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Inhaled Corticosteroids in Patients With Stable Chronic Obstructive Pulmonary Disease:

A Systematic Review and Meta-analysis

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Abstract

Context— Recent studies of inhaled corticosteroid (ICS) therapy for managing stable chronic obstructive pulmonary disease (COPD) have yielded conflicting results regarding survival and risk of adverse events.

Objective—To systematically review and quantitatively synthesize the effects of ICS therapy on mortality and adverse events in patients with stable COPD.

Data Sources—Search of MEDLINE, CENTRAL, EMBASE, CINAHL, Web of Science, and PsychInfo through February 9, 2008.

Study Selection—Eligible studies were double-blind, randomized controlled trials comparing ICS therapy for 6 or more months with nonsteroid inhaled therapy in patients with COPD.

Data Extraction—Two authors independently abstracted data including study characteristics, all-cause mortality, pneumonia, and bone fractures. The I^2 statistic was used to assess heterogeneity. Study-level data were pooled using a random-effects model (when I^2 50%) or a fixed-effects model (when I^2 <50%). For the primary outcome of all-cause mortality at 1 year, our meta-analysis was powered to detect a 1.0% absolute difference in mortality, assuming a 2-sided α of .05 and power of 0.80.

Results—Eleven eligible randomized controlled trials (14 426 participants) were included. In trials with mortality data, no difference was observed in 1-year all-cause mortality (128 deaths

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among 4636 patients in the treatment group and 148 deaths among 4597 patients in the control group; relative risk [RR], 0.86; 95% confidence interval [CI], 0.68–1.09; P=.20; I²=0%). In the trials with data on pneumonia, ICS therapy was associated with a significantly higher incidence of pneumonia (777 cases among 5405 patients in the treatment group and 561 cases among 5371 patients in the control group; RR, 1.34; 95% CI, 1.03–1.75; P=.03; I²=72%). Subgroup analyses indicated an increased risk of pneumonia in the following subgroups: highest ICS dose (RR, 1.46; 95% CI, 1.10–1.92; P=.008; I²=78%), shorter duration of ICS use (RR, 2.12; 95% CI, 1.47–3.05; P<.001; I²=0%), lowest baseline forced expiratory volume in the first second of expiration (RR, 1.90; 95% CI, 1.26–2.85; P=.002; I²=0%), and combined ICS and bronchodilator therapy (RR, 1.57; 95% CI, 1.35–1.82; P<.001; I²=24%).

Conclusions—Among patients with COPD, ICS therapy does not affect 1-year all-cause mortality. ICS therapy is associated with a higher risk of pneumonia. Future studies should determine whether specific subsets of patients with COPD benefit from ICS therapy.

Chronic obstructive pulmonary disease (COPD) is a smoking-related lung disease characterized by recurrent episodes of cough, sputum production, and breathlessness. The disease represents a substantial public health burden, affecting 10 million to 15 million persons in the United States. Currently, COPD is the fourth leading cause of death in the United States, accounting for 120 000 deaths annually, and it is expected to be the third leading cause of death by 2020. No pharmacotherapy and few interventions, other than smoking cessation and supplemental oxygen, have been shown to improve mortality in patients with COPD. Therefore, current medical interventions focus on reducing symptoms and complications associated with COPD. Relevated levels of systemic and pulmonary inflammation in patients with stable COPD have led to the development of inhaled corticosteroid (ICS) therapies for such patients. These therapies can be formulated as monotherapy or combination therapy (ie, an ICS combined with a long-acting β_2 -agonist [LABA]).

In patients with COPD and a forced expiratory volume in the first second of expiration (FEV₁) less than 50% predicted, several randomized controlled clinical trials have demonstrated that ICS therapy reduces the frequency of COPD exacerbation. ^{10–16} Thus, current US and international guidelines recommend ICS therapy in addition to bronchodilators for patients with symptomatic COPD, FEV₁ less than 50% predicted, and repeated exacerbations. ^{17,18} However, only 1 randomized clinical trial has demonstrated improved survival with combination therapy (ICS/LABA) as compared with long-acting anticholinergic therapy. ¹⁵ Additionally, ICS therapy has been associated with increased risk of pneumonia in several recent large randomized clinical trials. ^{12,13,15} Observational studies demonstrate an increased risk of osteopenia and fractures associated with long-term ICS use, particularly at high doses. ^{19,20} Prior meta-analyses have yielded conflicting results for the effect of ICS therapy on quality of life, ^{21–23} and none have incorporated the recently completed large randomized trials reporting the subsequent risk of pneumonia.

Given the high morbidity and mortality associated with pneumonia and fracture in patients with COPD, it is important to understand the association of ICS use with these adverse events. Additionally, the role of ICS therapy in patients with stable COPD is

controversial.^{24,25} Therefore, the objective of this report is to determine the potential benefit and harms of adding ICS therapy to placebo or non-ICS inhaled therapy for stable COPD. We conducted a systematic review and meta-analysis of randomized controlled trials to determine associations of ICS use of 6 or more months' duration with all-cause mortality and risks of fracture and pneumonia in patients with stable COPD. We also assessed associations of ICS therapy with outcomes among patients with COPD according to the following specific characteristics: high vs low ICS dose, shorter vs longer duration of ICS therapy, severity of COPD, and ICS monotherapy vs combined ICS and bronchodilator therapy.

