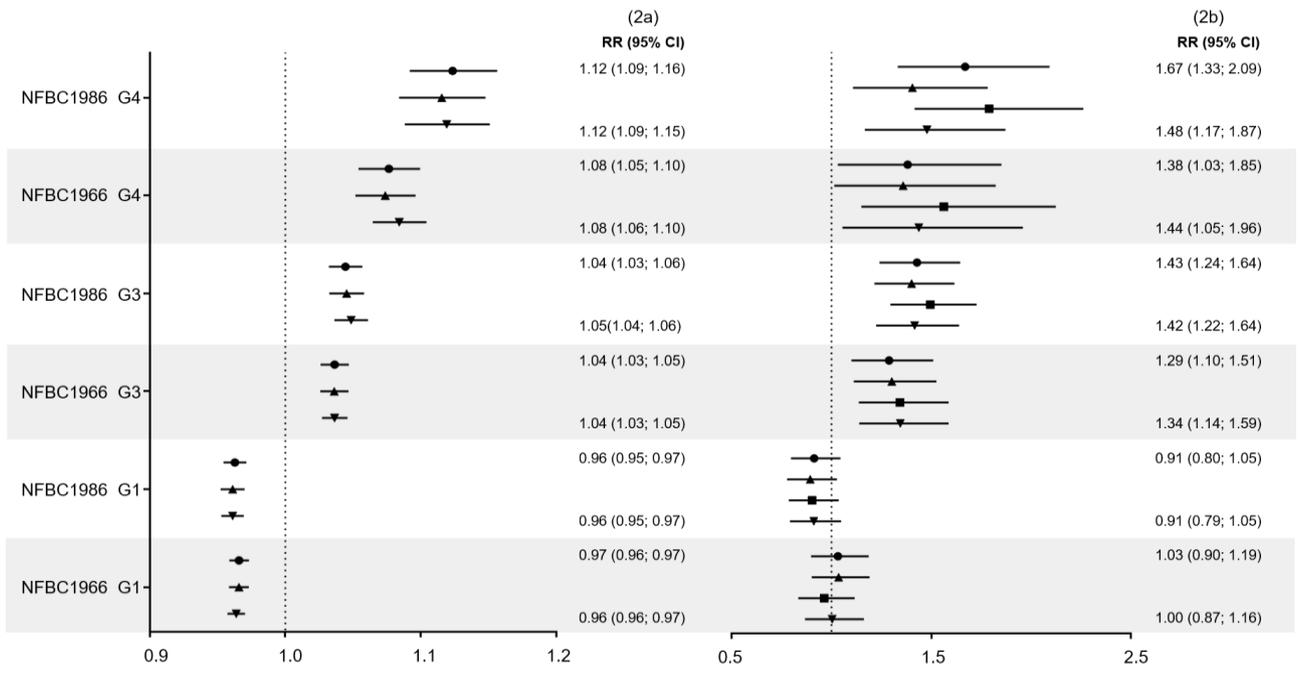


Figure 2

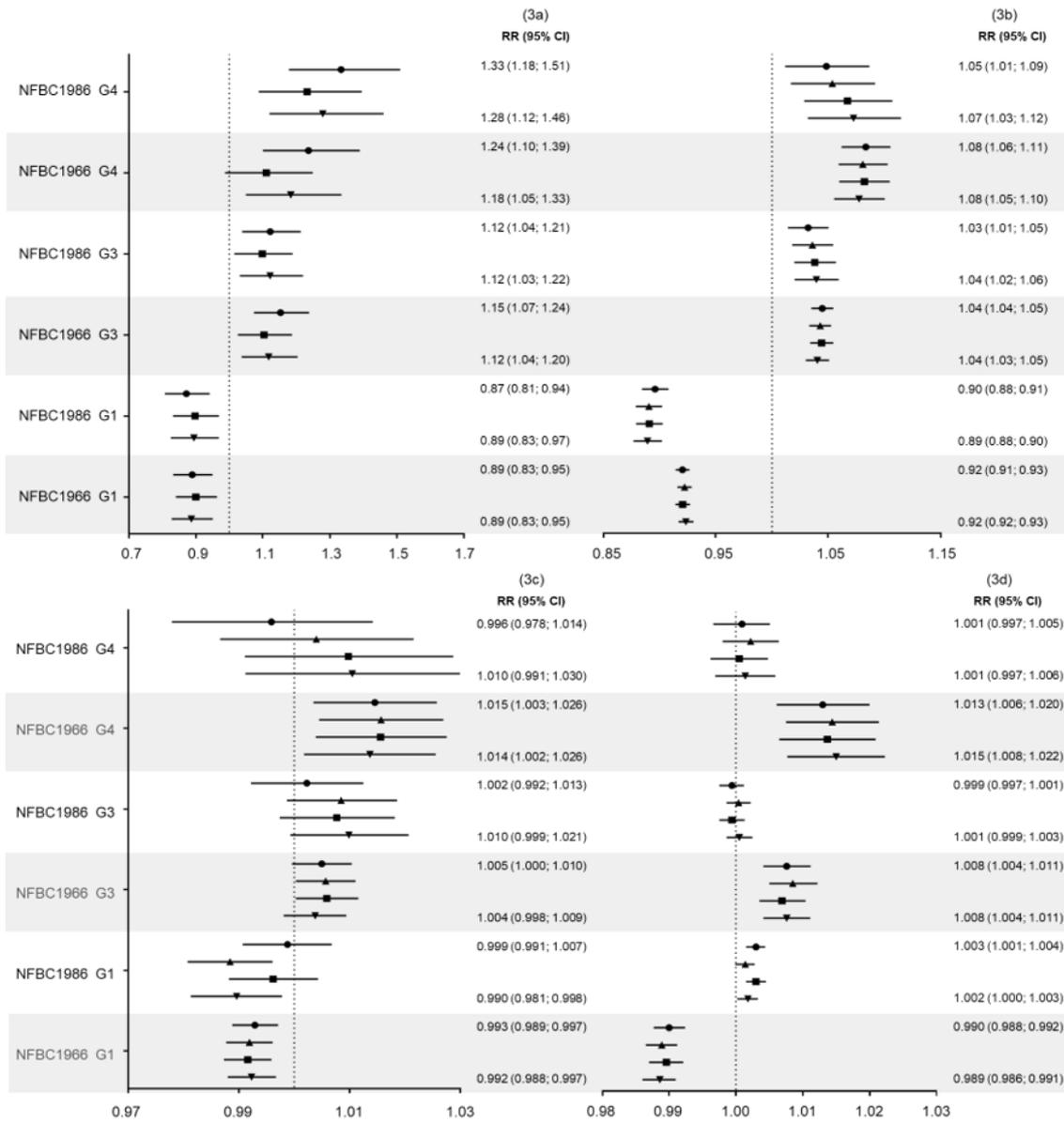


Figure 3

**Maternal and infant prediction of the child BMI trajectories;
studies across two generations of Northern Finland Birth Cohorts**

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Included in the Online Supplementary Material:

Figure S1: Flow chart of the study.

Table S1: Differences between the study participants included in the trajectory modelling and the study participants excluded of the modelling due to insufficiency of repeated BMI measures.

Supporting methods for assessing early growth indicators (BW z-score, PWV, PWV and BMI z-scores)

Table S2: Characteristics of the study by cohort.

Figure S2: Plots illustrating trajectory models for NFBC1966 and NFBC1986.

