- 1 Family's socioeconomic profile at birth and offspring mortality until midlife the
- 2 Northern Finland Birth Cohort 1966 study
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1 Abstract

Family's socioeconomic profile collected prenatally is known to predict offspring mortality during 2 3 early life, but it remains unclear whether it has the potential to predict offspring mortality until later 4 life. In this study, 12063 individuals belonging to the Northern Finland Birth Cohort 1966 were 5 followed up from mid-pregnancy for 52 years (570 000 person years). Five distinct socioeconomic profiles were identified by latent class analysis based on mother's marital status, education, and 6 7 occupation; father's occupation; number of family members; location of residence, room count, and 8 utilities; and family's wealth. The classes were highest status families (15.4% of the population), small families (22.1%), larger families (15.4%), average wealth families (23.4%), and rural families 9 10 (23.3%). Their associations to offspring mortality, via linkage to national offspring death records, were analysed by Cox regression, stratified by sex and age groups (0-19, 20-38 and 40-52 years). 11 12 In total, mortality was 9.2% among male and 5.0% among female offspring. Risk for midlife mortality 13 was higher among male offspring from larger families (hazard ratio 2.19, 95% confidence interval 1.32–3.63), average wealth families (1.66, 1.02–2.73) and rural families (1.63, 1.00–2.68), relative 14 15 to offspring from highest status families. It seems that family's socioeconomic profile constructed 16 prenatally has predictive value for midlife mortality among male offspring. Premature mortality of 17 men and women seem to be two distinct phenomena with differing underlying factors as 18 socioeconomic profile was not associated with mortality among female offspring.

19

20 Highlights:

Five distinct socioeconomic profiles of families were identified from general Finnish
 population in the 1960s.

- Family's socioeconomic profile had predictive value for premature mortality in male
- 2 offspring but not in female offspring.
- Premature mortality of men and women seem to be two distinct phenomena with differing
   underlying socioeconomical factors.
- 5 Key words
- 6 Cohort studies, Birth cohort, Socioeconomic status, Mortality, Latent class analysis, Parent-
- 7 offspring linkage

1 Introduction

The socioeconomic status of an individual is a key predictor of health and mortality throughout the life course[1-4]. Strikingly, the influence of low socioeconomic status on the risk of premature mortality seems to be comparable to that of tobacco use, alcohol consumption, physical inactivity, raised blood pressure, obesity, and diabetes[3]. Despite reducing health inequalities being an important policy objective[5], socioeconomic disparities persist even in welfare states[4,6].

Socioeconomic status is usually comprised of education, occupation, and income[7].
Also, housing conditions, household amenities, house location, family wealth, and family size have
been used as socioeconomic measures[8,9]. Because of the complexity of the construct, there is no
single indicator of socioeconomic status, but multiple indicators need to be combined[10].
Furthermore, the interpretation of socioeconomic measures might differ between sexes[11].

12 The socioeconomic environment one is exposed to in early life is known to influence 13 health-related behaviour and mortality across the life course[12,13]. A recent study[14] suggested 14 that parental and family-related socioeconomic factors have a prominent role in setting the environment during the formative years and are thus strongly associated with health-related 15 behaviour throughout the offspring's life course[12,15]. Moreover, there is evidence that 16 socioeconomic variables collected prenatally or at birth predict offspring's future health and 17 18 mortality in the early years [16-18]. However, it remains unclear whether not just individual variables 19 but comprehensive profiles of background families' socioeconomic status during pregnancy have 20 potential to predict offspring premature mortality later in life.

In this study, we first aimed to identify distinct socioeconomic profiles of families using cluster analysis in a Finnish birth cohort of 12 063 individuals. Second, we aimed to evaluate the association between family's socioeconomic profile at birth and offspring mortality until midlife. We

hypothesized that profiles of lower socioeconomic status would associate with increased mortality
in both female and male offspring compared to highest socioeconomic profile families. The
association was hypothesized to be seen throughout the life-course.