METHODS

We followed the Quality of Reporting of Meta-analyses guidelines for reporting our meta-analysis and results. 26

Data Sources and Search Strategy

We searched MEDLINE, CENTRAL, EMBASE, CINAHL, Web of Science, and PsychInfo (to February 9, 2008) to identify potentially relevant articles. Our search included combined Medical Subject Headings and keywords for adults with COPD (study population); use of an ICS (study intervention); and randomized clinical trials (study design). The search strategy used standard filters to identify randomized clinical trials. There were no language restrictions. We reviewed the bibliographies of all selected articles and relevant review articles to identify additional studies.

Study Selection

Two reviewers (M.B.D., M.W.P.) independently screened studies for inclusion, retrieved potentially relevant studies, and determined study eligibility. Disagreements were resolved by group consensus. Analysis was restricted to published double-blind, randomized controlled trials comparing ICS use (6 months) with control therapy (placebo or non-ICS inhaled medications). We included studies with 6 or more months' duration of ICS use, because we hypothesized that many adverse events (eg, fractures) would require a minimum duration of exposure. Included studies were those that enrolled adults (40 years) with COPD (defined as a clinical diagnosis of COPD or as current or former smoking [10 pack years] and an FEV₁ to forced vital capacity ratio <0.70). Trials enrolling patients with asthma or evidence of reversibility by standard bronchodilator testing were excluded. Agreement between reviewers regarding study inclusion was assessed using the Cohen κ statistic. Remarks the statistic of th

Data Extraction

Two reviewers (E.C.D., E.F.) independently abstracted data and methods from included studies using custom-made standardized forms. Extracted data were entered into Microsoft Excel (Microsoft Corp, Redmond, Washington) and were checked by a third reviewer (M.W.P.). Abstracted data included study design (eg, date of conduct, sample size), patient characteristics, study methodology (eg, eligibility criteria, method of randomization, and blinding), intervention (eg, ICS type, dose, and duration; outcome definitions), and main

results. To compare effects of different doses of ICS, the ICS doses in each study were converted to beclomethasone-equivalent doses.²⁹ Disagreements were resolved by group consensus. Methodological quality of included studies was evaluated by collecting data on sources of systematic bias according to published guidelines.³⁰ These data included description of sequence generation, allocation concealment, assessor blinding, incomplete outcome data, selective reporting, eligibility criteria, therapies, and excluded patients. We also quantified study quality using the Jadad score.³¹

The primary outcome of this meta-analysis was all-cause mortality 1 year after initiation of ICS therapy. Secondary outcomes included pneumonia, fractures, and mortality at 6-month, 2-year, and 3-year follow-up. Definitions of adverse events (pneumonia, fracture) used in included studies were heterogeneous and sometimes absent. However, studies reporting pneumonia and fracture outcomes without explicit criteria for these events were included in the analysis.

Statistical Analysis

We report dichotomous outcomes as relative risks (RRs) and their respective 95% confidence intervals (CIs). The I^2 statistic was calculated to determine the proportion of between-study variation due to heterogeneity, with suggested thresholds for low (25%–49%), moderate (50%–74%), and high (75%) values.³² Study-level data were pooled using a random-effects model. In analyses in which heterogeneity was low (I^2 <50%), a fixed-effects model was used. Studies with multiple therapeutic groups were analyzed as separate 2-group trials in each comparison, using the non-ICS comparator (LABA) or placebo therapy for comparison. For example, in a study including 4 treatment assignments (placebo, ICS, LABA, ICS/LABA), ICS was compared with placebo and ICS/LABA was compared with LABA. Publication bias was assessed using funnel plots and the Begg statistical test.³³ All statistical analyses were performed using Stata version 10.0 (StataCorp, College Station, Texas).

We defined several a priori subgroups to explore observed heterogeneity, including dose and duration of ICS therapy, baseline COPD severity (defined using FEV_1), and use of monotherapy vs combination therapy.

RESULTS

Search Results and Study Characteristics

The electronic database searches identified 3153 citations. After evaluating these citations, review articles, and the bibliographies of included studies, we included 11 randomized controlled trials $^{10-16,34-37}$ (Figure 1). The Cohen κ statistic for agreement on study inclusion was 0.92.

Characteristics of included trials are summarized in Table 1. A total of 14 426 participants (mean, 1312; range, 290–6112) with COPD were included in these studies. Although all studies were published in English, they represent an international experience, having been conducted in more than 40 countries. Ten studies were multicenter trials.