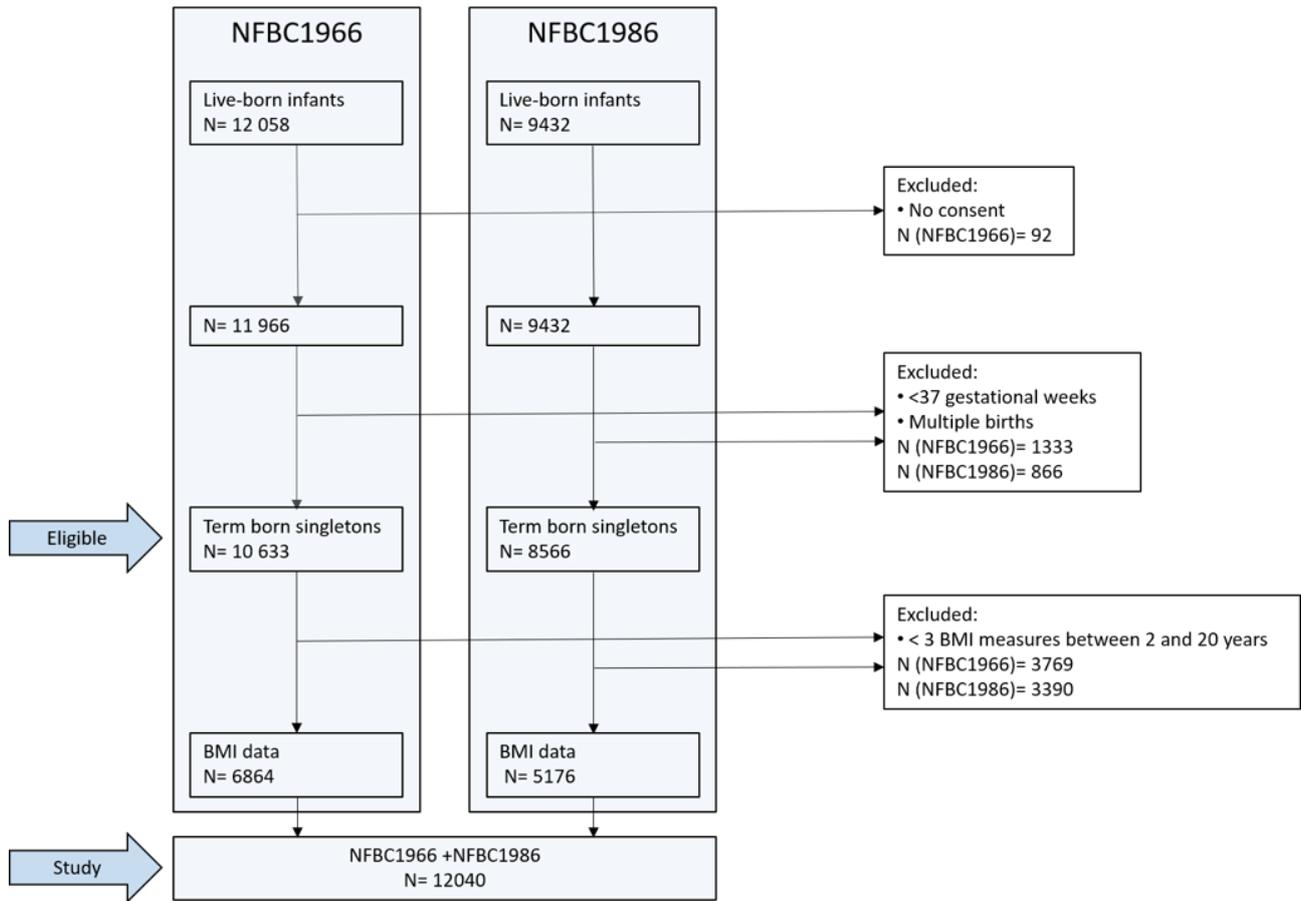
Table S3: Sensitivity analysis. Association between maternal and early growth parameters in NFBC1966 and NFBC1986 distinct group trajectories models.

Proc Traj codes for 4 groups model

Figure S3: Plots illustrating 1 to 7 trajectories models in the pooled Northern Finland Birth Cohort Studies

Table S4: Fit Statistics for the trajectory groups in the group-based trajectory models.

Table S5: Characteristics of the group trajectories according to cohort and sex.



Supplementary Figure S1: Flow chart of the study.

Supplementary table S1: Differences between the study participants included in the trajectory modelling and the study participants excluded of the modelling due to insufficiency of repeated BMI measures.

| | Inclusion in the model | | Exclusion from the model | |
|-------------------------|------------------------|------------------|--------------------------|------------------|
| | N | mean ± SD | N | mean ± SD |
| Mother's age (years) | 12004 | 28.1 ± 6.2 | 7144 | 27.5 ± 6.0 |
| Parity | 12018 | 2.3 ± 2.2 | 7139 | 2.1 ± 2.1 |
| Birth length (cm) | 11966 | 50.6 ± 2.0 | 7093 | 50.4 ± 2.2 |
| Birthweight (grams) | 12040 | 3581 ± 482 | 7158 | 3525 ± 533 |
| GA | 12040 | 40.2 ± 1.4 | 7159 | 40.2 ± 1.4 |
| | N | % | N | % |
| Sex | 12040 | | 7159 | |
| Male | 6243 | 51.8 | 3574 | 49.9 |
| Female | 5797 | 48.2 | 3585 | 50.1 |
| Maternal marital status | 12022 | | 3765 | |
| Married/cohabiting | 11596 | 96.5 | 3583 | 95.2 |
| Single | 339 | 2.8 | 156 | 4.1 |
| widowed/Divorced | 87 | 0.7 | 26 | 0.7 |
| Maternal education | 11330 | | 3709 | |
| Elementary | 5664 | 50.0 | 2371 | 63.9 |
| Vocational or secondary | 4003 | 35.3 | 1122 | 30.2 |
| Matriculation | 315 | 2.8 | 21 | 0.6 |
| Matriculation and more | 1348 | 11.9 | 195 | 5.3 |

Supporting methods for assessing early growth indicators.