1 Material and methods

2 Study population

3 We used the extensive Northern Finland Birth Cohort 1966 (NFBC1966) dataset as our material[19]. It is a prospective cohort study initiated in 1965–1966, when pregnant mothers who had expected 4 5 delivery dates during the calendar year of 1966 and were residing in the Northern Finnish provinces 6 of Oulu and Lapland were asked to participate in the data collections [20]. Originally, the NFBC1966 7 included 12 055 mothers and 12 231 deliveries, with a coverage of 96.3% against all deliveries in the 8 region in 1966. Although the current focus of the NFBC1966 study is on the offspring, the data 9 collections initially comprised both the parents and children born into the cohort. The data collections started during the 10-16th gestational week and are still ongoing with periodical follow-10 ups. 11

The data used in this study were collected from questionnaires issued to the mothers during the 24—28th weeks of pregnancy, from maternity cards and birth certificates of the offspring, and from the official death records of Statistics Finland from 1965—2018. Of the 12 231 individuals born into the cohort, excluded were 96 individuals that had completely missing pregnancy data, two individuals that had undetermined sex, and 70 individuals that declined the use of their data at later stages of the follow-up. Thus, the present analysis was based on 12 063 individuals (98.6% of the initial cohort base).

The study protocol follows the Declaration of Helsinki and has been approved by the Ethics Committee of the Northern Ostrobothnia Hospital District. Mothers of the NFBC1966 members gave informed verbal consent in the beginning of the NFBC1966 in 1965—1966. In the later stages of the study, written informed consent has been obtained from NFBC1966 members participating in the follow-ups. The NFBC1966 data are administered by the NFBC Project Center,

and researchers who have been granted access to the data are allowed to handle it in a
 pseudonymized format.

3

4 Parental, family, and offspring characteristics

5 Offspring's sex (male/female) and birthweight (in grams) were documented from maternity cards 6 and birth certificates. Mother's marital status was selected from the following response 7 alternatives: married, single, widow, divorced. As most mothers were married, the responses were 8 dichotomized as 'married' or 'unmarried'.

9 Mother's education level was reported according to the following alternatives: no 10 education or itinerant school; primary school for 1—4 years; primary school for 5—8 years; 11 vocational school for  $\frac{1}{2}$ —2 years; vocational school for > 2 years; secondary school; matriculation 12 examination; matriculation examination and further studies. As most mothers had completed 5—8 13 years of primary school and/or secondary school, the responses were combined into three 14 categories as follows: 'high' (matriculation examination), 'medium' (> 4 years of studies but not 15 matriculation examination), or 'low' ( $\leq$  4 years of primary school).

Mother's and father's occupational status were asked from participants and classified as following alternatives: entrepreneur or employer; manager or clerk; labourer; unpaid family worker; no occupation. A separate question also enquired the size and type of farm if applicable; these responses were used to identify farmer families as they were common in the 1960s Northern Finland. To ensure sufficient group sizes, occupational status was recategorized as 'high' (entrepreneur, employer, manager, or clerk), 'manual work' (labourer, or unpaid family worker), 'farmer' (farmers and their partners), or 'low' (no occupation).

1	Total number of family members, and the number of family members aged $\leq$ 15 years,
2	were reported as raw numbers. We categorized both variables into tertiles (total family members:
3	$\leq 2, 3-4, \text{ or } \geq 5;$ family members aged $\leq 15: 0, 1-2, \text{ or } \geq 3$ ).

4

#### 5 Residential characteristics and wealth

Location of the residence was selected from the following alternatives: city, market town, village
centre or other population centre, periphery. The responses were re-grouped as 'city', 'population
centre' (including market town), and 'periphery'.

9 Total number of rooms in the family's residence was enquired by the following 10 question: "How many rooms are available for the family to utilize (including kitchen but excluding 11 kitchenette, sleeping alcove, bathroom, toilet, and the rooms that are rented out)?" The responses 12 were categorized into tertiles as follows: 1, 2, or ≥ 3.

The questionnaire elicited whether several utilities were available at the family's residence. Electricity, telephone, running water, and television were each elicited separately (yes/no). Additionally, the families were asked whether they owned a residence, a car, or a summerhouse (yes/no for each separately).

17

18 Offspring mortality

In Finland, cause-of-death investigation is required for all deaths (the Act relating to cause-of-death investigation 459/1973). Once the investigation is completed, a death certificate is issued, stating the date and causes of death. An independent review of the certificates is conducted by forensic pathologists at the Finnish Institute for Health and Welfare; accepted certificates are forwarded to

Statistics Finland for archiving. The Finnish death certification practice has been concluded to
 function accurately and appropriately[21].