All studies defined COPD using pulmonary function testing as recommended by international guidelines. ¹⁷ Entry criteria differed regarding baseline FEV₁ reported as having been measured before or after bronchodilator administration. All studies excluded individuals with recent exacerbations. Mean baseline FEV₁ was 51% predicted (range, 36% to 86% predicted). Mean study duration was 24 months (range, 6–40 months). Six studies lasted less than 36 months. Three types of ICS were studied, as monotherapy or as combination therapy: fluticasone, ^{10,12,13,15,16,35} triamcinolone, ³⁴ and budesonide. ^{11,14,36,37} The mean ICS dose (in beclomethasone equivalents) was 930 μ g/d (range, 250–4800 μ g/d). Five trials were placebo-controlled. ^{10,16,34,36,37} Three studies had 3 comparator groups including placebo, LABA, and combination ICS/LABA. ^{11,12,14} Other studies used different comparator groups: tiotropium plus placebo vs tiotropium plus salmeterol vs tiotropium plus salmeterol/fluticasone vs tiotropium. ¹⁵

Quality Assessment

In general, methodological quality of included studies was good. All included studies had a Jadad score of 3 or greater (Table 2). All studies had clearly defined eligibility criteria, therapies, and reasons for patient exclusion. Allocation sequence generation and concealment were adequately described in 7 studies. ^{10–13,15,35,37} All studies demonstrated adequate description of assessor blinding and reported statistical methods including sample-size calculations. Nine studies reported intention-to-treat analysis. ^{11–15,34–37} Follow-up rates for all studies exceeded 95%. Discontinuation of assigned study medication was common (range, 25%–43%). Eight studies reported industry funding. ^{10–12,14–16,36,37} All studies reported some involvement with industry sponsors (eg, supply of active drug and placebo, statistical analysis).

Publication bias was assessed by inspection of funnel plots as well as analysis of the Begg statistic for each pooled analysis. There was no evidence of significant publication bias for the primary or secondary outcomes.

Evidence Synthesis

Findings of our main analysis were robust to sensitivity analyses restricted to studies with explicit outcome blinding, no industry funding, assessment of bronchodilator reversibility, and secular trends (ie, publication after 2003). Thus, we present only the results of our main analysis. No study found a statistically significant 1-year mortality benefit from ICS use as compared with placebo. Three studies reported a significantly higher rate of pneumonia with ICS use, ^{12,13,15} whereas no studies found a significantly increased risk of fractures.

All-Cause Mortality

All-cause mortality at 1-year follow- up (Figure 2) was reported in 5 studies (9233 patients). These studies included 128 deaths among 4636 individuals in the treatment group and 148 deaths among 4597 individuals in the control group. Inhaled corticosteroid therapy was not associated with a decreased risk of death at 1-year follow-up (RR, 0.86; 95% CI, 0.68–1.09; P=.20; P=.20; P=.20%).

Only 2 studies (6557 patients) reported all-cause mortality at 6-month follow-up, limiting pooled analysis. 12,16 Neither study demonstrated a survival difference with ICS use. Similarly, only 2 studies (7435 patients) reported all-cause mortality at 2-year follow-up. 12,15 Of these, only 1 study 15 reported that ICS therapy was associated with a mortality benefit (RR, 0.56; 95% CI, 0.33–0.94). When combined, these 2 studies demonstrated no difference in 2-year all-cause mortality for ICS therapy vs no ICS therapy (RR, 0.90; 95% CI, 0.69–1.12; P = .45; $I^2 = 55\%$). Five studies reported all-cause mortality rates at 3-year follow-up (9537 patients). These studies reported 498 deaths among 4777 individuals in the treatment group and 506 deaths among 4760 individuals in the control group. 10,12,34,36,37 There was no difference in all-cause mortality at 3-year follow-up for ICS therapy vs no ICS therapy (RR, 0.98; 95% CI, 0.87–1.10; P = .70; $I^2 = 0\%$).

Pneumonia

Seven studies (10 776 patients) reported pneumonia outcomes. These studies included 777 events among 5405 individuals in the treatment group and 561 events among 5371 individuals in the control group (Figure 2). $^{10-13,15,35,37}$ Patients receiving ICS therapy had a higher incidence of pneumonia (RR, 1.34; 95% CI, 1.03–1.75; P=.03). There was a moderate degree of heterogeneity for this analysis (P<.001; I²=72%). The 3 individual studies that reported a significantly higher pneumonia incidence contributed 61.9% of the overall weight in the pooled analysis. Each of these studies was published after 2006.

Fracture

Three studies (8131 patients) reported fracture events. These studies included 195 events among 4073 individuals in the ICS treatment groups and 178 events among 4058 individuals in the control groups. 10,12,36 There was no difference in the risk of fracture between ICS users vs nonusers (RR 1.09; 95% CI, 0.89–1.33; P=.40; I²=29%).

Subgroup Analyses

Subgroup analyses were performed to investigate heterogeneity for associations of ICS use with our primary and secondary outcomes within strata of 4 different study characteristics (Table 3). These 4 subgroups were ICS dose (low dose [<800 μg] vs medium dose [800–1000 μg] vs high dose [>1000 μg], in beclomethasone equivalents), duration of ICS use (2 years vs >2 years), baseline severity of illness (FEV $_1$ <40% predicted vs 40%–50% predicted vs >50% predicted), and ICS modality (monotherapy vs LABA-combination therapy).