Peak height velocity (PHV) and peak weight velocity (PWV) in infancy

PHV and PWV were derived from the postnatal data using the Reed1 model for boys and girls separately, using the previously described procedure^{1,2,3}. Reed1 was chosen because it showed a better fit to the early growth data than the Kouchi, Carlberg and Count models, and an equally good fit to the Reed2 model which has one more parameter than the Reed1 model⁴. Reed1 model allows the velocity to peak after birth, whereas other models force it to peak at birth. In normal individuals, weight may drop up to 10% in the first couple of weeks after birth, the PWV is thus usually not in the first weeks after birth but slightly later. Therefore, the Reed1 model is more realistic (especially for weight) and more flexible. The Reed1 model was fitted by sex on all weight and height measurements taken at 0–3 years of age, including birth weight and length. We assumed both a fixed and a random component for all four parameters in the model. For each person, the first derivative of the fitted distance curve was taken to get the weight or height velocity curve. Subsequently, the maximum of this curve was taken to obtain the PWV or PHV in infancy. This is a 4-parameter extension of the 3-parameter Count model [9] and since this model is not defined at birth ($t=0$), it was modified for this study in the same way as in Simondon et al⁵:

$$Y = A + Bt + C \ln(t+1) + D/(t+1)$$

where t = postnatal age, Y = height reached at age t and A , B , C and D the function parameters. Of the function parameters, A is related to the baseline height at birth, B to the linear component of the growth velocity, C to the decrease in the growth velocity over time, and D to the inflection point that allows growth velocity to peak after birth rather than exactly at birth. All subjects with at least one height measurement from birth to 24 months at least 0.1 month (~3 days) apart were included in the Reed1 model fitting. Although the model converged for the whole group, random effects for parameters b and c were estimated to be zero for those with only one measurement (typically birth length). Even having two measurements was inadequate to capture the shape of the growth curve and therefore we included those with a minimum of three measurements per person.

Adiposity peak and rebound

The BMI growth pattern is nonlinear as part of the normal growth and development of a child. Two periods are considered: infancy from 2 weeks to 18 months and childhood from 18 months to 13 years. Age and BMI at adiposity peak (AP) and rebound (AR) were derived from cubic models from the two age groups separately as described previously^{2,6}. In addition to fixed effects, random effects for the constant and the slope were included in the model. Individuals with fewer than three measurements per period were excluded. The models are written as:

Infancy model: $\log(\text{BMI}) = \beta_0 + \beta_1 \text{Age} + \beta_2 \text{Age}^2 + \beta_3 \text{Age}^3 + \beta_4 \text{Sex} + u_0 + u_1 (\text{Age}) + \varepsilon$

Childhood model: $\log(\text{BMI}) = \beta_0 + \beta_1 \text{Age} + \beta_2 \text{Age}^2 + \beta_3 \text{Age}^3 + \beta_4 \text{Sex} + \beta_5 \text{Age} * \text{Sex} + \beta_6 \text{Age}^2 * \text{Sex} + u_0 + u_1 (\text{Age}) + \varepsilon$

An autoregressive within person correlation structure between measurements was assumed.

For each participant, predicted BMI at AP and AR were calculated using the estimated fixed and random coefficients. Age at AP was defined as the age for a maximum BMI between 0.25 and 1.25 years and age at AR as the age for a minimum BMI between 2.5 and 8.5 years.

Z-scores

Z-scores are routinely used in clinical practices to monitor child growth but also in epidemiological studies to interpret growth measurements. Z-scores are statistical transformations, allowing to compare scores from different normal distributions. They are calculated by the formula:

$$z = (x - \mu) / \sigma$$

where x is value of the observation, μ is the population mean and σ is the population standard deviation. The Z-score system is used to compare a group to a reference population, it expresses a value as a number of standard deviations, also called z-scores or standard score, below or above the reference mean/median. BMI z-scores are measures of BMI adjusted for sex and the age of the child at the time of the measure and birthweight z-scores are measures of birthweight adjusted for sex and gestational age.

References:

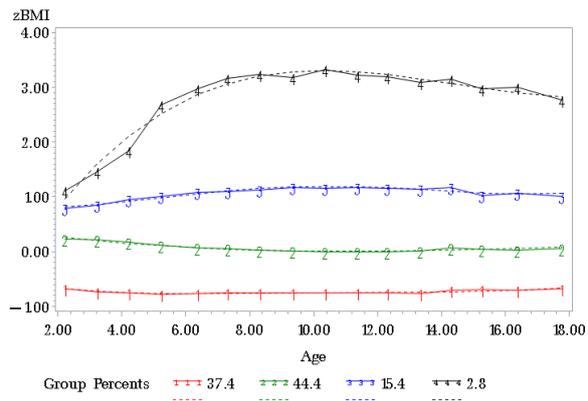
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Supplementary Table S2. Characteristics of the study by cohort.