3	For this study, we linked the NFBC1966 database with death records accumulated
4	from 1965 until the most recent update of the Statistics Finland official archive in the end of 2018.
5	As the Finnish authorities are generally informed of deaths occurring abroad, individuals with no
6	death records within the Statistics Finland database were assumed to be alive in Finland or abroad.
7	As our aim was to investigate all-cause mortality, we did not assess specific causes of death.

8

### 9 Statistical analysis

10 The data were accessed and analysed using SPSS version 27 (IBM, Armonk, NY, USA) unless 11 otherwise specified. Reclassifications relative to the original variables are presented above. 12 Distributions of the final variables were presented as percentages with frequencies (categorical 13 variables) or means with standard deviations after visual inspection for normality (continuous 14 variables).

To obtain a comprehensive view of the families' socioeconomic profiles, we applied LCA to the socioeconomic data. The overall number of classes needed to sufficiently explain the differences in response patterns is *a priori* unknown. Beginning with a one-class solution, several models are systematically tested, gradually increasing the number of classes in the model, to find the most parsimonious model which sufficiently discriminates between the individuals in each different class.

We performed LCA using Mplus version 8.4 (Muthén & Muthén, Los Angeles, CA, USA).
 All socioeconomic variables - mother's marital status, education, occupation; father's occupation;

1 total number of family members and those aged  $\leq$  15 years; location of residence, room count, utilities; family's wealth - were included in the models. Offspring sex was omitted in the LCA 2 3 procedure to ensure equal class structure and class definitions for both sexes. Models with one to six classes were tested as decided *a priori*, and model fit parameters were documented from the 4 5 data output. The choice of the best-fitting model was primarily based on the statistical significance 6 of the Lo-Mendell-Rubin and Vuong-Lo-Mendell-Rubin tests[22-24], entropy, Bayesian Information 7 Criterion, log-likelihood values, and adequate class sizes. Once the best-fitting solution was selected, 8 individuals were assigned to classes according to highest posterior membership probability. Lastly, the classes were named to reflect the distributions of the studied variables in each class. 9

The full sample of 12 063 individuals contributed to LCA and received class assignment, regardless of missing data. Percentages of missing data are shown for each variable in the results section (0.0—10.1%). Assuming data to be missing at random, Mplus used the full information maximum likelihood (FIML) method to compute parameter estimates based on all available data.

14 Finally, Cox regression was used to study the association between socioeconomic class 15 and offspring mortality. In order to comply with the proportional hazards assumption, piecewise Cox regression models were separately fitted for three age intervals, namely 16 0-19 years ("childhood and adolescence"), 20-39 years ("early adulthood"), and 40-52 years 17 ("midlife"). As there is clear discrepancy in mortality between the men and women of this 18 19 population[25], the analyses were stratified by sex. Hazard ratios (HRs) and their 95% confidence 20 intervals (CIs) were documented from the regression output. Stillbirths (n = 174) were excluded from the survival analysis. Additional survival analyses were performed with adjustment for 21 22 birthweight to test if the associations were irrespective from birth characteristics.

23

### 1 Data availability statement

NFBC1966 data is available from the University of Oulu, Infrastructure for Population Studies.
Permission to use the data can be applied for research purposes via electronic material request
portal. In the use of data, we follow the EU general data protection regulation (679/2016) and
Finnish Data Protection Act. The use of personal data is based on cohort participant's written
informed consent at his/her latest follow-up study, which may cause limitations to its use. Please,
contact NFBC project center (NFBCprojectcenter@oulu.fi) and visit the cohort website
(www.oulu.fi/nfbc) for more information.

1 Results

2 Study population

3 The background characteristics of the study population (N = 12 063) are presented in **Table 1**. Of the offspring, 48.7% were female. Most mothers were married when pregnant (95.5%) and had 4 5 attended school for > 4 years but did not complete matriculation examination (84.7%). Most 6 mothers and fathers had manual occupation (35.1% and 53.1%, respectively). A typical family during 7 the mother's pregnancy included  $\geq$  3 members (74.5%), of whom 1–2 were aged  $\leq$  15 years (42.3%). 8 The family's residence was commonly located in periphery (40.7%), had  $\geq$  3 rooms (49.4%), and was 9 equipped with electricity (83.8%) and running water (50.3%). A minority of the families owned their residence (44.0%), a car (35.9%), or a summerhouse (3.3%). 10

11

12 Latent class analysis and families' socioeconomic profiles

Goodness-of-fit statistics of LCA models with 1—6 classes are presented in **Table 2**. The five-class model was selected as the best-fitting solution. In contextual evaluation of the detected classes against the background variables, each of the classes was found to represent a distinct socioeconomic profile and named accordingly as follows:

a) Highest status families (relative class size 15.4% of the population): Families characterized by
high occupational status of the mother (41.9%) and the father (75.5%), located in cities or
population centres (83.1%). Large residences with ≥ 3 rooms (86.5%) and utilities widely
available (≥ 65.1%). Above-average wealth (own residence 40.6%, car 82.3%, summerhouse
15.7%).