No effect modification was demonstrated for 1-year all-cause mortality or fracture risk in any of the 4 explored subgroups, as expected from the lack of statistical heterogeneity in these analyses. However, there was a significantly higher risk of pneumonia in the following subgroups: highest ICS dose (RR, 1.46; 95% CI, 1.10–1.92; P=.008; I²=78%) 10,12,13,15,35 ; shorter (2 years) duration of ICS use (RR, 2.12; 95% CI, 1.47–3.05; P<.001; I²=0%) 11,13,15,35 ; higher baseline COPD severity (ie, mean FEV $_1$ < 40% predicted) (RR, 1.90; 95% CI, 1.26–2.85; P = .002; I²=0%) 11,15,35 ; and combination therapy (RR, 1.57; 95% CI, 1.35–1.82; P<.001; I²=24%). $^{11-13,15,35}$

COMMENT

This systematic review and meta-analysis of 11 randomized controlled trials of ICS therapy of 6 or more months' duration in patients with stable COPD demonstrated several important results. First, ICS therapy was not associated with improved survival during follow-up periods ranging from 6 months to 3 years. Second, ICS therapy was associated with a 34% increased risk of pneumonia. Third, ICS therapy was not associated with increased fracture risk, although few studies reported this outcome.

The lack of a survival benefit with ICS therapy reported here is consistent with the results of 2 previous meta-analyses. ^{21,23} In contrast to the prior meta-analyses, the current analysis includes a recent randomized controlled trial that demonstrated a survival benefit with ICS therapy. Nonetheless, our analysis demonstrated no short- or long-term mortality benefit with ICS therapy. Furthermore, our analysis did not reveal a specific subgroup of patients with COPD for whom a survival benefit might exist.

We demonstrated a significantly increased risk of pneumonia with ICS use. There has been increasing controversy in the literature regarding the association between ICS use and increased pneumonia risk. ^{38–40} Establishing and quantifying this risk is essential, given the high morbidity and mortality associated with pneumonia in patients with COPD. ^{41,42} Three recently published randomized trials demonstrated a potential association between ICS use and increased pneumonia incidence. ^{12,13,15} These studies contributed the majority of the weight to the pooled estimate. Two of these trials had clearly defined, albeit different, definitions of pneumonia. The increased risk of pneumonia in recent clinical trials may reflect improved detection of pneumonia associated with ICS use. Explicit definitions for pneumonia and acute COPD exacerbations in future clinical trials may prevent misclassification bias that could alter true associations of ICS with COPD exacerbation rates and pneumonia.

The mechanism of the association of ICS therapy with pneumonia remains unclear. One potential explanation is that ICS therapy may increase pneumonia risk in patients with COPD by increasing local airway immunosuppression.⁴³ This could diminish the ability of the innate immune system to defend against primary bacterial infections or postviral superinfections. This hypothesis is supported by a recent multicenter trial that suggested that ICS use reduces lung-specific but not generalized biomarkers of systemic inflammation in patients with COPD. 44 Consistent with this hypothesis, our results demonstrated that higher ICS doses are associated with increased pneumonia risk in a dose-dependent manner. Combined therapy with an ICS and a LABA may increase the distal ICS delivery to the alveolar bed when compared with monotherapy, which could further propagate the immunosuppressive effects of ICS. 45,46 Supporting this, our subgroup analysis showed a higher risk of pneumonia with combination therapy. Interestingly, ICS exposure for 2 years or less was associated with a higher risk of pneumonia than was ICS exposure for longer than 2 years. The biological explanation for this observation remains unclear. While the conclusions drawn from subgroup analyses are limited and hypothesis-generating, they highlight the need for further exploration of the association between ICS therapy, pneumonia risk, and subsequent morbidity and mortality in patients with COPD.

The association of ICS therapy with increased rates of pneumonia reported in our meta-analysis should be considered by clinicians and guideline developers when evaluating the role of ICS therapy in the management of stable COPD. This finding must be balanced with those of other reports describing beneficial effects of ICS therapy. A recent meta-analysis of 13 139 patients with COPD demonstrated that ICS use of longer than 6 months' duration slowed the rate of decline in quality of life as measured by the St George's Respiratory Questionnaire, a 51-item questionnaire scored from 0–100, with higher scores indicating worse quality of life (weighted mean difference, –1.22 units/y; 95% CI, –1.83 to –0.60; 2507 participants). Although data were not sufficient for pooling, the authors also observed that symptom scores decreased with ICS use. While different quality-of-life measures were used in the 11 studies included in our meta-analysis, 8 reported a statistically significant improvement in St George's Respiratory Questionnaire scores. Thus, the increased pneumonia risk associated with ICS therapy should be carefully balanced with the known favorable effects of ICS use on quality of life.