| | NFBC1966 | | NFBC1986 | | <i>p</i> value |
|---|----------|-------------|----------|-------------|-----------------------|
| | N | mean ± SD | N | mean ± SD | |
| Mother's height (cm) | 6539 | 160.2 ± 5.4 | 5133 | 163.2 ± 5.5 | <0.0001 |
| Mother's weight (kg) | 6542 | 59.7 ± 9.1 | 5084 | 59.5 ± 9.6 | 0.006 |
| Mother's BMI (kg/m ²) | 6291 | 23.2 ± 3.3 | 5062 | 22.3 ± 3.4 | <0.0001 |
| Mother's age (years) | 6828 | 28.2 ± 6.6 | 5176 | 28.0 ± 5.5 | 0.2 |
| Parity | 6858 | 2.9 ± 2.2 | 5160 | 1.5 ± 1.9 | <0.0001 |
| Birth length (cm) | 6809 | 50.5 ± 2.0 | 5157 | 50.8 ± 1.9 | <0.0001 |
| Birthweight (grams) | 6864 | 3544 ± 489 | 5176 | 3630 ± 468 | <0.0001 |
| Birthweight z-score | 6894 | 0.04 ± 0.97 | 5173 | 0.04 ± 0.98 | 0.7 |
| Peak Height Velocity Infancy (cm/year) | 6085 | 50.6 ± 3.8 | 5027 | 52.5 ± 6.7 | <0.0001 |
| Peak Weight Velocity Infancy (kg/year) | 6260 | 12.9 ± 1.7 | 5063 | 12.6 ± 2.9 | <0.0001 |
| Age at Adiposity Peak (year) | 5391 | 0.76 ± 0.03 | 4944 | 0.69 ± 0.02 | <0.0001 |
| BMI at Adiposity Peak (kg/m ²) | 5391 | 18.0 ± 0.8 | 4944 | 17.6 ± 0.7 | <0.0001 |
| Age at Adiposity Rebound (year) | 6755 | 5.63 ± 0.87 | 5130 | 4.99 ± 1.04 | <0.0001 |
| BMI at Adiposity Rebound (kg/m ²) | 6755 | 15.4 ± 1.0 | 5130 | 15.6 ± 1.1 | <0.0001 |
| | N | % | N | % | <i>p</i> value |
| Sex | 6864 | | 5176 | | <0.0001 |
| Male | 3685 | 53.7 | 2558 | 49.4 | |
| Female | 3179 | 46.3 | 2618 | 50.6 | |
| Maternal smoking (2 months pregnancy) | 6726 | | 5152 | | <0.0001 |
| Smoker | 932 | 13.9 | 1000 | 19.4 | |
| No smoker | 5794 | 86.1 | 4152 | 80.6 | |
| Maternal marital status | 6855 | | 5167 | | 0.0013 |
| Married/Cohabiting | 6647 | 97.0 | 4949 | 95.8 | |
| Single | 170 | 2.5 | 169 | 3.3 | |
| Widowed/Divorced | 38 | 0.5 | 49 | 0.9 | |
| Maternal education | 6762 | | 4568 | | <0.0001 |
| Elementary | 4535 | 67.1 | 1129 | 24.7 | |
| Vocational or secondary | 1935 | 28.6 | 2068 | 45.3 | |
| Matriculation | 49 | 0.7 | 266 | 5.8 | |
| Matriculation and more | 243 | 3.6 | 1105 | 24.2 | |
| Operative delivery | 2464 | | 5176 | | <0.0001 |
| Non-instrumental vaginal deliveries | 1846 | 74.9 | 4245 | 82.0 | |
| Caesarian Section | 301 | 12.2 | 619 | 12.0 | |
| Others (vacuum extraction, forceps) | 317 | 12.9 | 312 | 6.0 | |

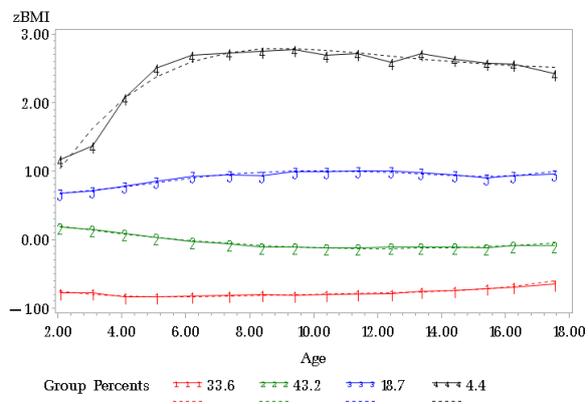
BMI course NFBC1966

4 groups from 2 to 20y



BMI course NFBC1986

4 groups from 2 to 20y



Supplementary Figure S2: Plots illustrating trajectory models for NFBC1966 and NFBC1986.