1	b) Small families (22.1%): Families with $\leq$ 2 members (95.4%) and no individuals aged $\leq$ 15 at time
2	of pregnancy (92.8%). Living in a residence with $1-2$ rooms (79.2%) in a city or population centre
3	(81.9%).
4	c) Larger families (15.4%): Families with $\geq$ 5 members (98.3%) and $\geq$ 3 individuals aged $\leq$ 15 (77.0%).
5	d) Average wealth families (23.4%): Medium-sized families living in medium-sized residences with
6	average utilities and wealth.
7	e) Rural families (23.3%): Relatively large families located in periphery (86.9%), with farming as the
8	most common occupation among mothers (93.9%) and fathers (74.8%).
9	
10	Association between family's socioeconomic profile and offspring mortality
11	During the follow-up of 52 years, a total of 862 deaths occurred among the offspring. Males had
12	higher mortality (9.2%) than females (5.0%). Among male offspring, premature death in early
13	adulthood was associated with average wealth families (HR 1.93, 95% CI 1.15 $-$ 3.24) relative to
14	highest status families (Table 3). Premature death during midlife was associated with background
15	in larger families (HR2.19, 95% Cl 1.32—3.63), average wealth families (HR 1.66, 95% Cl 1.02—2.73)
16	and rural families (HR 1.63, 95% CI 1.00 $-2.68$ ). Socioeconomic profile was not associated with
17	mortality in childhood and adolescence. The results remained highly similar after adjustment for
18	offspring's birthweight (Supplementary Table 1).
19	Among female offspring, socioeconomic profile was not associated with premature
20	mortality in early adulthood or midlife <b>(Table 3)</b> . Mortality in childhood and adolescence was higher
21	among females from rural families (HR 2.91, 95% CI 0.99—8.61) compared to highest status families,
22	however, the association did not quite reach statistical significance.

Associations between individual socioeconomic factors and offspring's premature mortality stratified by sex and age categories are seen in **Supplementary Table 2**. Beyond the clusters, also mother's marital status and occupation, father's occupation, family size, residency, utilities, and wealth all associated separately with offspring's premature mortality.

5

6 Sensitivity analyses

In addition to the primary analysis that adjusted for missing data by means of FIML (n = 12 063), we
also performed a complete-case analysis (n = 9422) as a sensitivity approach (data not shown). The
final class structure of LCA and associations with mortality were similar in both approaches.
Moreover, there was no statistically significant difference in overall mortality between the
complete-case sample and those with missing data (6.9% vs. 7.9%, P = 0.10).

1 Discussion

In this large birth cohort study, we found five distinct prenatal family socioeconomic profiles. Having
been born to families with more than five members, medium-sized families with average wealth,
and relatively large families living in rural areas associated with increased midlife mortality in male
offspring. Socioeconomic profile was not associated with mortality among female offspring.

6 Although previous literature has described the association between background 7 family's socioeconomic variables and offspring health and mortality during the early life 8 course[12,15-18], studies addressing the association in later life have been lacking. The findings of 9 this study suggest that family characteristics collected prenatally have predictive value for premature mortality among males until midlife. One possible mechanism to this association is the 10 11 family socioeconomic environment during offspring's formative years influencing later healthrelated behaviour across the life course[14]. Family socioeconomic environment may also be linked 12 13 to differential vulnerability to stressors during the life course. It should be noted that this analysis 14 did not include any life course data as we specifically aimed to characterize the role of family's socioeconomic characteristics at birth, and all life course variables were seen as potential effect 15 16 mediators. Future studies are encouraged to further investigate the potential mediating pathways during the life course. 17

The present data underline the predictive role of high socioeconomic status in lower premature mortality among men in particular. Among women, family's socioeconomic background does not seem to influence the risk for premature mortality until midlife. Contradictory, a previous French study has found that parental socioeconomic status, namely paternal occupation, predicts mortality in adult age among female offspring but not among males, most likely due to effects of the World Wars[26]. When concerning the lifetime socioeconomic status, men in low social classes