Our systematic review and meta-analysis has limitations. First, due to lack of mortality data from all 11 studies, our pooled results were likely underpowered to detect a significant difference in mortality. For example, for the primary outcome of all-cause mortality at 1 year, our meta-analysis was powered (assuming a 2-sided \alpha of .05 and power of 0.80) to detect a 1.0% absolute difference in mortality (2.2% vs 3.2%) but not the 0.5% difference in mortality identified (2.7% vs 3.2%). Demonstrating statistical significance for the identified 0.5% difference between groups would require 44 190 patients. Second, most studies used different definitions for pneumonia. Pneumonia can have a high degree of clinical overlap with COPD exacerbations. If a clinical outcome was misclassified as pneumonia rather than an exacerbation, one would observe an increased incidence of pneumonia with a decrease in acute exacerbations. Third, considerable heterogeneity existed in the type and dose of ICS used in studies of COPD. Given the different pharmacodynamics of these therapies, including drug metabolism and synergy with other inhaled therapies, pooling of different studies may attenuate the effect of a single agent. Fourth, considerable qualitative heterogeneity was present in the populations studied in these clinical trials. Different criteria for age, disease severity (ie, baseline FEV₁), and smoking status may limit the generalizability of our findings to a particular population. We used the I^2 statistic to quantify between-study heterogeneity created by these variations in study design and characteristics. The use of I^2 has limitations, because this statistic assesses only 1 aspect of study variability and does not consider the clinical implications of the observed heterogeneity. ³² Fifth, while the studies in our meta-analysis reported fracture risk, the total number of fracture events was low. Compared with randomized controlled trials, observational studies may be better suited to detect uncommon adverse events. This is highlighted by the report of a recent meta-analysis of case-control studies (43 783 cases and 259 936 controls) of ICS use that demonstrated an increased risk for nonvertebral fracture (RR, 1.12; 95% CI, 1.00–1.26).⁴⁷ Sixth, our analysis did not incorporate data on quality of life and symptom changes associated with ICS therapy.

Clinical Controversies

Results of this meta-analysis raise important questions regarding ICS therapy in patients with stable COPD. First, given that ICS therapy decreases exacerbations and has a neutral effect on mortality, should physicians be concerned about increased pneumonia risk? While reduced exacerbation rates is an important goal in patients with COPD, ICS-associated pneumonias could diminish the quality of life of patients with COPD through recurrent hospitalizations and decline in lung function. Furthermore, an increase in pneumonia could potentially contribute to the mixed observations of prior meta-analyses on global quality-of-life outcomes. Clinicians should weigh the adverse events and associated benefits of ICS therapy when considering initiating this treatment.

Second, if pneumonia is an important cause of death in patients with COPD, why is survival no different with ICS therapy? Given the heterogeneity of manifestations of COPD, certain populations of patients with COPD may benefit while others may be harmed, resulting in no net survival difference in survival with ICS therapy. In fact, a study of ICS therapy in patients with mild COPD reported a trend toward a protective effect.³⁷

Recognizing the adverse events associated with ICS use is especially important, since clinicians may increase ICS therapy from moderate to high doses in patients who are not responding. Our data suggest that increasing to higher ICS doses may place patients at greater risk for pneumonia. While the results of our subgroup analysis suggest the existence of subpopulations of patients with COPD who might be at higher risk for pneumonia, our evidence is not conclusive and is only hypothesis-generating. Analysis of existing observational studies and future clinical trials may help physicians determine an optimal ICS dose that balances the potential risks and benefits of this therapy. Until studies can confirm an unequivocal benefit of ICS therapy in a group of patients with COPD, patients should receive the lowest effective ICS dose to minimize potential adverse effects.

Conclusions and Future Considerations

Our systematic review and meta-analysis shows that while ICS use in patients with stable COPD does not significantly affect mortality, there is a significantly increased incidence of pneumonia. This risk may be greatest in patients with the lowest baseline FEV₁ and in those receiving the highest ICS dose, shortest duration of ICS therapy, and combination therapy. Further studies with rigorously defined pneumonia and exacerbation outcomes are warranted to better understand the potential harms associated with ICS use. Future studies are also needed to determine an optimal subgroup of patients with COPD that may benefit from ICS use. Finally, our study reinforces the importance for clinicians to weigh the potential risks and benefits of ICS therapy in patients with stable COPD.

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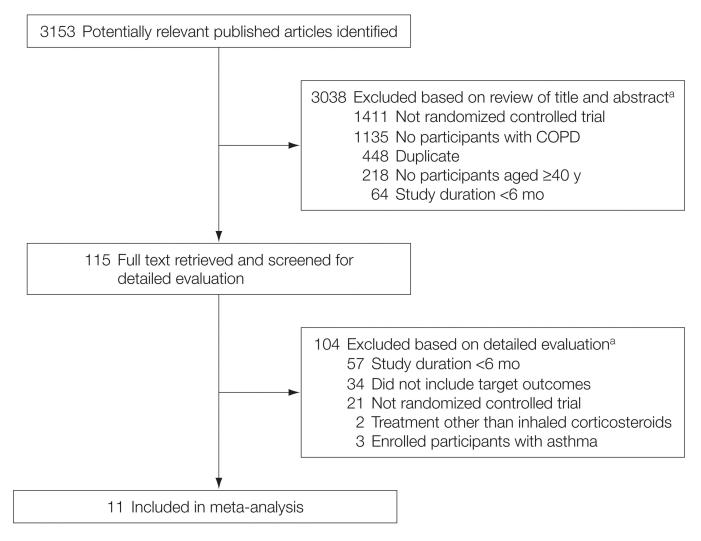


Figure 1. Flow of Study Selection

^a Articles could be excluded for more than 1 reason; therefore, summed exclusions exceed total. COPD indicates chronic obstructive pulmonary disease.