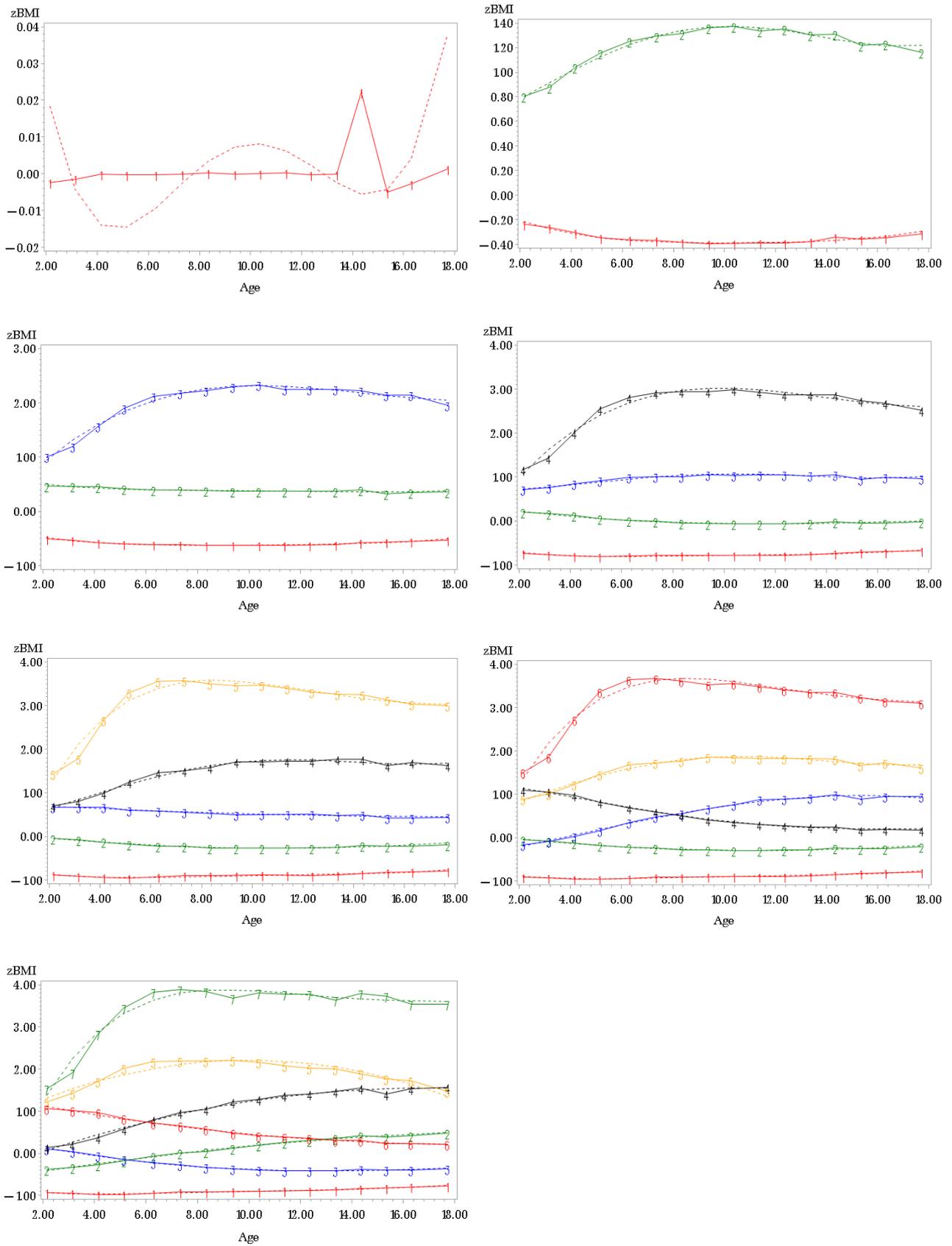
Supplementary Table S3: Sensitivity analysis. Association between maternal and early growth parameters in NFBC1966 and NFBC1986 distinct group trajectories models. Model 1: unadjusted, Model 2: adjusted for maternal age, maternal BMI, birthweight z-score, maternal education, parity, maternal smoking (according to the independent variable).