1 have higher mortality than women in same classes[27,28]. There is a smaller sex difference in higher 2 social classes suggesting that men are more vulnerable to adverse effects of low socioeconomical 3 status. Although we did not find significant associations between low socioeconomic status and premature mortality in female offspring, our current results suggest that having rural background 4 5 might heighten the risk for premature mortality in childhood and adolescence in women. However, 6 further studies with a higher number of outcome events are needed to confirm this potential 7 association. Overall, our findings indicate that premature mortality of men and women are clearly 8 two distinct phenomena with potentially substantially differing underlying socioeconomic factors. Patterns of timing and causes of premature mortality differ between women and men[25], which 9 may explain the current findings. 10

This study had several strengths, one being the use of a large, unselected, 11 12 representative, population-based birth cohort containing >12 000 families with nearly 570 000 13 person years of follow-up of the offspring. A wide range of socioeconomic variables was collected prospectively during the second trimester of pregnancy. The follow-up of offspring lasted up to 52 14 15 years of age and was complete in terms of mortality without any loss to follow-up. The validity of 16 the Finnish cause-of-death register is proven to be excellent. The NFBC1966 dataset is highly 17 valuable in the context of the present research question. Another strength is the use of LCA methodology which identified five distinct socioeconomic family profiles. LCA was used not only due 18 to the strong intercorrelations between the individual socioeconomic variables, but also to reduce 19 the number of statistical tests performed and thus minimize the risk for type 1 error. Supplemental 20 analyses confirmed that several socioeconomic factors associate with premature mortality also 21 22 when studied individually, indicating that the observed associations between latent profiles and 23 mortality are not driven by just one or a few variables.

There were some limitations to this study. Despite the long follow-up and large sample size, the number of deaths accumulated in each class was relatively low. Additionally, the socioeconomic factors were collected in 1960s and represent the Finnish society more than 50 years ago. The family socioeconomic factors related to increased mortality might be different for future generations.

In summary, the present results highlight the importance of considering early life predictors of mortality and suggest that family's socioeconomic profile constructed prenatally has predictive value for premature mortality among male offspring until midlife. This study provides further evidence that premature mortality of men and women of Northern Finnish origin seem to be two distinct phenomena with differing underlying factors. Future studies are encouraged to shed light on the potential life-course mediators of this difference.

## 1 References

2 1. Marmot M. Social determinants of health inequalities. *Lancet* 2005; **365:** 1099-104.

2. Mackenbach JP, Stirbu I, Roskam AJ, et al. Socioeconomic inequalities in health in 22 European
countries. *N Engl J Med* 2008; **358**: 2468-81.

3. Stringhini S, Carmeli C, Jokela M, et al. Socioeconomic status and the 25 × 25 risk factors as
determinants of premature mortality: a multicohort study and meta-analysis of 1.7 million men
and women. *Lancet* 2017; **389:** 1229-37.

4. Lewer D, Jayatunga W, Aldridge RW, et al. Premature mortality attributable to socioeconomic
inequality in England between 2003 and 2018: an observational study. *Lancet Public Health* 2020;
5: e33-41.

- 5. Closing the gap in a generation: Health equity through action on the social determinants of
  health. WHO Commission on Social Determinants of Health 2020.
- 6. Mackenbach JP. The persistence of health inequalities in modern welfare states: the explanation
  of a paradox. *Soc Sci Med* 2012; **75**: 761-9.
- 7. Galobardes B, Shaw M, Lawlor DA, Lynch JW, Davey Smith G. Indicators of socioeconomic
   position (part 1). *J Epidemiol Community Health* 2006; 60: 7-12.
- 8. Galobardes B, Shaw M, Lawlor DA, Lynch JW, Davey Smith G. Indicators of socioeconomic
  position (part 2). *J Epidemiol Community Health* 2006; **60:** 95-101.

9. Galobardes B, Morabia A. Measuring the habitat as an indicator of socioeconomic position:
 methodology and its association with hypertension. *J Epidemiol Community Health* 2003; 57: 248 53.