1-y Mortality

	No./7	Total	Relative Risk	Weight,	Favors : Favors
Source	Treatment	Control	(95% CI)	(%)	ICS Control
Aaron et al,35 2007	6/145	6/148	1.02 (0.34-3.09)	4.00	
Calverley et al,11 2003a	6/257	5/256	1.20 (0.37-3.87)	3.38	
Calverley et al,11 2003a	5/254	13/255	0.39 (0.14-1.07)	8.74	
Calverley et al,12 2007a	47/1534	60/1524	0.78 (0.53-1.13)	40.56	-
Calverley et al,12 2007a	46/1533	40/1521	1.14 (0.75-1.73)	27.06	 -
Kardos et al,13 2007	7/507	9/487	0.75 (0.28-1.99)	6.19	
Szafranski et al,14 2003	5/198	9/205	1.58 (0.20-1.69)	5.96	
Szafranski et al,14 2003a	6/208	6/201	0.97 (0.32-2.95)	4.11	
Overall	128/4636	148/4597	0.86 (0.68-1.09)	100.00	\Leftrightarrow
Test for heterogeneity: P	$0 = .60, I^2 = 0$	%			0.1 1.0 10.0
					Relative Risk (95% CI)

Pneumonia Risk

	No./1	Total	Odds Ratio	Weight	Favors : Favors
Source	Treatment	Control	(95% CI)	(%)	ICS Control
Aaron et al,35 2007	1/145	1/148	1.02 (0.06-16.16)	0.86	▼
Burge et al,10 2000	87/372	101/370	0.86 (0.67-1.10)	19.02	-
Calverley et al,11 2003a	5/257	2/256	2.49 (0.49-12.72)	2.31	
Calverley et al,11 2003a	8/254	7/255	1.15 (0.42-3.12)	5.28	
Calverley et al,12 2007a	284/1534	190/1524	1.48 (1.25-1.76)	20.93	
Calverley et al,12 2007a	303/1533	205/1521	1.47 (1.25-1.73)	21.08	
Kardos et al,13 2007	23/507	7/487	3.16 (1.37-7.29)	6.85	—
Vestbo et al,37 1999	16/145	24/145	0.67 (0.37-1.20)	10.60	
Wedzicha et al,15 2008	50/658	24/665	2.11 (1.31-3.38)	13.07	
Overall	777/5405	561/5371	1.34 (1.03-1.75)	100.00	\Diamond
Test for heterogeneity: F	$P < .001, I^2 = 7$	2%			
					0.1 1.0 10.0
					Relative Risk (95% CI)

Figure 2.

Effects of Inhaled Corticosteroids on 1-Year All-Cause Mortality and Risk of Pneumonia ^aStudy assessed 2 treatment vs control comparisons; see Table 1 for details.

Size of data markers is proportional to the weight of the study in the analysis. CI indicates confidence interval; ICS, inhaled corticosteroid.

Table 1

Study Characteristics of Randomized Controlled Trials Included in Meta-analysis

Source/Intervention	Medication	No.	FEV ₁ , % Predicted, Mean (SD)	Current Smoker, No. (%)	Trial Length, mo	Definition of Pneumonia in Primary Manuscript	Primary Outcome	Outcome Results (Treatment vs Control)
Aaron et al, 35 2007 Treatment	Fluticasone (500 µg)/ salmeterol (50 µg) wice daily + tiotropium (18 µg) daily	145	39 (11.9)	46 (32)	12	Definition not provided	Proportion of participants experiencing 1	60% vs 64.8% (absolute risk
Control	Salmeterol (50 µg) twice daily/tiotropium (18 µg) daily	148	38 (13.1)	36 (24)			exacerbation	reduction, 4.8)
Burge et al, ¹⁰ 2000								
Treatment	Fluticasone (500 µg) twice daily	372	50 (14.9)	145 (36)	36	Lower respiratory tract	Decline in FEV ₁	-50 (SE, 4.1) mL/y vs -59 (SE,
Control	Placebo twice daily	370	50 (14.9)	144 (39)		CVCIII		4:4) (IIII.) (F. 1-10)
Calverley et al, ¹¹ 2003								
Treatment	Budesonide (400 μg) twice daily	257	36 (10)	100 (39)	12	Definition not provided	Days to first exacerbation; percent	178 vs 96 (P = .51); 2% (P = .15)
Control	Placebo twice daily	256	36 (10)	77 (30)			change in ${ m FEV}_1$	
Treatment	Budesonide (320 μg)/ formoterol (9 μg) twice daily	254	36 (10)	84 (33)	12	Definition not provided	Days to first exacerbation; percent	254 vs 154 ($P < .01$); 5% ($P <$
Control	Formoterol (9 µg) twice daily	255	36 (10)	92 (36)			change in FEV ₁	(10)
Calverley et al, ¹² 2007								
Treatment	Fluticasone (500 µg) twice daily	1534	44 (12.3)	661 (43)	36	Definition not provided	All-cause mortality	$16.0\% \text{ vs } 15.2\% \ (P = .53)$
Control	Placebo twice daily	1524	44 (12.3)	658 (43)				