| Variables | Cohort | Model | Groups | | | | | | | |
|---------------------------------|----------|-------|--------|---------------|-----|--------|---------------|--------|---------------|--------|
| | | | RR | 95% CI | RR | 95% CI | RR | 95% CI | RR | 95% CI |
| Maternal BMI | NFBC1966 | 1 | 0.96 | 0.96 ; 0.97 | Ref | 1.04 | 1.03 ; 1.05 | 1.07 | 1.05 ; 1.10 | |
| | | 2 | 0.97 | 0.96 ; 0.97 | Ref | 1.04 | 1.03 ; 1.05 | 1.08 | 1.06 ; 1.10 | |
| | NFBC1986 | 1 | 0.97 | 0.90 ; 1.18 | Ref | 1.26 | 1.07 ; 1.49 | 1.47 | 1.08 ; 1.99 | |
| | | 2 | 0.94 | 0.81 ; 1.08 | Ref | 1.33 | 1.12 ; 1.58 | 1.53 | 1.11 ; 2.12 | |
| Maternal smoking | NFBC1966 | 1 | 0.97 | 0.90 ; 1.18 | Ref | 1.26 | 1.07 ; 1.49 | 1.47 | 1.08 ; 1.99 | |
| | | 2 | 0.94 | 0.81 ; 1.08 | Ref | 1.33 | 1.12 ; 1.58 | 1.53 | 1.11 ; 2.12 | |
| | NFBC1986 | 1 | 0.93 | 0.81 ; 1.07 | Ref | 1.41 | 1.22 ; 1.62 | 1.62 | 1.30 ; 2.02 | |
| | | 2 | 0.87 | 0.76 ; 1.01 | Ref | 1.46 | 1.26 ; 1.69 | 1.47 | 1.15 ; 1.88 | |
| BW z-score | NFBC1966 | 1 | 0.87 | 0.82 ; 0.93 | Ref | 1.15 | 1.07 ; 1.24 | 1.18 | 1.04 ; 1.34 | |
| | | 2 | 0.88 | 0.82 ; 0.94 | Ref | 1.12 | 1.04 ; 1.21 | 1.13 | 0.99 ; 1.29 | |
| | NFBC1986 | 1 | 0.87 | 0.80 ; 0.94 | Ref | 1.11 | 1.03 ; 1.20 | 1.30 | 1.15 ; 1.48 | |
| | | 2 | 0.89 | 0.82 ; 0.97 | Ref | 1.12 | 1.03 ; 1.21 | 1.24 | 1.08 ; 1.42 | |
| Age at Adiposity Peak | NFBC1966 | 1 | 0.989 | 0.987 ; 0.991 | Ref | 1.008 | 1.004 ; 1.011 | 1.012 | 1.004 ; 1.020 | |
| | | 2 | 0.987 | 0.985 ; 0.990 | Ref | 1.008 | 1.004 ; 1.011 | 1.014 | 1.006 ; 1.022 | |
| | NFBC1986 | 1 | 1.003 | 1.001 ; 1.004 | Ref | 0.999 | 0.997 ; 1.001 | 1.001 | 0.997 ; 1.005 | |
| | | 2 | 1.002 | 1.000 ; 1.003 | Ref | 1.000 | 0.998 ; 1.002 | 1.002 | 0.998 ; 1.006 | |
| Peak Weight Velocity in infancy | NFBC1966 | 1 | 0.92 | 0.91 ; 0.93 | Ref | 1.04 | 1.03 ; 1.05 | 1.08 | 1.06 ; 1.10 | |
| | | 2 | 0.92 | 0.92 ; 0.93 | Ref | 1.04 | 1.03 ; 1.05 | 1.07 | 1.05 ; 1.10 | |
| | NFBC1986 | 1 | 0.89 | 0.88 ; 0.91 | Ref | 1.03 | 1.02 ; 1.05 | 1.05 | 1.02 ; 1.09 | |
| | | 2 | 0.89 | 0.87 ; 0.90 | Ref | 1.04 | 1.02 ; 1.06 | 1.08 | 1.04 ; 1.11 | |
| Peak Height Velocity in infancy | NFBC1966 | 1 | 0.992 | 0.988 ; 0.996 | Ref | 1.004 | 0.998 ; 1.009 | 1.017 | 1.006 ; 1.028 | |
| | | 2 | 0.992 | 0.987 ; 0.996 | Ref | 1.004 | 0.998 ; 1.010 | 1.017 | 1.005 ; 1.029 | |
| | NFBC1986 | 1 | 0.996 | 0.988 ; 1.004 | Ref | 1.002 | 0.992 ; 1.012 | 0.995 | 0.979 ; 1.012 | |
| | | 2 | 0.985 | 0.976 ; 0.993 | Ref | 1.010 | 1.000 ; 1.021 | 1.012 | 0.994 ; 1.031 | |

RR: Risk Ratios, CI: Confidence Interval, BMI: Body Mass Index, BW z-core: Birthweight z-score.