- 10. Geyer S, Hemström O, Peter R, Vågerö D. Education, income, and occupational class cannot be
   used interchangeably in social epidemiology. Empirical evidence against a common practice. J
- 24 *Epidemiol Community Health* 2006; **60:** 804-10.
- 11. Krieger N, Williams DR, Moss NE. Measuring social class in US public health research: concepts,
  methodologies, and guidelines. *Annu Rev Public Health* 1997; 18: 341-78.
- 12. Viner RM, Ozer EM, Denny S, et al. Adolescence and the social determinants of health. *Lancet*2012; **379:** 1641-52.
- 13. Pavalko EK, Caputo J. Social Inequality and Health Across the Life Course. *Am Behav Sci* 2013;
  57: 1040-56.
- 14. Marttila-Tornio K, Männikkö N, Ruotsalainen H, Miettunen J, Kääriäinen M. Lower parental
   socioeconomic status in childhood and adolescence predicts unhealthy health behaviour patterns
   in adolescence in Northern Finland. *Scand J Caring Sci* 2020;.

- 15. Quon EC, McGrath JJ. Subjective socioeconomic status and adolescent health: a meta-analysis.
   *Health Psychol* 2014; 33: 433-47.
- 16. Ericson A, Eriksson M, Källén B, Zetterström R. Socio-economic variables and pregnancy
   outcome. 2. Infant and child survival. *Acta Paediatr Scand* 1990; **79:** 1009-16.
- 17. Osler M, Andersen AM, Due P, Lund R, Damsgaard MT, Holstein BE. Socioeconomic position in
  early life, birth weight, childhood cognitive function, and adult mortality. A longitudinal study of
  Danish men born in 1953. *J Epidemiol Community Health* 2003; 57: 681-6.
- 18. Kim J, Son M, Kawachi I, Oh J. The extent and distribution of inequalities in childhood mortality
  by cause of death according to parental socioeconomic positions: a birth cohort study in South
- 10 Korea. Soc Sci Med 2009; **69:** 1116-26.
- 11 19. University of Oulu: Northern Finland Birth Cohort 1966. University of
- 12 Oulu. http://urn.fi/urn:nbn:fi:att:bc1e5408-980e-4a62-b899-43bec3755243
- 20. Rantakallio P. The longitudinal study of the northern Finland birth cohort of 1966. *Paediatr Perinat Epidemiol* 1988; 2: 59-88.
- 15 21. Lahti RA, Penttilä A. The validity of death certificates: routine validation of death certification
  and its effects on mortality statistics. *Forensic Sci Int* 2001; **115**: 15-32.
- 17 22. Geiser C. Data analysis with Mplus. New York, NY, USA: The Guildford Press; 2013.
- 18 23. Nylund KL, Asparouhov T, Muthén BO. Deciding on the Number of Classes in Latent Class
   19 Analysis and Growth Mixture Modeling: A Monte Carlo Simulation Study. *null* 2007; 14: 535-69.
- 20 24. Asparouhov T MB. Using Mplus TECH11 and TECH14 to test the number of latent
- 21 classes. Mplus Web Notes: No 14. 2020;.
- 22 25. Junno JA, Pakanen L, Oura P. Unnatural-cause mortality patterns of Northern Finnish men and
  23 women diverge in adolescence A 52-year follow-up. *Prev Med Rep* 2021; 22: 101337.
- 26. Todd N, Le Fur S, Bougnères P, Valleron AJ. Impact of social inequalities at birth on the
  longevity of children born 1914-1916: A cohort study. *PLoS One* 2017; **12**: e0185848.
- 27. Ahrenfeldt LJ, Pedersen JK, Thinggaard M, Christensen K, Lindahl-Jacobsen R. Sex differences in
  health and mortality by income and income changes. *J Epidemiol Community Health* 2020; 74:
  225-31.
- 28. Mustard CA, Etches J. Gender differences in socioeconomic inequality in mortality. *J Epidemiol Community Health* 2003; 57: 974-80.
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- 32