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Source/Intervention	Medication	Š.	FEV ₁ , % Predicted, Mean (SD)	Current Smoker, No. (%)	Trial Length, mo	of Pneumonia in Primary Manuscript	Primary Outcome	Outcome Results (Treatment vs Control)
Treatment	Fluticasone (500 μg)/ salmeterol (50 μg) twice daily	1533	44 (12.3)	660 (43)	36	Definition not provided	All-cause mortality	12.6% vs 13.5% (P = .48)
Control	Salmeterol (50 µg) twice daily	1521	44 (12.6)	651 (43)				
Kardos et al, ¹³ 2007 Treatment	Fluticasone (500 µg)/ salmeterol (50 µg) twice daily	507	40 (8.9)	208 (41)	12	Suspected pneumonia; confirmation by radioeraph	No. of moderate and	334 vs 464 exacerbations (P <
Control	Salmeterol (50 µg) twice daily	487	40 (8.5)	214 (44)		not required.	severe exacerbations	(001)
Lung Health Study Research Group, ³⁴ 2000 Treatment Triamcinolone (6	arch Group, ³⁴ 2000 Triamcinolone (600 µg)	559	65 (13.5)	509 (91)				
Control	Placebo twice daily	557	63 (13.2)	501 (90)	40a	Pneumonia outcome not reported	Decline in FEV ₁	44.2 (SE, 2.9) mL/y vs 47.0 (SE, 3.0) mL/y (P = .50)
Pauwels et al, ³⁶ 1999								
Treatment	Budesonide (400 µg) twice daily	634	77 (12.4)	634 (100)	36	Pneumonia outcome not	Change in FEV_1 over 3	140 mL vs 180 mL ($P = .05$)
Control	Placebo twice daily	643	77 (13.2)	643 (100)		reported	years	
Szafranski et al, ¹⁴ 2003								
Treatment	Budesonide (200 μg) twice daily	198	37	71 (36)	12	Pneumonia outcome not	Percent reduction in severe exacerbation rate;	15% (95% CI, -10.3% to 34.1%; P = .22); 5% (95% CI, 1.5% to
Control	Placebo twice daily	205	36	70 (34)		reported	percent change in FEV_1	9.1%; P = .005)
Treatment	Budesonide (160 μg)/ formoterol (4.5 μg) twice daily	208	36	62 (30)	12	Pneumonia outcome not	Percent reduction in severe exacerbation rate;	23% (95% CI, 0.8% to 40.1%; <i>P</i> = .043); 1% (95% CI, –2.2% to
Control	Formoterol (4.5 µg)	201	36	76 (38)		nanodar	percent change in FEV_1	4.9%; P = .49)

			FEV ₁ , %	Current	- F	Definition of Pneumonia	D.	Outcome Results
Source/Intervention	Medication	No.	Mean (SD)	No. (%)	Length, mo	m r mary Manuscript	Outcome	Control)
Vestbo et al, ³⁷ 1999								
Treatment	Budesonide (800 μg every morning/400 μg every evening $\times 4$ mo; then 400 μg twice daily)	145	86 (20.8)	110 (76)	ć	:	Pooling in DDV	45.1 mL/y vs 41.8 mL/y ($P =$
Control	Placebo (2 puffs every moming/1 puff every evening × 4 mo; then 1 puff twice daily)	145	87 (21.1)	112 (77)	30	Definition not provided		70)
Wedzicha et al, ¹⁵ 2008								
Treatment	Fluticasone (500 μg)/ salmeterol (50 μg) twice daily	658	39	250 (38)	24	Clinical judgment; confirmation by radiograph	Exacerbation rate	1.28 vs 1.32 exacerbations/y (<i>P</i> = .66)
Control	Tiotropium (18 µg) daily	999	39	253 (38)		not required		
Zheng et al, ¹⁶ 2007 Treatment	Fluticasone (500 µg)/ salmeterol (50 µg) twice daily	297	47	62 (21)	9	Pneumonia outcome not reported	6-mo change in FEV_1	180 mL (95% CI, 91 to 268; <i>P</i> < .001) higher comparing
Control	Placebo twice daily	148	47	34 (23)				reaunent vs contor

Abbreviations: CI, confidence interval; FEV1, forced expiratory volume in first second of expiration.