Proc Traj Codes for 4 groups model :

```
proc traj data=trajNFBC out=of4 outplot=op4 outstat=os4; id project_id; var
zBMI1-zBMI16; indep windage1-windage16; model CNORM; min -30; max 30; ngroups 4;
risk gender; order 4 3 4 3; run;
%trajplot (OP4, OS4, 'BMIzscore course NFBC1966-NFBC1986', '4 groups from 2 to
20y', 'zBMI', 'Age');
```



Supplementary Figure S3: Plots illustrating 1 to 7 trajectories models in the pooled Northern Finland Birth Cohort Studies.

Supplementary Table S4: Fit Statistics for the trajectory groups in the group-based trajectory models. Bold numbers indicate the chosen model.

| Number of groups | BIC | AvPP | OCC | Estimated Percentage | Observed Percentage | Sample size |
|------------------|----------------|-------------|-------------|----------------------|---------------------|-------------|
| 1 | -176237 | | | 100 | 100 | 12040 |
| 2 | -147593 | 0.99 | 28 | 77.7 | 77.8 | 9368 |
| | | 0.97 | 113 | 22.3 | 22.2 | 2672 |
| 3 | -134036 | 0.96 | 22 | 52.7 | 52.8 | 6364 |
| | | 0.95 | 29 | 39.6 | 39.5 | 4752 |
| | | 0.97 | 388 | 7.7 | 7.7 | 924 |
| 4 | -127135 | 0.94 | 29 | 35.0 | 34.8 | 4195 |
| | | 0.93 | 17 | 43.9 | 44.0 | 5299 |
| | | 0.95 | 90 | 17.5 | 17.5 | 2106 |
| | | 0.98 | 1275 | 3.7 | 3.7 | 440 |
| 5 | -123438 | 0.93 | 41 | 24.5 | 24.2 | 2909 |
| | | 0.92 | 16 | 42.4 | 42.9 | 5168 |
| | | 0.93 | 44 | 23.0 | 22.9 | 2755 |
| | | 0.96 | 269 | 8.2 | 8.1 | 978 |
| | | 0.98 | 2530 | 1.9 | 1.9 | 230 |
| 6 | -120217 | 0.93 | 42 | 24.0 | 23.8 | 2860 |
| | | 0.92 | 17 | 41.0 | 41.6 | 5014 |
| | | 0.9 | 71 | 11.3 | 11.1 | 1333 |
| | | 0.89 | 45 | 15.3 | 15.1 | 1820 |
| | | 0.96 | 334 | 6.7 | 6.7 | 810 |
| | | 0.99 | 5725 | 1.7 | 1.7 | 203 |
| 7 | -117987 | 0.93 | 43 | 23.4 | 23.3 | 2802 |
| | | 0.84 | 29 | 15.4 | 14.9 | 1796 |
| | | 0.88 | 15 | 33.0 | 33.6 | 4051 |
| | | 0.92 | 146 | 7.3 | 7.3 | 876 |
| | | 0.96 | 547 | 4.2 | 4.2 | 508 |
| | | 0.90 | 49 | 15.6 | 15.5 | 1863 |
| | | 0.99 | 8151 | 1.2 | 1.2 | 144 |

BIC: Bayesian Information Criterion; AvPP: Average Posterior Probabilities; OCC: Odds of Correct Classification

Supplementary table S5. Characteristics of the group trajectories according to cohort and sex.

| | | Stable-low Group | | Normal Group | | Stable-high group | | Early-increase group | |
|----------|--------|------------------|------|--------------|------|-------------------|------|----------------------|------|
| | | N | % | N | % | N | % | N | % |
| Sex | Male | 2116 | 50.4 | 2874 | 54.2 | 1021 | 48.5 | 232 | 52.7 |
| | Female | 2079 | 49.6 | 2425 | 45.8 | 1085 | 51.5 | 208 | 47.3 |
| NFBC1966 | | | 56.3 | | 58.1 | | 56.4 | | 53.2 |
| | Male | 1223 | 51.8 | 1733 | 56.3 | 601 | 50.6 | 128 | 54.7 |
| | Female | 1138 | 48.2 | 1348 | 43.7 | 587 | 49.4 | 106 | 45.3 |
| NFBC1986 | | | 43.7 | | 41.9 | | 43.6 | | 46.8 |
| | Male | 893 | 48.7 | 1141 | 51.4 | 420 | 45.7 | 104 | 50.5 |
| | Female | 941 | 51.3 | 1077 | 48.6 | 498 | 54.3 | 102 | 49.5 |