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# Tables

<b>Table 1</b> . Socioeconomic characteristics of the full sample and the five classes. Values are percentages and frequencies unless otherwise specified.
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	ALL (N = 1	2 063)	Highes familie (n = 18		Small f (n = 26	amilies 60)	Larger familie (n = 18	es	Avera wealti familio (n = 23	h es	Rural (n = 28	families 814)
Offspring characteristics												
Female sex	48.7	5880	48.2	894	49.2	1310	47.6	884	48.6	1399	49.5	1393
Birth weight (kg)*	3.44	0.59	3.53	0.59	3.33	0.56	3.46	0.64	3.43	0.59	3.49	0.59
Missing (at least one)	0.0	0	0.0	0	0.0	0	0.0	0	0.0	0	0.0	0
Maternal characteristics												
Mother's marital status												
Married	95.5	11 519	98.7	1831	91.7	2439	92.5	1717	97.5	2804	96.9	2728
Not married	4.4	528	1.1	20	8.1	215	7.5	139	2.5	72	2.9	82
Missing	0.1	16	0.2	4	0.2	6	0.1	1	0.0	1	0.1	4
Mother's education level												
High	4.5	543	20.7	384	5.1	136	0.2	4	0.5	14	0.2	5
Medium	84.7	10 222	75.9	1386	92.6	2398	82.6	1534	94.9	2685	78.9	2219
Low	9.0	1087	3.1	57	2.1	55	15.5	287	4.5	129	19.9	559
Missing	1.7	211	1.5	28	2.7	71	1.7	32	1.7	49	1.1	31
Mother's occupational status												
High	9.9	1192	41.9	777	11.2	298	1.0	19	2.8	80	0.6	18
Manual work	35.1	4240	30.0	557	61.1	1625	36.2	672	44.1	1270	4.1	116
Farmer	22.4	2702	0.6	12	0.8	20	0.1	2	0.9	25	93.9	2643
Low	30.4	3668	25.3	469	23.4	623	61.0	1133	49.9	1437	0.2	6
Missing	2.2	261	2.2	40	3.5	94	1.7	31	2.3	65	1.1	31
Paternal characteristics												
Father's occupational status												
High	20.4	2455	75.5	1400	15.7	418	11.6	216	11.6	334	3.1	87
Manual work	53.1	6410	17.7	329	67.0	1781	78.5	1458	81.2	2336	18.0	506
Farmer	18.2	2199	2.4	45	1.6	42	0.0	0	0.2	7	74.8	2105
Low	2.6	313	1.5	27	5.9	157	1.8	34	2.9	83	0.4	12
Missing	5.7	686	2.9	54	9.8	262	8.0	149	4.1	117	3.7	104
Family characteristics												
Family member count												
≤ 2	23.5	2837	12.5	231	95.4	2538	0.0	0	0.0	0	2.4	68
3—4	37.3	4494	55.3	1025	0.0	0	0.4	8	98.4	2832	22.4	629
≥ 5	37.2	4484	31.2	579	0.0	0	98.3	1826	0.0	0	73.9	2079
Missing	2.1	248	1.1	20	4.6	122	1.2	23	1.6	45	1.4	38
Family members aged ≤ 15 years												
0	27.5	3317	17.4	323	92.8	2469	2.9	54	5.6	162	11.0	309
1—2	42.3	5105	62.5	1159	1.4	37	17.0	315	91.1	2620	34.6	974
≥ 3	26.6	3211	17.7	328	0.0	0	77.0	1430	0.0	0	51.6	1453
Missing	3.6	430	2.4	45	5.8	154	3.1	58	3.3	95	2.8	78
Residential characteristics												
Location												
City	30.4	3670	41.5	770	50.6	1345	24.0	446	38.2	1099	0.4	10
Population center	28.8	3479	41.6	772	31.3	832	29.9	556	33.4	960	12.8	359
Periphery	40.7	4914	16.9	313	18.2	483	46.0	855	28.4	818	86.9	2445
Missing	0.0	0	0.0	0	0.0	0	0.0	0	0.0	0	0.0	0
Room count												
1	14.8	1787	0.6	11	41.2	1096	5.1	95	17.9	515	2.5	70
2	30.9	3725	9.4	175	38.0	1010	30.9	574	50.3	1448	18.4	518
≥ 3	49.4	5959	86.5	1605	13.3	354	59.2	1100	27.3	784	75.2	2116
Missing	4.9	592	3.5	64	7.5	200	4.7	88	4.5	130	3.9	119
Utilities												
Electricity	83.8	10 113	99.0	1837	91.4	2430	80.9	1503	88.3	2814	64.1	1804
Telephone	20.6	2483	65.1	1207	11.3	301	12.7	235	9.9	285	16.2	455
Running water	50.3	6072	90.7	1683	57.9	1540	35.5	659	46.8	1345	30.0	845
Television	46.4	5598	78.0	1447	31.5	839	48.5	901	52.0	1496	32.5	915
Missing (at least one)	7.7	924	7.0	130	10.5	278	7.1	132	8.1	233	5.4	151
Wealth												
Own residence	44.0	5312	40.6	754	8.8	233	56.0	1040	28.0	806	88.1	2479
Own car	35.9	4328	82.3	1527	31.3	832	26.9	499	32.9	946	18.6	524
Own summerhouse	3.3	393	15.7	291	1.0	27	1.1	21	1.3	37	0.6	17
Missing (at least one)	10.1	1216	11.5	213	10.8	286	10.7	199	8.9	256	9.3	262

\*Values are means and standard deviations.