 $[^]a$ Average length of follow-up.

Table 2

Qualitative Analysis of Included Studies^a

Source	Adequate Sequence Generation Described	Allocation Concealment Described	Assessor Blinding Described	Incomplete Outcome Data Addressed	Free From Selective Reporting	Free From Other Bias	Eligibility Criteria Defined	Excluded Patients Described	All Therapies Described	$\begin{array}{c} \text{Jadad} \\ \text{Quality} \\ \text{Score}^b \end{array}$
Aaron et al, ³⁵ 2007	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	5
Burge et al, ¹⁰ 2000	Yes	Yes	Yes	Yes	Yes	No	Yes	Yes	Yes	4
Calverley et al, 11 2003	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	3
Calverley et al, 12 2007	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	5
Kardos et al, ¹³ 2007	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	4
Lung Health Study Research Group,34 2000	Unclear	Unclear	Yes	Unclear	Yes	Yes	Yes	Yes	Yes	3
Pauwels et al, ³⁶ 1999	Unclear	Unclear	Yes	Yes	Yes	Yes	Yes	Yes	Yes	3
Szafranski et al, ¹⁴ 2003	Unclear	Unclear	Yes	Unclear	Yes	No	Yes	Yes	Yes	3
Vestbo et al, ³⁷ 1999	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	5
Wedzicha et al, 15 2008	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	5
Zheng et al, ¹⁶ 2007	Unclear	Unclear	Yes	Yes	Yes	Yes	Yes	Yes	Yes	3

^aQualitative assessment was performed using the Cochrane bias assessment method.³⁰ Assessment of each criteria could be "yes," "no," or "unclear."

 b Range, 1–5; 5 indicates highest quality. 31

Drummond et al. Page 19

Subgroup Analysis of Inhaled Corticosteroids and 1-Year Mortality, Pneumonia, and Fracture

Table 3

		No. of Participants	cipants		
	No. of Studies	Treatment	Control	RR (95% CI)	I^2 , %
		1-y Mortality			
Overall	5	4636	4597	0.86 (0.68–1.09)	0
Dose, mg ^a					
<800	2	462	456	0.57 (0.28–1.20)	30
800-1000	2	455	461	0.80 (0.37–1.74)	0
>1000	3	3719	3680	0.91 (0.71–1.19)	0
Duration, y					
2	4	1569	1552	0.73 (0.47–1.12)	0
>2	1	3067	3045	0.92 (0.70–1.22)	4
Baseline FEV ₁ (% predicted)					
<40	3	1062	1065	0.72 (0.45–1.16)	0
40–50	2	3574	3532	0.91 (0.70–1.19)	0
>50	0	0	0		
Modality					
Monotherapy	3	1989	1985	0.78 (0.56–1.10)	0
Combination	5	2647	2612	0.94 (0.68–1.29)	0
		Pneumonia			
Overall	7	5405	5371	1.34 (1.03–1.75)	72
Dose, mg ^a					
<800	1	254	255	1.15 (0.42–3.12)	0
800–1000	2	402	401	1.03 (0.30–3.47)	55

Drummond et al.

		No. of Participants	cipants		
	No. of Studies	Treatment	Control	RR (95% CI)	I^2 , %
>1000	5	4749	4715	1.46 (1.10–1.92)	78
Duration, y	4	1821	1811	2.12 (1.47–3.05)	0
>2	ю	3584	3560	1.14 (0.84–1.54)	85
Baseline FEV ₁ (% predicted)					
<40	3	1314	1324	1.90 (1.26–2.85)	0
40–50	3	3946	3902	1.36 (1.00–1.86)	84
>50	1	145	145	0.67 (0.37–1.20)	0
Modality					
Monotherapy	4	2308	2295	1.06 (0.67–1.67)	83
Combination	5	3097	3076	1.57 (1.35–1.82)	24
		Fracture			
Overall	3	4073	4058	1.09 (0.89–1.33)	29
Dose, mg ^a					
<800	0	0	0		
800-1000	1	634	643	1.69 (0.41–7.04)	48
>1000	2	3439	3415	1.08 (0.88–1.32)	29
Duration, y					
2	0	0	0		
>2	3	4073	4058	1.09 (0.89–1.33)	29
Baseline FEV ₁ (% predicted) <40	0	0	0		

Page 20

Drummond et al.

		No. of Participants	ipants		
	No. of Studies	Treatment Control	Control	RR (95% CI) 1 ² , %	I^2 , %
40–50	2	3439	3415	3415 1.08 (0.88–1.31)	48
>50	1	634	643	643 1.69 (0.41–7.04)	29
Modality	,	1			
Monotherapy	3	2540	2537	2537 0.98 (0.75–1.29)	36
Combination	1	1533	1521	1521 1.22 (0.91–1.63)	0

Abbreviations: CI, confidence interval; FEV1, forced expiratory volume in first second of expiration; RR, relative risk.

^aICS dose in each study converted to beclomethasone equivalents. ²⁹

Page 21