Number of classes	Class size	Average posterior membership probability	Log- likelihood	Bayesian information criterion	Entropy	Vuong-Lo- Mendell- Rubin likelihood ratio test	Lo- Mendell- Rubin likelihood ratio test
1	1.00	1.00	-129 221.7	258 668.9	-	-	-
2	0.35/0.65	0.93/0.97	-117 329.5	235 119.5	0.860	p < 0.001	p < 0.001
3	0.43/0.24/0.34	0.96/0.98/0.95	-111 908.1	224 511.7	0.913	p < 0.001	p < 0.001
4	0.30/0.23/0.27/0.20	0.95/0.98/0.89/0.87	-109 345.9	219 622.2	0.864	p < 0.001	p < 0.001
5	0.23/0.15/0.24/0.15/0.22	0.97/0.85/0.95/0.94/0.98	-107 227.0	215 619.3	0.911	p < 0.001	p < 0.001
6	0.21/0.12/0.22/0.14/0.21/0.10	0.96/0.84/0.93/0.93/0.97/0.83	-105 746.6	212 893.5	0.896	p = 0.758	p = 0.758

Table 2. Fit statistics of the latent class analysis models with 1-6 classes.

	Male offspring		Female offspring			
	Deaths per class	HR (95% CI)	Deaths per class	HR (95% CI)		
Childhood and	n=6090		n=5799			
adolescence (age 0-19)						
Highest status families	11/947 (1.2%)	1 (reference)	4/886 (0.5%)	1 (reference)		
Small families	17/1326 (1.3%)	1.10 (0.52; 2.36)	9/1284 (0.7%)	1.56 (0.48; 5.05)		
Larger families	14/955 (1.5%)	1.27 (0.57; 2.79)	7/869 (0.8%)	1.79 (0.52; 6.10)		
Average wealth families	17/1460 (1.2%)	1.00 (0.47; 2.14)	14/1385 (1.0%)	2.24 (0.74; 6.81)		
Rural families	27/1402 (1.9%)	1.66 (0.83; 3.35)	18/1375 (1.3%)	2.91 (0.99; 8.61)		
Early adulthood (age	n=5985		n=5733			
20—39)						
Highest status families	19/933 (2.0%)	1 (reference)	11/879 (1.3%)	1 (reference)		
Small families	35/1307 (2.7%)	1.32 (0.76; 2.31)	11/1270 (0.9%)	0.69 (0.30; 1.59)		
Larger families	33/937 (3.5%)	1.75 (0.99; 3.07)	9/862 (1.0%)	0.83 (0.35; 2.01)		
Average wealth families	56/1436 (3.9%)	1.93 (1.15; 3.24)	15/1368 (1.1%)	0.88 (0.40; 1.91)		
Rural families	36/1372 (2.6%)	1.29 (0.74; 2.25)	13/1354 (1.0%)	0.77 (0.34; 1.71)		
Midlife (age 40—52)	n=5825		n=5688			
Highest status families	22/917 (2.4%)	1 (reference)	22/871 (2.5%)	1 (reference)		
Small families	36/1274 (2.8%)	1.18 (0.70; 2.01)	21/1264 (1.7%)	0.66 (0.36; 1.19)		
Larger families	47/908 (5.2%)	2.19 (1.32; 3.63)	14/853 (1.6%)	0.65 (0.33; 1.27)		
Average wealth families	55/1387 (4.0%)	1.66 (1.02; 2.73)	21/1356 (1.5%)	0.61 (0.34; 1.11)		
Rural families	52/1339 (3.9%)	1.63 (1.00; 2.68)	22/1344 (1.6%)	0.65 (0.36; 1.17)		

Table 3. Hazard ratios (HR) and 95% confidence intervals (CIs) for offspring mortality in different phases of life-course.